# Package 'mev'

November 11, 2025

```
Type Package
Title Modelling of Extreme Values
Version 2.1
Description Various tools for the analysis of univariate, multivariate and functional extremes. Ex-
     act simulation from max-stable processes (Dombry, Engelke and Oest-
     ing, 2016, <doi:10.1093/biomet/asw008>, R-Pareto processes for various parametric models, in-
     cluding Brown-Resnick (Wadsworth and Tawn, 2014, <doi:10.1093/biomet/ast042>) and Ex-
     tremal Student (Thibaud and Opitz, 2015, <doi:10.1093/biomet/asv045>). Threshold selec-
     tion methods, includ-
     ing Wadsworth (2016) <doi:10.1080/00401706.2014.998345>, and Northrop and Cole-
     man (2014) <doi:10.1007/s10687-014-0183-z>. Multivariate extreme diagnostics. Estima-
     tion and likelihoods for univariate extremes, e.g., Coles (2001) <doi:10.1007/978-1-4471-3675-
     0>.
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BugReports https://github.com/lbelzile/mev/issues/
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```

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abisko

Abisko rainfall

### **Description**

Daily non-zero rainfall measurements in Abisko (Sweden) from January 1913 until December 2014.

### **Arguments**

```
date Date of the measurement precip rainfall amount (in mm)
```

### **Format**

a data frame with 15132 rows and two variables

### **Source**

Abisko Scientific Research Station

### References

A. Kiriliouk, H. Rootzen, J. Segers and J.L. Wadsworth (2019), *Peaks over thresholds modeling with multivariate generalized Pareto distributions*, Technometrics, **61**(1), 123–135, <doi:10.1080/00401706.2018.1462738>

adf

Estimation of the bivariate angular dependence function

# Description

Estimation of the bivariate angular dependence function

# Usage

```
adf(
   xdat,
   qlev = 0.95,
   estimator = c("hill", "mle", "bayes"),
   level = 0.95,
   ties.method = "random",
   angles = seq(0, 1, by = 0.02),
   plot = TRUE
)
```

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#### **Arguments**

xdat an n by 2 matrix of multivariate observations

qlev quantile level on uniform scale at which to threshold data. Default to 0.95

estimator string indicating the estimation method

level level for confidence intervals, default to 0.95

ties.method method for handling of ties in rank transformation

angles vector of angles at which to evaluate the angular dependence function The con-

fidence intervals are based on normal quantiles. The standard errors for the hill are based on the asymptotic covariance and that of the mle derived using the delta-method. Bayesian posterior predictive interval estimates are obtained using ratio-of-uniform sampling with flat priors: the shape parameters are constrained to lie within the triangle, as are frequentist point estimates which are

adjusted post-inference.

plot logical indicating whether to plot the function, defaults to TRUE

#### Value

a plot of the angular dependence function if plot=TRUE, plus an invisible list with components

- angle the sequence of angles in (0,1) at which the lambda values are evaluated
- coef point estimates of the angular dependence function
- lower level% confidence interval for lambda (lower bound)
- upper level% confidence interval for lambda (upper bound)

#### References

J.L. Wadsworth and J.A. Tawn (2013). A new representation for multivariate tail probabilities, *Bernoulli*, 19(**5B**), 2689-2714.

# **Examples**

```
set.seed(12)
dat <- mev::rmev(n = 1000, d = 2, model = "log", param = 0.1)
adf(xdat = dat, estimator = 'hill')</pre>
```

angmeas

Rank-based transformation to angular measure

### Description

The method uses the pseudo-polar transformation for suitable norms, transforming the data to pseudo-observations, than marginally to unit Frechet or unit Pareto. Empirical or Euclidean weights are computed and returned alongside with the angular and radial sample for values above threshold(s) thresh, specified in terms of quantiles of the radial component R or marginal quantiles. Only complete tuples are kept.

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### Usage

```
angmeas(
   xdat,
   thresh,
   Rnorm = c("l1", "l2", "linf"),
   Anorm = c("l1", "l2", "linf", "arctan"),
   marg = c("frechet", "pareto"),
   wgt = c("empirical", "euclidean"),
   region = c("sum", "min", "max"),
   is.angle = FALSE,
   ...
)
```

### **Arguments**

xdat	an n by d sample matrix
thresh	threshold of length 1 for 'sum', or d marginal thresholds otherwise.
Rnorm	character string indicating the norm for the radial component.
Anorm	character string indicating the norm for the angular component. arctan is only implemented for $d=2$
marg	character string indicating choice of marginal transformation, either to Frechet or Pareto scale
wgt	character string indicating weighting function for the equation. Can be based on Euclidean or empirical likelihood for the mean
region	character string specifying which observations to consider (and weight). 'sum' corresponds to a radial threshold $\sum x_i$ >thresh, 'min' to $\min x_i$ >thresh and 'max' to $\max x_i$ >thresh.
is.angle	logical indicating whether observations are already angle with respect to region. Default to FALSE.
	additional arguments

### **Details**

The empirical likelihood weighted mean problem is implemented for all thresholds, while the Euclidean likelihood is only supported for diagonal thresholds specified via region=sum.

### Value

a list with arguments ang for the d-1 pseudo-angular sample, rad with the radial component and possibly wts if Rnorm='11' and the empirical likelihood algorithm converged. The Euclidean algorithm always returns weights even if some of these are negative.

a list with components

- ang matrix of pseudo-angular observations
- rad vector of radial contributions
- wts empirical or Euclidean likelihood weights for angular observations

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#### Author(s)

Leo Belzile

#### References

Einmahl, J.H.J. and J. Segers (2009). Maximum empirical likelihood estimation of the spectral measure of an extreme-value distribution, *Annals of Statistics*, **37**(5B), 2953–2989.

de Carvalho, M. and B. Oumow and J. Segers and M. Warchol (2013). A Euclidean likelihood estimator for bivariate tail dependence, *Comm. Statist. Theory Methods*, **42**(7), 1176–1192.

Owen, A.B. (2001). Empirical Likelihood, CRC Press, 304p.

### **Examples**

```
x \leftarrow rmev(n = 25, d = 3, param = 0.5, model = 'log') wts <- angmeas(xdat = x, Rnorm = 'l1', Anorm = 'l1', marg = 'frechet', wgt = 'empirical') wts2 <- angmeas(xdat = x, Rnorm = 'l2', Anorm = 'l2', marg = 'pareto')
```

angmeasdir

Dirichlet mixture smoothing of the angular measure

### **Description**

This function computes the empirical or Euclidean likelihood estimates of the spectral measure and uses the points returned from a call to angmeas to compute the Dirichlet mixture smoothing of de Carvalho, Warchol and Segers (2012), placing a Dirichlet kernel at each observation.

### Usage

```
angmeasdir(
  xdat,
  thresh,
  Rnorm = c("11", "12", "linf"),
  Anorm = c("11", "12", "linf", "arctan"),
  marg = c("frechet", "pareto"),
  wgt = c("empirical", "euclidean"),
  region = c("sum", "min", "max"),
  is.angle = FALSE,
  ...
)
```

### **Arguments**

xdat an n by d sample matrix

thresh threshold of length 1 for 'sum', or d marginal thresholds otherwise.

Rnorm character string indicating the norm for the radial component.

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Anorm	character string indicating the norm for the angular component. arctan is only implemented for $d=2$
marg	character string indicating choice of marginal transformation, either to Frechet or Pareto scale
wgt	character string indicating weighting function for the equation. Can be based on Euclidean or empirical likelihood for the mean
region	character string specifying which observations to consider (and weight). 'sum' corresponds to a radial threshold $\sum x_i$ >thresh, 'min' to $\min x_i$ >thresh and 'max' to $\max x_i$ >thresh.
is.angle	logical indicating whether observations are already angle with respect to region. Default to FALSE.
	additional arguments

### **Details**

The cross-validation bandwidth is the solution of

$$\max_{\nu} \sum_{i=1}^{n} \log \left\{ \sum_{k=1, k \neq i}^{n} p_{k,-i} f(\mathbf{w}_i; \nu \mathbf{w}_k) \right\},\,$$

where f is the density of the Dirichlet distribution,  $p_{k,-i}$  is the Euclidean weight obtained from estimating the Euclidean likelihood problem without observation i.

# Value

an invisible list with components

- nu bandwidth parameter obtained by cross-validation;
- dirparmat n by d matrix of Dirichlet parameters for the mixtures;
- wts mixture weights.

# **Examples**

```
set.seed(123) x \leftarrow rmev(n = 100, d = 2L, param = 0.5, model = 'log') out \leftarrow angmeasdir(x)
```

confint.eprof

Confidence intervals for profile likelihood objects

### **Description**

Computes confidence intervals for the parameter psi for profile likelihood objects. This function uses spline interpolation to derive level confidence intervals

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### Usage

```
## S3 method for class 'eprof'
confint(
  object,
  parm,
  level = 0.95,
  prob = c((1 - level)/2, 1 - (1 - level)/2),
  print = FALSE,
  method = c("cobs", "smooth.spline"),
  boundary = FALSE,
  ...
)
```

### **Arguments**

object an object of class eprof, normally the output of gpd.pll or gev.pll. a specification of which parameters are to be given confidence intervals, either parm a vector of numbers or a vector of names. If missing, all parameters are considconfidence level, with default value of 0.95 level prob percentiles, with default giving symmetric 95% confidence intervals should a summary be printed. Default to FALSE. print method string for the method, either cobs (constrained robust B-spline from eponym package) or smooth.spline boundary logical; if TRUE, the null distribution is assumed to be a mixture of a point mass and half a chi-square with one degree of freedom. additional arguments passed to functions. Providing a logical warn=FALSE turns off warning messages when the lower or upper confidence interval for psi are

#### Value

returns a 2 by 3 matrix containing point estimates, lower and upper confidence intervals based on the likelihood root and modified version thereof

extrapolated beyond the provided calculations.

dgeoaniso

Distance matrix with geometric anisotropy

### **Description**

The function computes the distance between locations, with geometric anisotropy. Consider real parameters  $\theta_1$  and  $\theta_2$ , and the transformation  $\psi = \arctan(\theta_1/\theta_2)/2$  and  $r = 1 + \theta_1^2 + \theta_2^2$ . The dilation and rotation matrix is

$$\begin{pmatrix} \sqrt{r}\cos(\rho) & -\sqrt{r}\sin(\rho) \\ \sin(\rho)/\sqrt{r} & \cos(\rho)/\sqrt{r} \end{pmatrix}.$$

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The parametrization is convenient for optimization purposes, as the parameter vector is unconstrained and the transformation has unit Jacobian.

### Usage

```
dgeoaniso(loc, theta)
```

### **Arguments**

loc a d by 2 matrix of locations giving the coordinates of a site per row.

theta numeric vector of length 2, real parameters

#### Value

a d by d square matrix of pairwise distance

#### References

Rai, K. and Brown, P.E. (2025), A parameter transformation of the anisotropic Matérn covariance function. Canadian Journal of Statistics e11839. doi:10.1002/cjs.11839

egp

Extended generalised Pareto families

### Description

This function provides the log-likelihood and quantiles for the three different families presented in Papastathopoulos and Tawn (2013) and the two proposals of Gamet and Jalbert (2022), plus exponential tilting. All of the models contain an additional parameter,  $\kappa \geq 0$ . All families share the same tail index as the generalized Pareto distribution, while allowing for lower thresholds. For most models, the distribution reduce to the generalised Pareto when  $\kappa = 1$  (for models gj-tnorm and logist, on the boundary of the parameter space when  $\kappa \to 0$ ).

egp.retlev gives the return levels for the extended generalised Pareto distributions

#### **Arguments**

xdat	vector of observations, greater than the threshold
thresh	threshold value
par	parameter vector $(\kappa, \sigma, \xi)$ .
model	a string indicating which extended family to fit
show	logical; if TRUE, print the results of the optimization
p	extreme event probability; $p$ must be greater than the rate of exceedance for the calculation to make sense. See <b>Details</b> .
plot	logical; if TRUE, a plot of the return levels

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#### **Details**

For return levels, the p argument can be related to T year exceedances as follows: if there are  $n_y$  observations per year, than take p to equal  $1/(Tn_y)$  to obtain the T-years return level.

### Value

egp.11 returns the log-likelihood value, while egp.retlev returns a plot of the return levels if plot=TRUE and a list with tail probabilities p, return levels retlev, thresholds thresh and model name model.

### Usage

```
egp.ll(xdat, thresh, model, par)
egp.retlev(xdat, thresh, par, model, p, plot=TRUE)
```

#### Author(s)

Leo Belzile

#### References

Papastathopoulos, I. and J. Tawn (2013). Extended generalised Pareto models for tail estimation, *Journal of Statistical Planning and Inference* **143**(3), 131–143, <doi:10.1016/j.jspi.2012.07.001>.

Gamet, P. and Jalbert, J. (2022). A flexible extended generalized Pareto distribution for tail estimation. *Environmetrics*, **33**(6), <doi:10.1002/env.2744>.

### **Examples**

```
set.seed(123)
xdat <- rgp(1000, loc = 0, scale = 2, shape = 0.5)
par <- fit.egp(xdat, thresh = 0, model = 'gj-beta')$par
p <- c(1/1000, 1/1500, 1/2000)
# With multiple thresholds
th <- c(0, 0.1, 0.2, 1)
opt <- tstab.egp(xdat, thresh = th, model = 'gj-beta')
egp.retlev(xdat = xdat, thresh = th, model = 'gj-beta', p = p)
opt <- tstab.egp(xdat, th, model = 'pt-power', plots = NA)
egp.retlev(xdat = xdat, thresh = th, model = 'pt-power', p = p)</pre>
```

egp.pll

Profile log likelihood for extended generalized Pareto models

### **Description**

Computes the profile log likelihood over a grid of values of  $\psi$  for various parameters, including return levels.

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# Usage

```
egp.pll(
   psi,
   model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
        "logist"),
   param = c("kappa", "scale", "shape", "retlev"),
   mle = NULL,
   xdat,
   thresh = NULL,
   plot = FALSE,
   method = c("Nelder", "nlminb", "BFGS"),
   p,
   ...
)
```

### **Arguments**

ļ	osi	grid of values for the parameter to profile
r	model	string; choice of extended eneralized Pareto model.
ļ	oaram	string; parameter to profile
r	nle	a vector or matrix with maximum likelihood estimates of kappa, scale, shape. This can be a matrix if there are multiple threshold
)	xdat	vector of observations
1	thresh	vector of positive thresholds. If NULL, defaults to zero.
ļ	olot	logical; if TRUE, returns a plot of the profile log likelihood
r	method	string giving the optimization method for the outer optimization in the augmented Lagrangian routine; one of $nlminb$ or BFGS
ļ	0	tail probability for return level if param="retlev".
	• • •	additional arguments, currently ignored

### Value

an object of class eprof

egpdist	Extended generalized Pareto distribution
egpdist	Extended generalized Pareto distribution

# Description

Density function, distribution function, quantile function and random number generation for various extended generalized Pareto distributions

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### Usage

```
pegp(
  q,
  scale,
  shape,
  kappa,
 model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
 lower.tail = TRUE,
 log.p = FALSE
)
degp(
  Х,
  scale,
  shape,
 kappa,
 model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
 log = FALSE
)
qegp(
 р,
  scale,
  shape,
 model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist"),
  lower.tail = TRUE,
 log.p = FALSE
)
regp(
  n,
  scale,
  shape,
 kappa,
 model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
    "logist")
)
```

### **Arguments**

```
scale scale parameter, strictly positive.

shape shape parameter.

kappa shape parameter for the tilting distribution.
```

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model	string giving the distribution of the model
lower.tail	logical; if TRUE (default), the lower tail probability $\Pr(X \leq x)$ is returned.
log.p,log	logical; if FALSE (default), values are returned on the probability scale.
x, q	vector of quantiles
p	vector of probabilities
n	scalar number of observations

### References

Papastathopoulos, I. and J. Tawn (2013). Extended generalised Pareto models for tail estimation, *Journal of Statistical Planning and Inference* **143**(3), 131–143, <doi:10.1016/j.jspi.2012.07.001>.

Gamet, P. and Jalbert, J. (2022). A flexible extended generalized Pareto distribution for tail estimation. *Environmetrics*, **33**(6), <doi:10.1002/env.2744>.

emplik

Self-concordant empirical likelihood for a vector mean

### **Description**

Self-concordant empirical likelihood for a vector mean

# Usage

```
emplik(
   dat,
   mu = rep(0, ncol(dat)),
   lam = rep(0, ncol(dat)),
   eps = 1/nrow(dat),
   M = 1e+30,
   thresh = 1e-30,
   itermax = 100
)
```

### **Arguments**

dat	n by d matrix of d-variate observations
mu	d vector of hypothesized mean of dat
lam	starting values for Lagrange multiplier vector, default to zero vector
eps	lower cutoff for $-\log$ , with default 1/nrow(dat)
М	upper cutoff for $-\log$ .
thresh	convergence threshold for log likelihood (default of 1e-30 is aggressive)
itermax	upper bound on number of Newton steps.

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### Value

a list with components

- logelr log empirical likelihood ratio.
- lam Lagrange multiplier (vector of length d).
- wts n vector of observation weights (probabilities).
- conv boolean indicating convergence.
- niter number of iteration until convergence.
- ndec Newton decrement.
- gradnorm norm of gradient of log empirical likelihood.

### Author(s)

Art Owen, C++ port by Leo Belzile

### References

Owen, A.B. (2013). Self-concordance for empirical likelihood, *Canadian Journal of Statistics*, **41**(3), 387–397.

eskrain

Eskdalemuir Observatory Daily Rainfall

# **Description**

This dataset contains exceedances of 30mm for daily cumulated rainfall observations over the period 1970-1986. These data were aggregated from hourly series.

### **Format**

a vector with 93 daily cumulated rainfall measurements exceeding 30mm.

### **Details**

The station is one of the rainiest of the whole UK, with an average 1554m of cumulated rainfall per year. The data consisted of 6209 daily observations, of which 4409 were non-zero. Only the 93 largest observations are provided.

### Source

Met Office.

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expme

Exponent measure for multivariate generalized Pareto distributions

### **Description**

Integrated intensity over the region defined by  $[0, z]^c$  for logistic, Huesler-Reiss, Brown-Resnick and extremal Student processes.

### Usage

```
expme(
   z,
   par,
   model = c("log", "neglog", "hr", "br", "xstud"),
   method = c("TruncatedNormal", "mvtnorm", "mvPot")
)
```

# **Arguments**

z vector at which to estimate exponent measure

par list of parameters

model string indicating the model family

method string indicating the package from which to extract the numerical integration

routine

### Value

numeric giving the measure of the complement of [0, z].

### Note

The list par must contain different arguments depending on the model. For the Brown–Resnick model, the user must supply the conditionally negative definite matrix Lambda following the parametrization in Engelke *et al.* (2015) or the covariance matrix Sigma, following Wadsworth and Tawn (2014). For the Husler–Reiss model, the user provides the mean and covariance matrix, m and Sigma. For the extremal student, the covariance matrix Sigma and the degrees of freedom df. For the logistic model, the strictly positive dependence parameter alpha.

### **Examples**

```
## Not run:
# Extremal Student
Sigma <- stats::rWishart(n = 1, df = 20, Sigma = diag(10))[, , 1]
expme(z = rep(1, ncol(Sigma)), par = list(Sigma = cov2cor(Sigma), df = 3), model = "xstud")
# Brown-Resnick model
D <- 5L
loc <- cbind(runif(D), runif(D))</pre>
```

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```
di <- as.matrix(dist(rbind(c(0, ncol(loc)), loc)))
semivario <- function(d, alpha = 1.5, lambda = 1) {
    (d / lambda)^alpha
}
Vmat <- semivario(di)
Lambda <- Vmat[-1, -1] / 2
expme(z = rep(1, ncol(Lambda)), par = list(Lambda = Lambda), model = "br", method = "mvPot")
Sigma <- outer(Vmat[-1, 1], Vmat[1, -1], "+") - Vmat[-1, -1]
expme(z = rep(1, ncol(Lambda)), par = list(Lambda = Lambda), model = "br", method = "mvPot")
## End(Not run)</pre>
```

extcoef

Estimators of the extremal coefficient

#### **Description**

These functions estimate the extremal coefficient using an approximate sample from the Frechet distribution.

### Usage

```
extcoef(
  xdat,
  coord = NULL,
  thresh = NULL,
  estimator = c("schlather", "smith", "fmado"),
  margtrans = c("emp", "gev", "none"),
  ties.method = "random",
  prob = 0,
  plot = TRUE,
  ...
)
```

### **Arguments**

xdat an n by D matrix of unit Frechet observations coord an optional d by D matrix of location coordinates threshold parameter (default is to keep all data if prob = 0). thresh string indicating which estimator to compute, one of smith, schlather or fmado. estimator string indicating which method to use to transform the margins to unit Frechet margtrans scale, either emp for nonparametric transformation via rank transform, gev for fit of generalized extreme value distribution to marginals, or none method for handling of ties in rank transformation ties.method prob probability of not exceeding threshold thresh logical; should cloud or matrix of pairwise empirical estimates be plotted? Deplot fault to TRUE.

additional parameters passed to the function, currently ignored.

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#### **Details**

The **Smith** estimator: suppose Z(x) is simple max-stable vector (i.e., with unit Frechet marginals). Then 1/Z is unit exponential and  $1/\max(Z(s_1), Z(s_2))$  is exponential with rate  $\theta = \max\{Z(s_1), Z(s_2)\}$ . The extremal index for the pair can therefore be calculated using the reciprocal mean.

The **Schlather and Tawn** estimator: the likelihood of the naive estimator for a pair of two sites A is

$$\operatorname{card}\left\{j: \max_{i \in A} X_i^{(j)} \bar{X}_i) > z\right\} \log(\theta_A) - \theta_A \sum_{i=1}^n \left[ \max\left\{z, \max_{i \in A} (X_i^{(j)} \bar{X}_i)\right\} \right]^{-1},$$

where  $\bar{X}_i = n^{-1} \sum_{j=1}^n 1/X_i^{(j)}$  is the harmonic mean and z is a threshold on the unit Frechet scale. The search for the maximum likelihood estimate for every pair A is restricted to the interval [1,3]. A binned version of the extremal coefficient cloud is also returned. The Schlather estimator is not self-consistent. The Schlather and Tawn estimator includes as special case the Smith estimator if we do not censor the data (p = 0) and do not standardize observations by their harmonic mean.

The **F-madogram** estimator is a non-parametric estimate based on a stationary process Z; the extremal coefficient satisfies

$$\theta(h) = \frac{1 + 2\nu(h)}{1 - 2\nu(h)},$$

where

$$\nu(h) = \frac{1}{2}\mathsf{E}[|F(Z(s+h) - F(Z(s))|]$$

The implementation only uses complete pairs to calculate the relative ranks.

All estimators are coded in plain R and computations are not optimized. The estimation time can therefore be large for large data sets. If there are no missing observations, the routine fmadogram from the SpatialExtremes package should be preferred as it is noticeably faster.

The data will typically consist of max-stable vectors or block maxima. Both of the Smith and the Schlather-Tawn estimators require unit Frechet margins; the margins will be standardized to the unit Frechet scale, either parametrically or nonparametrically unless margtrans="none". If margtrans = "gev", a parametric GEV model is fitted to each column of dat using maximum likelihood estimation and transformed back using the probability integral transform. If method = "emp", using the empirical distribution function. The latter is the default, as it is appreciably faster.

### Value

an invisible list with vectors dist if coord is non-null or else a matrix of pairwise indices ind, extcoef and the supplied estimator, fmado and binned. If estimator == "schlather", an additional matrix with 2 columns containing the binned distance binned with the h and the binned extremal coefficient.

# References

Schlather, M. and J. Tawn (2003). A dependence measure for multivariate and spatial extremes, *Biometrika*, **90**(1), pp.139–156.

Cooley, D., P. Naveau and P. Poncet (2006). Variograms for spatial max-stable random fields, In: Bertail P., Soulier P., Doukhan P. (eds) *Dependence in Probability and Statistics*. Lecture Notes in Statistics, vol. 187. Springer, New York, NY

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R. J. Erhardt, R. L. Smith (2012), Approximate Bayesian computing for spatial extremes, *Computational Statistics and Data Analysis*, **56**, pp.1468–1481.

# **Examples**

```
## Not run:
coord <- 10 * cbind(runif(50), runif(50))</pre>
di <- as.matrix(dist(coord))</pre>
dat <- rmev(</pre>
 n = 1000.
  d = 100,
  param = 3,
  sigma = exp(-di / 2),
  model = 'xstud'
res <- extcoef(xdat = dat, coord = coord)
# Extremal Student extremal coefficient function
XT.extcoeffun <- function(h, nu, corrfun, ...) {</pre>
  if (!is.function(corrfun)) {
    stop('Invalid function \"corrfun\".')
  h <- unique(as.vector(h))</pre>
  rhoh <- sapply(h, corrfun, ...)</pre>
  cbind(
    h = h,
    extcoef = 2 * pt(sqrt((nu + 1) * (1 - rhoh) / (1 + rhoh)), nu + 1)
  )
}
#This time, only one graph with theoretical extremal coef
plot(res$dist, res$extcoef, ylim = c(1, 2), pch = 20)
abline(v = 2, col = 'gray')
extcoefxt <- XT.extcoeffun(</pre>
  seq(0, 10, by = 0.1),
  corrfun = function(x) {
    exp(-x / 2)
  }
)
lines(
  extcoefxt[, 'h'],
  extcoefxt[, 'extcoef'],
  type = '1',
  col = 'blue',
  1wd = 2
# Brown--Resnick extremal coefficient function
BR.extcoeffun <- function(h, vario, ...) {</pre>
  if (!is.function(vario)) {
    stop('Invalid function \"vario\".')
  h <- unique(as.vector(h))</pre>
  gammah <- sapply(h, vario, ...)</pre>
```

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```
cbind(h = h, extcoef = 2 * pnorm(sqrt(gammah / 4)))
}
extcoefbr <- BR.extcoeffun(</pre>
  seq(0, 20, by = 0.25),
  vario = function(x) {
    2 * x^0.7
)
lines(
  extcoefbr[, 'h'],
  extcoefbr[, 'extcoef'],
  type = '1',
  col = 'orange',
  1wd = 2
)
coord <- 10 * cbind(runif(20), runif(20))</pre>
di <- as.matrix(dist(coord))</pre>
dat <- rmev(</pre>
 n = 1000,
  d = 20,
 param = 3,
  sigma = exp(-di / 2),
  model = 'xstud'
)
res <- extcoef(
  xdat = dat,
  coord = coord,
  estimator = "smith"
)
## End(Not run)
```

extgp

Extended generalised Pareto families of Naveau et al. (2016)

### **Description**

Density function, distribution function, quantile function and random generation for the extended generalized Pareto distribution (GPD) with scale and shape parameters.

### **Arguments**

q	vector of quantiles
Х	vector of observations
p	vector of probabilities
n	sample size
prob	mixture probability for model type 4

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kappa	shape parameter for type 1, 3 and 4
delta	additional parameter for type 2, 3 and 4
sigma	scale parameter
xi	shape parameter
type	integer between 0 to 5 giving the model choice
step	function of step size for discretization with default $\emptyset$ , corresponding to continuous quantiles
log	logical; should the log-density be returned (default to FALSE)?
unifsamp	sample of uniform; if provided, the data will be used in place of new uniform random variates
censoring	numeric vector of length 2 containing the lower and upper bound for censoring

#### **Details**

The extended generalized Pareto families proposed in Naveau *et al.* (2016) retain the tail index of the distribution while being compliant with the theoretical behavior of extreme low rainfall. There are five proposals, the first one being equivalent to the GP distribution.

- type 0 corresponds to uniform carrier, G(u) = u.
- type 1 corresponds to a three parameters family, with carrier  $G(u) = u^{\kappa}$ .
- type 2 corresponds to a three parameters family, with carrier  $G(u) = 1 V_{\delta}((1-u)^{\delta})$ .
- type 3 corresponds to a four parameters family, with carrier

$$G(u) = 1 - V_{\delta}((1 - u)^{\delta}))^{\kappa/2}$$

• type 4 corresponds to a five parameter model (a mixture of type 2, with  $G(u)=pu^\kappa+(1-p)*u^\delta$ 

#### Usage

```
pextgp(q, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1)
dextgp(x, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1, log=FALSE)
qextgp(p, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1)
rextgp(n, prob=NA, kappa=NA, delta=NA, sigma=NA, xi=NA, type=1, unifsamp=NULL, censoring=c(0,Inf))
```

#### Author(s)

Raphael Huser and Philippe Naveau

#### References

Naveau, P., R. Huser, P. Ribereau, and A. Hannart (2016), Modeling jointly low, moderate, and heavy rainfall intensities without a threshold selection, *Water Resour. Res.*, 52, 2753-2769, doi:10.1002/2015WR018552.

extgp.G

extgp.G	Carrier distribution for the extended GP distributions of Naveau et al.

# Description

Density, distribution function, quantile function and random number generation for the carrier distributions of the extended Generalized Pareto distributions.

# Arguments

u	vector of observations (dextgp.G), probabilities (qextgp.G) or quantiles (pextgp.G), in $[0,1]$
prob	mixture probability for model type 4
kappa	shape parameter for type 1, 3 and 4
delta	additional parameter for type 2, 3 and 4
type	integer between 0 to 5 giving the model choice
log	logical; should the log-density be returned (default to FALSE)?
n	sample size
unifsamp	sample of uniform; if provided, the data will be used in place of new uniform random variates
censoring	numeric vector of length 2 containing the lower and upper bound for censoring
direct	logical; which method to use for sampling in model of type 4?

# Usage

```
pextgp.G(u, type=1, prob, kappa, delta)
dextgp.G(u, type=1, prob=NA, kappa=NA, delta=NA, log=FALSE)
qextgp.G(u, type=1, prob=NA, kappa=NA, delta=NA)
rextgp.G(n, prob=NA, kappa=NA, delta=NA, type=1, unifsamp=NULL, direct=FALSE, censoring=c(0,1))
```

### Author(s)

Raphael Huser and Philippe Naveau

### See Also

extgp

24 fit.egp

fit.egp	Parameter stability plot and maximum likelihood routine for extended GP models
TIL.egp	

# Description

The function tstab.egp provides classical threshold stability plot for  $(\kappa, \sigma, \xi)$ . The fitted parameter values are displayed with pointwise normal 95% confidence intervals. The function returns an invisible list with parameter estimates and standard errors, and p-values for the Wald test that  $\kappa=1$ . The plot is for the modified scale (as in the generalised Pareto model) and as such it is possible that the modified scale be negative. tstab.egp can also be used to fit the model to multiple thresholds.

### Usage

```
fit.egp(
  xdat,
  thresh = 0,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
      "logist"),
  start = NULL,
  method = c("Nelder", "nlminb", "BFGS"),
  fpar = NULL,
  show = FALSE,
  ...
)
```

### **Arguments**

xdat	vector of observations, greater than the threshold
thresh	threshold value
model	a string indicating which extended family to fit
start	optional named list of initial values, with $\kappa$ , $sigma$ or $xi$ .
method	the method to be used. See <b>Details</b> . Can be abbreviated.
fpar	a named list with fixed parameters, either scale or shape
show	logical; if TRUE, print the results of the optimization
	additional parameters, for backward compatibility purposes

### **Details**

fit.egp is a numerical optimization routine to fit the extended generalised Pareto models of Papastathopoulos and Tawn (2013), using maximum likelihood estimation.

fit.egp 25

### Value

fit.egp outputs the list returned by optim, which contains the parameter values, the hessian and in addition the standard errors

tstab.egp returns a plot(s) of the parameters fit over the range of provided thresholds, with pointwise normal confidence intervals; the function also returns an invisible list containing notably the matrix of point estimates (par) and standard errors (se).

### Author(s)

Leo Belzile

#### References

Papastathopoulos, I. and J. Tawn (2013). Extended generalised Pareto models for tail estimation, *Journal of Statistical Planning and Inference* **143**(3), 131–143.

# Examples

```
xdat <- mev::rgp(</pre>
  n = 100,
  loc = 0,
  scale = 1,
  shape = 0.5)
fitted <- fit.egp(</pre>
  xdat = xdat,
  thresh = 1,
  model = "pt-gamma",
  show = TRUE)
thresh <- mev::qgp(seq(0.1, 0.5, by = 0.05), 0, 1, 0.5)
tstab.egp(
   xdat = xdat,
   thresh = thresh,
   model = "pt-gamma")
xdat <- regp(</pre>
  n = 100,
  scale = 1,
  shape = 0.1,
  kappa = 0.5,
  model = "pt-power"
fit.egp(
 xdat = xdat,
 model = "pt-power",
 show = TRUE,
 fpar = list(kappa = 1),
 method = "Nelder"
)
```

26 fit.extgp

fit.extgp

Fit an extended generalized Pareto distribution of Naveau et al.

# Description

This is a wrapper function to obtain PWM or MLE estimates for the extended GP models of Naveau et al. (2016) for rainfall intensities. The function calculates confidence intervals by means of nonparametric percentile bootstrap and returns histograms and QQ plots of the fitted distributions. The function handles both censoring and rounding.

# Usage

```
fit.extgp(
  data,
  model = 1,
  method = c("mle", "pwm"),
  init,
  censoring = c(0, Inf),
  rounded = 0,
  confint = FALSE,
  R = 1000,
  ncpus = 1,
  plots = TRUE
)
```

### **Arguments**

data	data vector.
model	integer ranging from 0 to 4 indicating the model to select (see extgp).
method	string; either 'mle' for maximum likelihood, or 'pwm' for probability weighted moments, or both.
init	vector of initial values, comprising of $p$ , $\kappa$ , $\delta$ , $\sigma$ , $\xi$ (in that order) for the optimization. All parameters may not appear depending on model.
censoring	numeric vector of length 2 containing the lower and upper bound for censoring; censoring=c(0, Inf) is equivalent to no censoring.
rounded	numeric giving the instrumental precision (and rounding of the data), with default of $\boldsymbol{0}$ .
confint	logical; should confidence interval be returned (percentile bootstrap).
R	integer; number of bootstrap replications.
ncpus	integer; number of CPUs for parallel calculations (default: 1).
plots	logical; whether to produce histogram and density plots.

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### **Details**

The different models include the following transformations:

- model 0 corresponds to uniform carrier, G(u) = u.
- model 1 corresponds to a three parameters family, with carrier  $G(u) = u^{\kappa}$ .
- model 2 corresponds to a three parameters family, with carrier  $G(u) = 1 V_{\delta}((1-u)^{\delta})$ .
- model 3 corresponds to a four parameters family, with carrier

$$G(u) = 1 - V_{\delta}((1-u)^{\delta}))^{\kappa/2}$$

.

• model 4 corresponds to a five parameter model (a mixture of type 2, with  $G(u)=pu^\kappa+(1-p)*u^\delta$ 

### Author(s)

Raphael Huser and Philippe Naveau

#### References

Naveau, P., R. Huser, P. Ribereau, and A. Hannart (2016), Modeling jointly low, moderate, and heavy rainfall intensities without a threshold selection, *Water Resour. Res.*, 52, 2753-2769, doi:10.1002/2015WR018552.

# See Also

```
egp.fit, egp, extgp
```

# **Examples**

```
## Not run:
data(rain, package = "ismev")
fit.extgp(
   rain[rain > 0],
   model = 1,
   method = 'mle',
   init = c(0.9, fit.gpd(rain)$est),
   rounded = 0.1,
   confint = TRUE,
   R = 20
)
## End(Not run)
```

28 fit.gev

fit.gev	Maximum likelihood estimation for the generalized extreme value distribution
	irioution

# **Description**

This function returns an object of class mev\_gev, with default methods for printing and quantile-quantile plots. The default starting values are the solution of the probability weighted moments.

### Usage

```
fit.gev(
  xdat,
  start = NULL,
  method = c("nlminb", "BFGS"),
  show = FALSE,
  fpar = NULL,
  warnSE = FALSE
)
```

### **Arguments**

xdat	a numeric vector of data to be fitted.
start	named list of starting values
method	string indicating the outer optimization routine for the augmented Lagrangian. One of $nlminb$ or BFGS.
show	logical; if TRUE (the default), print details of the fit.
fpar	a named list with optional fixed components loc, scale and shape
warnSE	logical; if TRUE, a warning is printed if the standard errors cannot be returned from the observed information matrix when the shape is less than -0.5.

### Value

a list containing the following components:

- estimate a vector containing the maximum likelihood estimates.
- std.err a vector containing the standard errors.
- vcov the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- method the method used to fit the parameter.
- nllh the negative log-likelihood evaluated at the parameter estimate.
- convergence components taken from the list returned by auglag. Values other than 0 indicate that the algorithm likely did not converge.
- counts components taken from the list returned by auglag.
- · xdat vector of data

fit.gpd 29

### **Examples**

```
xdat <- mev::rgev(n = 100)
fit.gev(xdat, show = TRUE)
# Example with fixed parameter
fit.gev(xdat, show = TRUE, fpar = list(shape = 0))</pre>
```

fit.gpd

Maximum likelihood estimation for the generalized Pareto distribution

### **Description**

Numerical optimization of the generalized Pareto distribution for data exceeding threshold. This function returns an object of class mev\_gpd, with default methods for printing and quantile-quantile plots.

### Usage

```
fit.gpd(
  xdat,
  threshold = 0,
  method = "Grimshaw",
  show = FALSE,
  MCMC = NULL,
  k = 4,
  tol = 1e-08,
  fpar = NULL,
  warnSE = FALSE
)
```

### **Arguments**

xdat a numeric vector of data to be fitted.

threshold the chosen threshold.

method the method to be used. See **Details**. Can be abbreviated. show logical; if TRUE (the default), print details of the fit.

MCMC NULL for frequentist estimates, otherwise a boolean or a list with parameters

passed. If TRUE, runs a Metropolis-Hastings sampler to get posterior mean estimates. Can be used to pass arguments niter, burnin and thin to the sampler

as a list.

k bound on the influence function (method = "obre"); the constant k is a robust-

ness parameter (higher bounds are more efficient, low bounds are more robust).

Default to 4, must be larger than  $\sqrt{2}$ .

tol numerical tolerance for OBRE weights iterations (method = "obre"). Default

to 1e-8.

fpar a named list with fixed parameters, either scale or shape

warnSE logical; if TRUE, a warning is printed if the standard errors cannot be returned

from the observed information matrix when the shape is less than -0.5.

30 fit.gpd

#### **Details**

The default method is 'Grimshaw', which maximizes the profile likelihood for the ratio scale/shape. Other options include 'obre' for optimal *B*-robust estimator of the parameter of Dupuis (1998), vanilla maximization of the log-likelihood using constrained optimization routine 'auglag', 1-dimensional optimization of the profile likelihood using nlm and optim. Method 'ismev' performs the two-dimensional optimization routine gpd. fit from the ismev library, with in addition the algebraic gradient. The approximate Bayesian methods ('zs' and 'zhang') are extracted respectively from Zhang and Stephens (2009) and Zhang (2010) and consists of a approximate posterior mean calculated via importance sampling assuming a GPD prior is placed on the parameter of the profile likelihood.

#### Value

If method is neither 'zs' nor 'zhang', a list containing the following components:

- estimate a vector containing the scale and shape parameters (optimized and fixed).
- std.err a vector containing the standard errors. For method = "obre", these are Huber's robust standard errors.
- vcov the variance covariance matrix, obtained as the numerical inverse of the observed information matrix. For method = "obre", this is the sandwich Godambe matrix inverse.
- threshold the threshold.
- method the method used to fit the parameter. See details.
- nllh the negative log-likelihood evaluated at the parameter estimate.
- nat number of points lying above the threshold.
- pat proportion of points lying above the threshold.
- convergence components taken from the list returned by optim. Values other than 0 indicate that the algorithm likely did not converge (in particular 1 and 50).
- counts components taken from the list returned by optim.
- exceedances excess over the threshold.

Additionally, if method = "obre", a vector of OBRE weights.

Otherwise, a list containing

- threshold the threshold.
- method the method used to fit the parameter. See Details.
- nat number of points lying above the threshold.
- pat proportion of points lying above the threshold.
- approx.mean a vector containing containing the approximate posterior mean estimates.

and in addition if MCMC is neither FALSE, nor NULL

- post.mean a vector containing the posterior mean estimates.
- post.se a vector containing the posterior standard error estimates.
- accept.rate proportion of points lying above the threshold.
- niter length of resulting Markov Chain
- burnin amount of discarded iterations at start, capped at 10000.
- · thin thinning integer parameter describing

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#### Note

Some of the internal functions (which are hidden from the user) allow for modelling of the parameters using covariates. This is not currently implemented within gp.fit, but users can call internal functions should they wish to use these features.

#### Author(s)

Scott D. Grimshaw for the Grimshaw option. Paul J. Northrop and Claire L. Coleman for the methods optim, nlm and ismev. J. Zhang and Michael A. Stephens (2009) and Zhang (2010) for the zs and zhang approximate methods and L. Belzile for methods auglag and obre, the wrapper and MCMC samplers.

If show = TRUE, the optimal B robust estimated weights for the largest observations are printed alongside with the p-value of the latter, obtained from the empirical distribution of the weights. This diagnostic can be used to guide threshold selection: small weights for the r-largest order statistics indicate that the robust fit is driven by the lower tail and that the threshold should perhaps be increased.

#### References

Davison, A.C. (1984). Modelling excesses over high thresholds, with an application, in *Statistical extremes and applications*, J. Tiago de Oliveira (editor), D. Reidel Publishing Co., 461–482.

Grimshaw, S.D. (1993). Computing Maximum Likelihood Estimates for the Generalized Pareto Distribution, *Technometrics*, **35**(2), 185–191.

Northrop, P.J. and C. L. Coleman (2014). Improved threshold diagnostic plots for extreme value analyses, *Extremes*, **17**(2), 289–303.

Zhang, J. (2010). Improving on estimation for the generalized Pareto distribution, *Technometrics* **52**(3), 335–339.

Zhang, J. and M. A. Stephens (2009). A new and efficient estimation method for the generalized Pareto distribution. *Technometrics* **51**(3), 316–325.

Dupuis, D.J. (1998). Exceedances over High Thresholds: A Guide to Threshold Selection, *Extremes*, **1**(3), 251–261.

#### See Also

```
fpot and gpd.fit
```

### **Examples**

```
data(eskrain)
fit.gpd(eskrain, threshold = 35, method = 'Grimshaw', show = TRUE)
fit.gpd(eskrain, threshold = 30, method = 'zs', show = TRUE)
```

32 fit.pp

fit.pp

Maximum likelihood estimation of the point process of extremes

### **Description**

Data above threshold is modelled using the limiting point process of extremes.

# Usage

```
fit.pp(
  xdat,
  threshold = 0,
  npp = 1,
  np = NULL,
  method = c("nlminb", "BFGS"),
  start = NULL,
  show = FALSE,
  fpar = NULL,
  warnSE = FALSE
)
```

### Arguments

xdat a numeric vector of data to be fitted.

threshold the chosen threshold.

npp number of observation per period. See **Details** 

np number of periods of data, if xdat only contains exceedances.

method the method to be used. See **Details**. Can be abbreviated.

start named list of starting values

show logical; if TRUE (the default), print details of the fit.

fpar a named list with optional fixed components loc, scale and shape

warnSE logical; if TRUE, a warning is printed if the standard errors cannot be returned

from the observed information matrix when the shape is less than -0.5.

#### **Details**

The parameter npp controls the frequency of observations. If data are recorded on a daily basis, using a value of npp = 365.25 yields location and scale parameters that correspond to those of the generalized extreme value distribution fitted to block maxima.

fit.rho 33

#### Value

a list containing the following components:

- estimate a vector containing all parameters (optimized and fixed).
- std.err a vector containing the standard errors.
- vcov the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- threshold the threshold.
- method the method used to fit the parameter. See details.
- nllh the negative log-likelihood evaluated at the parameter estimate.
- nat number of points lying above the threshold.
- pat proportion of points lying above the threshold.
- convergence components taken from the list returned by optim. Values other than 0 indicate that the algorithm likely did not converge (in particular 1 and 50).
- counts components taken from the list returned by optim.

#### References

Coles, S. (2001), An introduction to statistical modelling of extreme values. Springer: London, 208p.

### **Examples**

```
data(eskrain)
pp_mle <- fit.pp(eskrain, threshold = 30, np = 6201)
plot(pp_mle)</pre>
```

fit.rho

Estimator of the second order tail index parameter

### **Description**

Estimator of the second order tail index parameter

### Usage

```
fit.rho(xdat, k, method = c("fagh", "dk", "ghp", "gbw"), ...)
```

# **Arguments**

xdat vector of positive observations

k number of highest order statistics to use for estimation

method string for the estimator

. . . additional arguments passed to individual routinescurrently ignored.

34 fit.rlarg

### **Examples**

```
# Example with rho = -0.2
n <- 1000
xdat <- mev::rgp(n = n, shape = 0.2)
kmin <- floor(n^0.995)
kmax <- ceiling(n^0.999)
rho_est <- fit.rho(
    xdat = xdat,
    k = n - kmin:kmax)
rho_med <- mean(rho_est$rho)</pre>
```

fit.rlarg

Maximum likelihood estimates of point process for the r-largest observations

### **Description**

This uses a constrained optimization routine to return the maximum likelihood estimate based on an n by r matrix of observations. Observations should be ordered, i.e., the r-largest should be in the last column.

### Usage

```
fit.rlarg(
  xdat,
  start = NULL,
  method = c("nlminb", "BFGS"),
  show = FALSE,
  fpar = NULL,
  warnSE = FALSE
)
```

### **Arguments**

xdat a numeric vector of data to be fitted.

start named list of starting values

method the method to be used. See **Details**. Can be abbreviated.

show logical; if TRUE (the default), print details of the fit.

fpar a named list with fixed parameters, either scale or shape

warnSE logical; if TRUE, a warning is printed if the standard errors cannot be returned

from the observed information matrix when the shape is less than -0.5.

fit.shape 35

#### Value

a list containing the following components:

- estimate a vector containing all the maximum likelihood estimates.
- std.err a vector containing the standard errors.
- vcov the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- method the method used to fit the parameter.
- nllh the negative log-likelihood evaluated at the parameter estimate.
- convergence components taken from the list returned by auglag. Values other than 0 indicate that the algorithm likely did not converge.
- counts components taken from the list returned by auglag.
- xdat an n by r matrix of data

# **Examples**

```
xdat \leftarrow rrlarg(n = 10, loc = 0, scale = 1, shape = 0.1, r = 4) fit.rlarg(xdat)
```

fit.shape

Shape parameter estimates

### **Description**

Wrapper to estimate the tail index or shape parameter of an extreme value distribution. Each function has similar sets of arguments, a vector or scalar number of order statistics k and a vector of positive observations xdat. The method argument allows users to choose between different indicators, including the Hill estimator (hill, for positive observations and shape only), the moment estimator of Dekkers and de Haan (mom or dekkers), the de Vries estimator of de Haan and Peng (vries), the generalized jackknife estimator of Gomes et al. (genjack), the Beirlant, Vynckier and Teugels generalized quantile estimator (bvt or genquant), the Pickands estimator (pickands), the extreme *U*-statistics estimator of Oorschot, Segers and Zhou (osz), or the exponential rgression model of Beirlant et al. (erm).

### Usage

36 fit.wgpd

### **Arguments**

xdat vector of positive observations of length n

k number of largest order statistics

method estimation method.

... additional parameters passed to functions

#### Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar.

fit.wgpd Maximum likelihood estimation for weighted generalized Pareto distribution

### **Description**

Weighted maximum likelihood estimation, with user-specified vector of weights.

### Usage

```
fit.wgpd(xdat, threshold = 0, weightfun = Stein_weights, start = NULL, ...)
```

### **Arguments**

xdat vector of observations

threshold numeric, value of the threshold

weightfun function whose first argument is the length of the weight vector

start optional vector of scale and shape parameters for the optimization routine, de-

faults to NULL

... additional arguments passed to the weighting function weightfun

### Value

a list with components

- estimate a vector containing the scale and shape parameters (optimized and fixed).
- std.err a vector containing the standard errors.
- vcov the variance covariance matrix, obtained as the numerical inverse of the observed information matrix.
- threshold the threshold.
- method the method used to fit the parameter. See details.
- nllh the negative log-likelihood evaluated at the parameter estimate.
- nat number of points lying above the threshold.

frwind 37

- pat proportion of points lying above the threshold.
- convergence logical indicator of convergence.
- weights vector of weights for exceedances.
- exceedances excess over the threshold, sorted in decreasing order.

frwind

French wind data

## **Description**

Daily mean wind speed (in km/h) at four stations in the south of France, namely Cap Cepet (S1), Lyon St-Exupery (S2), Marseille Marignane (S3) and Montelimar (S4). The data includes observations from January 1976 until April 2023; days containing missing values are omitted.

#### **Format**

A data frame with 17209 observations and 8 variables:

date date of measurement

- S1 wind speed (in km/h) at Cap Cepet
- S2 wind speed (in km/h) at Lyon Saint-Exupery
- S3 wind speed (in km/h) at Marseille Marignane
- S4 wind speed (in km/h) at Montelimar
- H2 humidity (in percentage) at Lyon Saint-Exupery
- T2 mean temperature (in degree Celcius) at Lyon Saint-Exupery

The metadata attribute includes latitude and longitude (in degrees, minutes, seconds), altitude (in m), station name and station id.

## Source

European Climate Assessment and Dataset project https://www.ecad.eu/

## References

Klein Tank, A.M.G. and Coauthors, 2002. Daily dataset of 20th-century surface air temperature and precipitation series for the European Climate Assessment. Int. J. of Climatol., 22, 1441-1453.

### **Examples**

```
data(frwind, package = "mev")
head(frwind)
attr(frwind, which = "metadata")
```

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Magnetic storms		
	Magnetic storms	Magnetic storms

# Description

Absolute magnitude of 373 geomagnetic storms lasting more than 48h with absolute magnitude (dst) larger than 100 in 1957-2014.

#### **Format**

a vector of size 373

## Note

For a detailed article presenting the derivation of the Dst index, see http://wdc.kugi.kyoto-u.ac.jp/dstdir/dst2/onDs

### **Source**

Aki Vehtari

### References

World Data Center for Geomagnetism, Kyoto, M. Nose, T. Iyemori, M. Sugiura, T. Kamei (2015), *Geomagnetic Dst index*, <doi:10.17593/14515-74000>.

g	gev	Generalized extreme value distribution

## **Description**

Likelihood, score function and information matrix, bias, approximate ancillary statistics and sample space derivative for the generalized extreme value distribution

par	vector of loc, scale and shape
dat	sample vector
method	string indicating whether to use the expected ('exp') or the observed ('obs'-the default) information matrix.
V	vector calculated by gev. Vfun
n	sample size
р	vector of probabilities

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#### Usage

```
gev.ll(par, dat)
gev.score(par, dat)
gev.score(par, dat)
gev.infomat(par, dat, method = c('obs','exp'))
gev.retlev(par, p)
gev.bias(par, n)
gev.Fscore(par, dat, method=c('obs','exp'))
gev.Vfun(par, dat)
gev.phi(par, dat, V)
gev.dphi(par, dat, V)
```

#### **Functions**

- gev.11: log likelihood
- gev.ll.optim: negative log likelihood parametrized in terms of location, log(scale) and shape in order to perform unconstrained optimization
- gev.score: score vector
- gev.infomat: observed or expected information matrix
- gev. retlev: return level, corresponding to the (1-p)th quantile
- gev.bias: Cox-Snell first order bias
- gev.Fscore: Firth's modified score equation
- gev. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gev.phi: canonical parameter in the local exponential family approximation
- gev.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

## References

Firth, D. (1993). Bias reduction of maximum likelihood estimates, *Biometrika*, **80**(1), 27–38.

Coles, S. (2001). An Introduction to Statistical Modeling of Extreme Values, Springer, 209 p.

Cox, D. R. and E. J. Snell (1968). A general definition of residuals, *Journal of the Royal Statistical Society: Series B (Methodological)*, **30**, 248–275.

Cordeiro, G. M. and R. Klein (1994). Bias correction in ARMA models, *Statistics and Probability Letters*, **19**(3), 169–176.

40 gev.bcor

gev	ah	1	as

Asymptotic bias of block maxima for fixed sample sizes

## **Description**

Asymptotic bias of block maxima for fixed sample sizes

## Usage

```
gev.abias(shape, rho)
```

## **Arguments**

shape shape parameter

rho second-order parameter, non-positive

### Value

a vector of length three containing the bias for location, scale and shape (in this order)

#### References

Dombry, C. and A. Ferreira (2017). Maximum likelihood estimators based on the block maxima method. https://arxiv.org/abs/1705.00465

gev.bcor

Bias correction for GEV distribution

# Description

Bias corrected estimates for the generalized extreme value distribution using Firth's modified score function or implicit bias subtraction.

# Usage

```
gev.bcor(par, dat, corr = c("subtract", "firth"), method = c("obs", "exp"))
```

# Arguments

dat sample of observations

corr string indicating which correction to employ either subtract or firth

method string indicating whether to use the expected ('exp') or the observed ('obs'—

the default) information matrix. Used only if corr='firth'

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### **Details**

Method subtractsolves

$$\tilde{\boldsymbol{\theta}} = \hat{\boldsymbol{\theta}} + b(\tilde{\boldsymbol{\theta}})$$

for  $\tilde{\theta}$ , using the first order term in the bias expansion as given by gev. bias.

The alternative is to use Firth's modified score and find the root of

$$U(\tilde{\boldsymbol{\theta}}) - i(\tilde{\boldsymbol{\theta}})b(\tilde{\boldsymbol{\theta}}),$$

where U is the score vector, b is the first order bias and i is either the observed or Fisher information.

The routine uses the MLE (bias-corrected) as starting values and proceeds to find the solution using a root finding algorithm. Since the bias-correction is not valid for  $\xi < -1/3$ , any solution that is unbounded will return a vector of NA as the solution does not exist then.

### Value

vector of bias-corrected parameters

## **Examples**

```
set.seed(1)
dat <- mev::rgev(n=40, loc = 1, scale=1, shape=-0.2)
par <- mev::fit.gev(dat)$estimate
gev.bcor(par, dat, 'subtract')
gev.bcor(par, dat, 'firth') #observed information
gev.bcor(par, dat, 'firth','exp')</pre>
```

gev.mle

Generalized extreme value maximum likelihood estimates for various quantities of interest

### **Description**

This function calls the fit.gev routine on the sample of block maxima and returns maximum likelihood estimates for all quantities of interest, including location, scale and shape parameters, quantiles and mean and quantiles of maxima of N blocks.

## Usage

```
gev.mle(
   xdat,
   args = c("loc", "scale", "shape", "quant", "Nmean", "Nquant"),
   N,
   p,
   q
)
```

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## **Arguments**

values for.  N size of block over which to take maxima. Required only for args Nmean and Nquant.  p tail probability. Required only for arg quant.	xdat	sample vector of maxima
Nquant. p tail probability. Required only for arg quant.	args	vector of strings indicating which arguments to return the maximum likelihood values for.
	N	
q level of quantile for maxima of N exceedances. Required only for args Nquant.	p	tail probability. Required only for arg quant.
	q	level of quantile for maxima of N exceedances. Required only for $\operatorname{\sf args}$ Nquant.

## Value

named vector with maximum likelihood estimated parameter values for arguments args

# Examples

```
dat <- mev::rgev(n = 100, shape = 0.2)
gev.mle(xdat = dat, N = 100, p = 0.01, q = 0.5)</pre>
```

gev.Nyr

N-year return levels, median and mean estimate

# Description

N-year return levels, median and mean estimate

# Usage

```
gev.Nyr(par, nobs, N, type = c("retlev", "median", "mean"), p = 1/N)
```

# Arguments

par	vector of location, scale and shape parameters for the GEV distribution
nobs	integer number of observation on which the fit is based
N	integer number of observations for return level. See Details
type	string indicating the statistic to be calculated (can be abbreviated).
р	probability indicating the return level, corresponding to the quantile at 1-1/p

## **Details**

If there are  $n_y$  observations per year, the L-year return level is obtained by taking N equal to  $n_yL$ .

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## Value

a list with components

- est point estimate
- var variance estimate based on delta-method
- type statistic

gev.pll

Profile log-likelihood for the generalized extreme value distribution

# Description

This function calculates the profile likelihood along with two small-sample corrections based on Severini's (1999) empirical covariance and the Fraser and Reid tangent exponential model approximation.

# Usage

```
gev.pll(
   psi,
   param = c("loc", "scale", "shape", "quant", "Nmean", "Nquant"),
   mod = "profile",
   dat,
   N = NULL,
   p = NULL,
   q = NULL,
   correction = TRUE,
   plot = TRUE,
   ...
)
```

psi	parameter vector over which to profile (unidimensional)
param	string indicating the parameter to profile over
mod	string indicating the model, one of profile, tem or modif. See <b>Details</b> .
dat	sample vector
N	size of block over which to take maxima. Required only for param $\mbox{\sc Nmean}$ and $\mbox{\sc Nquant.}$
р	tail probability. Required only for param quant.
q	probability level of quantile. Required only for param Nquant.
correction	logical indicating whether to use ${\tt spline.corr}$ to smooth the tem approximation.
plot	logical; should the profile likelihood be displayed? Default to TRUE
	additional arguments such as output from call to Vfun if mode='tem'.

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#### **Details**

The two additional mod available are tem, the tangent exponential model (TEM) approximation and modif for the penalized profile likelihood based on  $p^*$  approximation proposed by Severini. For the latter, the penalization is based on the TEM or an empirical covariance adjustment term.

#### Value

### a list with components

- mle: maximum likelihood estimate
- psi.max: maximum profile likelihood estimate
- param: string indicating the parameter to profile over
- std.error: standard error of psi.max
- psi: vector of parameter  $\psi$  given in psi
- pll: values of the profile log likelihood at psi
- · maxpl1: value of maximum profile log likelihood

### In addition, if mod includes tem

- ullet normal: maximum likelihood estimate and standard error of the interest parameter  $\psi$
- r: values of likelihood root corresponding to  $\psi$
- q: vector of likelihood modifications
- rstar: modified likelihood root vector
- rstar.old: uncorrected modified likelihood root vector
- tem.psimax: maximum of the tangent exponential model likelihood

## In addition, if mod includes modif

- tem.mle: maximum of tangent exponential modified profile log likelihood
- tem.profll: values of the modified profile log likelihood at psi
- tem.maxpll: value of maximum modified profile log likelihood
- empcov.mle: maximum of Severini's empirical covariance modified profile log likelihood
- empcov.profll: values of the modified profile log likelihood at psi
- empcov.maxpll: value of maximum modified profile log likelihood

### References

Fraser, D. A. S., Reid, N. and Wu, J. (1999), A simple general formula for tail probabilities for frequentist and Bayesian inference. *Biometrika*, **86**(2), 249–264.

Severini, T. (2000) Likelihood Methods in Statistics. Oxford University Press. ISBN 9780198506508.

Brazzale, A. R., Davison, A. C. and Reid, N. (2007) Applied asymptotics: case studies in small-sample statistics. Cambridge University Press, Cambridge. ISBN 978-0-521-84703-2

gev.tem 45

## **Examples**

```
## Not run:
set.seed(123)
dat <- rgev(n = 100, loc = 0, scale = 2, shape = 0.3)
gev.pll(psi = seq(0,0.5, length = 50), param = 'shape', dat = dat)
gev.pll(psi = seq(-1.5, 1.5, length = 50), param = 'loc', dat = dat)
gev.pll(psi = seq(10, 40, length = 50), param = 'quant', dat = dat, p = 0.01)
gev.pll(psi = seq(12, 100, length = 50), param = 'Nmean', N = 100, dat = dat)
gev.pll(psi = seq(12, 90, length = 50), param = 'Nquant', N = 100, dat = dat, q = 0.5)
## End(Not run)</pre>
```

gev.tem

Tangent exponential model approximation for the GEV distribution

## Description

The function gev.tem provides a tangent exponential model (TEM) approximation for higher order likelihood inference for a scalar parameter for the generalized extreme value distribution. Options include location scale and shape parameters as well as value-at-risk (or return levels). The function attempts to find good values for psi that will cover the range of options, but the fail may fit and return an error.

## Usage

```
gev.tem(
  param = c("loc", "scale", "shape", "quant", "Nmean", "Nquant"),
  dat,
  psi = NULL,
  p = NULL,
  q = 0.5,
  N = NULL,
  n.psi = 50,
  plot = TRUE,
  correction = TRUE
)
```

param	parameter over which to profile
dat	sample vector for the GEV distribution
psi	scalar or ordered vector of values for the interest parameter. If NULL (default), a grid of values centered at the MLE is selected $$
p	tail probability for the (1-p)th quantile (return levels). Required only if $param = retlev'$
q	probability level of quantile. Required only for param Nquant.

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N	size of block over which to take maxima. Required only for param $\mbox{\sc Nmean}$ and $\mbox{\sc Nquant.}$
n.psi	number of values of psi at which the likelihood is computed, if psi is not supplied (NULL). Odd values are more prone to give rise to numerical instabilities near the MLE. If psi is a vector of length 2 and n.psi is greater than 2, these are taken to be endpoints of the sequence.
plot	logical indicating whether plot.fr should be called upon exit
correction	logical indicating whether spline.corr should be called.

#### Value

an invisible object of class fr (see tem in package hoa) with elements

- normal: maximum likelihood estimate and standard error of the interest parameter  $\psi$
- par.hat: maximum likelihood estimates
- par.hat.se: standard errors of maximum likelihood estimates
- th. rest: estimated maximum profile likelihood at  $(\psi, \hat{\lambda})$
- r: values of likelihood root corresponding to  $\psi$
- psi: vector of interest parameter
- q: vector of likelihood modifications
- rstar: modified likelihood root vector
- rstar.old: uncorrected modified likelihood root vector
- param: parameter

# Author(s)

Leo Belzile

## **Examples**

```
## Not run:
set.seed(1234)
dat <- rgev(n = 40, loc = 0, scale = 2, shape = -0.1)
gev.tem('shape', dat = dat, plot = TRUE)
gev.tem('quant', dat = dat, p = 0.01, plot = TRUE)
gev.tem('scale', psi = seq(1, 4, by = 0.1), dat = dat, plot = TRUE)
dat <- rgev(n = 40, loc = 0, scale = 2, shape = 0.2)
gev.tem('loc', dat = dat, plot = TRUE)
gev.tem('Nmean', dat = dat, p = 0.01, N=100, plot = TRUE)
gev.tem('Nquant', dat = dat, q = 0.5, N=100, plot = TRUE)
## End(Not run)</pre>
```

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Generalized extreme value distribution

## Description

Density function, distribution function, quantile function and random number generation for the generalized extreme value distribution.

### Usage

```
qgev(p, loc = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
rgev(n, loc = 0, scale = 1, shape = 0)
dgev(x, loc = 0, scale = 1, shape = 0, log = FALSE)
pgev(q, loc = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
```

## **Arguments**

p	vector of probabilities
loc	scalar or vector of location parameters whose length matches that of the input
scale	scalar or vector of positive scale parameters whose length matches that of the input
shape	scalar shape parameter
lower.tail	logical; if TRUE (default), returns the distribution function, otherwise the survival function $$
n	scalar number of observations
x, q	vector of quantiles
log, log.p	logical; if TRUE, probabilities $p$ are given as $\log(p)$ .

### **Details**

The distribution function of a GEV distribution with parameters  $loc = \mu$ ,  $scale = \sigma$  and  $shape = \xi$  is

$$F(x) = \exp\{-[1 + \xi(x - \mu)/\sigma]^{-1/\xi}\}\$$

for  $1 + \xi(x - \mu)/\sigma > 0$ . If  $\xi = 0$  the distribution function is defined as the limit as  $\xi$  tends to zero.

The quantile function, when evaluated at zero or one, returns the lower and upper endpoint, whether the latter is finite or not.

## Author(s)

Leo Belzile, with code adapted from Paul Northrop

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#### References

Jenkinson, A. F. (1955) The frequency distribution of the annual maximum (or minimum) of meteorological elements. *Quart. J. R. Met. Soc.*, **81**, 158-171. Chapter 3: doi:10.1002/qj.49708134804

Coles, S. G. (2001) An Introduction to Statistical Modeling of Extreme Values, Springer-Verlag, London. doi:10.1007/9781447136750 3

gevN Generalized extreme value distribution (quantile/mean of N-block maxima parametrization)

### **Description**

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized extreme value distribution parametrized in terms of the quantiles/mean of N-block maxima parametrization z, scale and shape.

#### **Arguments**

par	vector of loc, quantile/mean of N-block maximum and shape
dat	sample vector
٧	vector calculated by gevN. Vfun
q	probability, corresponding to $q$ th quantile of the N-block maximum
qty	string indicating whether to calculate the q quantile or the mean

### Usage

```
gevN.ll(par, dat, N, q, qty = c('mean', 'quantile'))
gevN.ll.optim(par, dat, N, q = 0.5, qty = c('mean', 'quantile'))
gevN.score(par, dat, N, q = 0.5, qty = c('mean', 'quantile'))
gevN.infomat(par, dat, qty = c('mean', 'quantile'), method = c('obs', 'exp'), N, q = 0.5, nobs = length(of gevN.Vfun(par, dat, N, q = 0.5, qty = c('mean', 'quantile'))
gevN.phi(par, dat, N, q = 0.5, qty = c('mean', 'quantile'), V)
gevN.dphi(par, dat, N, q = 0.5, qty = c('mean', 'quantile'), V)
```

## **Functions**

- gevN.11: log likelihood
- gevN. score: score vector
- gevN. infomat: expected and observed information matrix
- gevN. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gevN.phi: canonical parameter in the local exponential family approximation
- gevN.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

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#### Author(s)

Leo Belzile

gevr Generalized extreme value distribution (return level parametrization)

## **Description**

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized extreme value distribution parametrized in terms of the return level z, scale and shape.

## **Arguments**

```
par vector of retlev, scale and shape dat sample vector  \begin{array}{ll} \text{p} & \text{tail probability, corresponding to } (1-p) \text{th quantile for } z \\ \text{method} & \text{string indicating whether to use the expected ('exp') or the observed ('obs'-the default) information matrix.} \\ \text{nobs} & \text{number of observations} \\ \text{V} & \text{vector calculated by gevr.Vfun} \\ \end{array}
```

### Usage

```
gevr.ll(par, dat, p)
gevr.ll.optim(par, dat, p)
gevr.score(par, dat, p)
gevr.infomat(par, dat, p, method = c('obs', 'exp'), nobs = length(dat))
gevr.Vfun(par, dat, p)
gevr.phi(par, dat, p, V)
gevr.dphi(par, dat, p, V)
```

## **Functions**

- gevr.11: log likelihood
- gevr.ll.optim: negative log likelihood parametrized in terms of return levels, log(scale) and shape in order to perform unconstrained optimization
- gevr.score: score vector
- gevr.infomat: observed information matrix
- · gevr. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gevr.phi: canonical parameter in the local exponential family approximation
- gevr.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

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## Author(s)

Leo Belzile

## **Description**

Likelihood, score function and information matrix, bias, approximate ancillary statistics and sample space derivative for the generalized Pareto distribution

### **Arguments**

```
par vector of scale and shape

dat sample vector

tol numerical tolerance for the exponential model

method string indicating whether to use the expected ('exp') or the observed ('obs' -
the default) information matrix.

V vector calculated by gpd. Vfun

n sample size
```

## Usage

```
gpd.ll(par, dat, tol=1e-5)
gpd.ll.optim(par, dat, tol=1e-5)
gpd.score(par, dat)
gpd.infomat(par, dat, method = c('obs','exp'))
gpd.bias(par, n)
gpd.Fscore(par, dat, method = c('obs','exp'))
gpd.Vfun(par, dat)
gpd.phi(par, dat, V)
gpd.dphi(par, dat, V)
```

#### **Functions**

- gpd.11: log likelihood
- gpd.ll.optim: negative log likelihood parametrized in terms of log(scale) and shape in order to perform unconstrained optimization
- gpd.score: score vector
- gpd.infomat: observed or expected information matrix
- gpd.bias: Cox-Snell first order bias
- gpd.Fscore: Firth's modified score equation
- gpd. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gpd.phi: canonical parameter in the local exponential family approximation
- gpd.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

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### Author(s)

Leo Belzile

### References

Firth, D. (1993). Bias reduction of maximum likelihood estimates, *Biometrika*, **80**(1), 27–38.

Coles, S. (2001). An Introduction to Statistical Modeling of Extreme Values, Springer, 209 p.

Cox, D. R. and E. J. Snell (1968). A general definition of residuals, *Journal of the Royal Statistical Society: Series B (Methodological)*, **30**, 248–275.

Cordeiro, G. M. and R. Klein (1994). Bias correction in ARMA models, *Statistics and Probability Letters*, **19**(3), 169–176.

Giles, D. E., Feng, H. and R. T. Godwin (2016). Bias-corrected maximum likelihood estimation of the parameters of the generalized Pareto distribution, *Communications in Statistics - Theory and Methods*, **45**(8), 2465–2483.

gpd.abias

Asymptotic bias of threshold exceedances for k order statistics

### **Description**

The formula given in de Haan and Ferreira, 2007 (Springer). Note that the latter differs from that found in Drees, Ferreira and de Haan.

## Usage

```
gpd.abias(shape, rho)
```

## **Arguments**

shape shape parameter

rho second-order parameter, non-positive

## Value

a vector of length containing the bias for scale and shape (in this order)

## References

Dombry, C. and A. Ferreira (2017). Maximum likelihood estimators based on the block maxima method. https://arxiv.org/abs/1705.00465

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gr	od	b	C	O	r

Bias correction for GP distribution

## **Description**

Bias corrected estimates for the generalized Pareto distribution using Firth's modified score function or implicit bias subtraction.

#### **Usage**

```
gpd.bcor(par, dat, corr = c("subtract", "firth"), method = c("obs", "exp"))
```

## **Arguments**

par parameter vector (scale, shape)

dat sample of observations

corr string indicating which correction to employ either subtract or firth

method string indicating whether to use the expected ('exp') or the observed ('obs'—

the default) information matrix. Used only if corr='firth'

#### **Details**

Method subtract solves

$$\tilde{\boldsymbol{\theta}} = \hat{\boldsymbol{\theta}} + b(\tilde{\boldsymbol{\theta}})$$

for  $\hat{\theta}$ , using the first order term in the bias expansion as given by gpd. bias.

The alternative is to use Firth's modified score and find the root of

$$U(\tilde{\boldsymbol{\theta}}) - i(\tilde{\boldsymbol{\theta}})b(\tilde{\boldsymbol{\theta}}),$$

where U is the score vector, b is the first order bias and i is either the observed or Fisher information.

The routine uses the MLE as starting value and proceeds to find the solution using a root finding algorithm. Since the bias-correction is not valid for  $\xi < -1/3$ , any solution that is unbounded will return a vector of NA as the bias correction does not exist then.

### Value

vector of bias-corrected parameters

## **Examples**

```
set.seed(1)
dat <- rgp(n=40, scale=1, shape=-0.2)
par <- gp.fit(dat, threshold=0, show=FALSE)$estimate
gpd.bcor(par,dat, 'subtract')
gpd.bcor(par,dat, 'firth') #observed information
gpd.bcor(par,dat, 'firth','exp')</pre>
```

gpd.boot 53

gpd.boot	Bootstrap approximation for generalized Pareto parameters	

#### **Description**

Given an object of class mev\_gpd, returns a matrix of parameter values to mimic the estimation uncertainty.

## Usage

```
gpd.boot(object, B = 1000L, method = c("post", "norm"))
```

### **Arguments**

object of class mev\_gpd

B number of pairs to sample

method string; one of 'norm' for the normal approximation or 'post' (default) for

posterior sampling

#### **Details**

Two options are available: a normal approximation to the scale and shape based on the maximum likelihood estimates and the observed information matrix. This method uses forward sampling to simulate from a bivariate normal distribution that satisfies the support and positivity constraints

The second approximation uses the ratio-of-uniforms method to obtain samples from the posterior distribution with uninformative priors, thus mimicking the joint distribution of maximum likelihood. The benefit of the latter is that it is more reliable in small samples and when the shape is negative.

#### Value

a matrix of size B by 2 whose columns contain scale and shape parameters

gpd.lmom	Estimation of generalized Pareto parameters via L-moments	
gpd.lmom	Estimation of generalized Pareto parameters via L-moments	

## **Description**

Given a sample of exceedances, compute the first four L-moments and use either the first two to obtain the scale and shape (default), or else use L-skewness and L-scale to compute the scale and shape of the generalized Pareto distribution

## Usage

```
gpd.lmom(xdat, thresh, sorted = FALSE, Lskew = FALSE)
```

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## **Arguments**

xdat	[numeric] vector of observations
thresh	[numeric] optional threshold argument
sorted	[logical] if TRUE, observations are sorted in increasing order
Lskew	[logical]; if TRUE, shape is obtained from L-skewness rather than first two moments.

## Value

a vector of length two with the scale and shape estimates

gpd.mle	Generalized Pareto maximum likelihood estimates for various quantities of interest

# Description

This function calls the fit.gpd routine on the sample of excesses and returns maximum likelihood estimates for all quantities of interest, including scale and shape parameters, quantiles and value-atrisk, expected shortfall and mean and quantiles of maxima of N threshold exceedances

# Usage

```
gpd.mle(
  xdat,
  args = c("scale", "shape", "quant", "VaR", "ES", "Nmean", "Nquant"),
  m,
  N,
  p,
  q
)
```

xdat	sample vector of excesses
args	vector of strings indicating which arguments to return the maximum likelihood values for
m	number of observations of interest for return levels. Required only for args values 'VaR' or 'ES' $$
N	size of block over which to take maxima. Required only for $\mathop{\hbox{args}}\nolimits$ Nmean and Nquant.
р	tail probability, equivalent to $1/m$ . Required only for args quant.
q	level of quantile for N-block maxima. Required only for args Nquant.

gpd.pll 55

## Value

named vector with maximum likelihood values for arguments args

## **Examples**

```
xdat <- mev::rgp(n = 30, shape = 0.2)
gpd.mle(xdat = xdat, N = 100, p = 0.01, q = 0.5, m = 100)
```

gpd.pll

Profile log-likelihood for the generalized Pareto distribution

## **Description**

This function calculates the (modified) profile likelihood based on the  $p^*$  formula. There are two small-sample corrections that use a proxy for  $\ell_{\lambda;\hat{\lambda}}$ , which are based on Severini's (1999) empirical covariance and the Fraser and Reid tangent exponential model approximation.

## Usage

```
gpd.pll(
    psi,
    param = c("scale", "shape", "quant", "retlev", "VaR", "ES", "Nmean", "Nquant"),
    mod = "profile",
    mle = NULL,
    dat,
    m = NULL,
    N = NULL,
    p = NULL,
    q = NULL,
    correction = TRUE,
    thresh = NULL,
    plot = TRUE,
    ...
)
```

psi	parameter vector over which to profile (unidimensional)
param	string indicating the parameter to profile over
mod	string indicating the model. See <b>Details</b> .
mle	maximum likelihood estimate in $(\psi,\xi)$ parametrization if $\psi\neq\xi$ and $(\sigma,\xi)$ otherwise (optional).
dat	sample vector of excesses, unless thresh is provided (in which case user provides original data)
m	number of observations of interest for return levels. Required only for args values 'VaR' or 'ES' $$

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N	size of block over which to take maxima. Required only for args Nmean and Nquant.
р	tail probability, equivalent to $1/m$ . Required only for args quant.
q	level of quantile for N-block maxima. Required only for args Nquant.
correction	logical indicating whether to use $spline.corr$ to smooth the tem approximation.
thresh	numerical threshold above which to fit the generalized Pareto distribution
plot	logical; should the profile likelihood be displayed? Default to TRUE
	additional arguments such as output from call to Vfun if mode='tem'.

### **Details**

The three mod available are profile (the default), tem, the tangent exponential model (TEM) approximation and modif for the penalized profile likelihood based on  $p^*$  approximation proposed by Severini. For the latter, the penalization is based on the TEM or an empirical covariance adjustment term.

#### Value

### a list with components

- mle: maximum likelihood estimate
- psi.max: maximum profile likelihood estimate
- param: string indicating the parameter to profile over
- std.error: standard error of psi.max
- psi: vector of parameter  $\psi$  given in psi
- pll: values of the profile log likelihood at psi
- maxpll: value of maximum profile log likelihood
- family: a string indicating "gpd"
- thresh: value of the threshold, by default zero

#### In addition, if mod includes tem

- ullet normal: maximum likelihood estimate and standard error of the interest parameter  $\psi$
- r: values of likelihood root corresponding to  $\psi$
- q: vector of likelihood modifications
- rstar: modified likelihood root vector
- rstar.old: uncorrected modified likelihood root vector
- tem.psimax: maximum of the tangent exponential model likelihood

#### In addition, if mod includes modif

- tem.mle: maximum of tangent exponential modified profile log likelihood
- tem. profll: values of the modified profile log likelihood at psi

gpd.tem 57

- tem.maxpll: value of maximum modified profile log likelihood
- empcov.mle: maximum of Severini's empirical covariance modified profile log likelihood
- empcov.profll: values of the modified profile log likelihood at psi
- empcov.maxpll: value of maximum modified profile log likelihood

### **Examples**

```
## Not run:
dat <- rgp(n = 100, scale = 2, shape = 0.3)
gpd.pll(psi = seq(-0.5, 1, by=0.01), param = 'shape', dat = dat)
gpd.pll(psi = seq(0.1, 5, by=0.1), param = 'scale', dat = dat)
gpd.pll(psi = seq(20, 35, by=0.1), param = 'quant', dat = dat, p = 0.01)
gpd.pll(psi = seq(20, 80, by=0.1), param = 'ES', dat = dat, m = 100)
gpd.pll(psi = seq(15, 100, by=1), param = 'Nmean', N = 100, dat = dat)
gpd.pll(psi = seq(15, 90, by=1), param = 'Nquant', N = 100, dat = dat, q = 0.5)
## End(Not run)</pre>
```

gpd.tem

Tangent exponential model approximation for the GP distribution

### **Description**

The function gpd.tem provides a tangent exponential model (TEM) approximation for higher order likelihood inference for a scalar parameter for the generalized Pareto distribution. Options include scale and shape parameters as well as value-at-risk (also referred to as quantiles, or return levels) and expected shortfall. The function attempts to find good values for psi that will cover the range of options, but the fit may fail and return an error. In such cases, the user can try to find good grid of starting values and provide them to the routine.

## Usage

```
gpd.tem(
  dat,
  param = c("scale", "shape", "quant", "VaR", "retlev", "ES", "Nmean", "Nquant"),
  psi = NULL,
  m = NULL,
  thresh = 0,
  n.psi = 50,
  N = NULL,
  p = NULL,
  q = NULL,
  plot = FALSE,
  correction = TRUE,
  ...
)
```

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## **Arguments**

dat	sample vector for the GP distribution
param	parameter over which to profile
psi	scalar or ordered vector of values for the interest parameter. If NULL (default), a grid of values centered at the MLE is selected. If psi is of length 2 and n.psi>2, it is assumed to be the minimal and maximal values at which to evaluate the profile log likelihood.
m	number of observations of interest for return levels. See <b>Details</b> . Required only for param = 'VaR' or param = 'ES'.
thresh	threshold value corresponding to the lower bound of the support or the location parameter of the generalized Pareto distribution.
n.psi	number of values of psi at which the likelihood is computed, if psi is not supplied (NULL). Odd values are more prone to give rise to numerical instabilities near the MLE
N	size of block over which to take maxima. Required only for args $\ensuremath{Nmean}$ and $\ensuremath{Nquant}.$
p	tail probability, equivalent to $1/m$ . Required only for args quant.
q	level of quantile for N-block maxima. Required only for args Nquant.
plot	logical indicating whether plot.fr should be called upon exit
correction	logical indicating whether spline.corr should be called.
• • •	additional arguments, for backward compatibility

### **Details**

As of version 1.11, this function is a wrapper around gpd.pll.

The interpretation for m is as follows: if there are on average  $m_y$  observations per year above the threshold, then  $m=Tm_y$  corresponds to T-year return level.

# Value

an invisible object of class fr (see tem in package hoa) with elements

- normal: maximum likelihood estimate and standard error of the interest parameter  $\psi$
- par.hat: maximum likelihood estimates
- par.hat.se: standard errors of maximum likelihood estimates
- th. rest: estimated maximum profile likelihood at  $(\psi, \hat{\lambda})$
- r: values of likelihood root corresponding to  $\psi$
- psi: vector of interest parameter
- q: vector of likelihood modifications
- rstar: modified likelihood root vector
- rstar.old: uncorrected modified likelihood root vector
- param: parameter

gpde 59

### Author(s)

Leo Belzile

## **Examples**

```
set.seed(123)
dat < - rgp(n = 40, scale = 1, shape = -0.1)
#with plots
m1 <- gpd.tem(param = 'shape', n.psi = 50, dat = dat, plot = TRUE)
## Not run:
m2 <- gpd.tem(param = 'scale', n.psi = 50, dat = dat)</pre>
m3 <- gpd.tem(param = 'VaR', n.psi = 50, dat = dat, m = 100)
#Providing psi
psi \leftarrow c(seq(2, 5, length = 15), seq(5, 35, length = 45))
m4 <- gpd.tem(param = 'ES', dat = dat, m = 100, psi = psi, correction = FALSE)
mev:::plot.fr(m4, which = c(2, 4))
plot(fr4 <- spline.corr(m4))</pre>
confint(m1)
confint(m4, parm = 2, warn = FALSE)
m5 <- gpd.tem(param = 'Nmean', dat = dat, N = 100, psi = psi, correction = FALSE)
m6 \leftarrow gpd.tem(param = 'Nquant', dat = dat, N = 100, q = 0.7, correction = FALSE)
## End(Not run)
```

gpde

Generalized Pareto distribution (expected shortfall parametrization)

### **Description**

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized Pareto distribution parametrized in terms of expected shortfall.

The parameter m corresponds to  $\zeta_u/(1-\alpha)$ , where  $\zeta_u$  is the rate of exceedance over the threshold u and  $\alpha$  is the percentile of the expected shortfall. Note that the actual parametrization is in terms of excess expected shortfall, meaning expected shortfall minus threshold.

par	vector of length 2 containing $e_m$ and $\xi$ , respectively the expected shortfall at probability $1/(1-\alpha)$ and the shape parameter.
dat	sample vector
m	number of observations of interest for return levels. See Details
tol	numerical tolerance for the exponential model
method	string indicating whether to use the expected ('exp') or the observed ('obs'-the default) information matrix.
nobs	number of observations
V	vector calculated by gpde. Vfun

60 gpdist

#### **Details**

The observed information matrix was calculated from the Hessian using symbolic calculus in Sage.

#### Usage

```
gpde.ll(par, dat, m, tol=1e-5)
gpde.ll.optim(par, dat, m, tol=1e-5)
gpde.score(par, dat, m)
gpde.infomat(par, dat, m, method = c('obs', 'exp'), nobs = length(dat))
gpde.Vfun(par, dat, m)
gpde.phi(par, dat, V, m)
gpde.dphi(par, dat, V, m)
```

#### **Functions**

- gpde.11: log likelihood
- gpde.ll.optim: negative log likelihood parametrized in terms of log expected shortfall and shape in order to perform unconstrained optimization
- gpde.score: score vector
- gpde.infomat: observed information matrix for GPD parametrized in terms of rate of expected shortfall and shape
- gpde. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gpde.phi: canonical parameter in the local exponential family approximation
- gpde.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

## Author(s)

Leo Belzile

gpdist

Generalized Pareto distribution

## **Description**

Density function, distribution function, quantile function and random number generation for the generalized Pareto distribution.

## Usage

```
pgp(q, loc = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)

dgp(x, loc = 0, scale = 1, shape = 0, log = FALSE)

qgp(p, loc = 0, scale = 1, shape = 0, lower.tail = TRUE)

rgp(n, loc = 0, scale = 1, shape = 0)
```

gpdN

## **Arguments**

loc location parameter.

scale scale parameter, strictly positive.

shape shape parameter.

lower.tail logical; if TRUE (default), the lower tail probability  $\Pr(X \leq x)$  is returned. log.p, log logical; if FALSE (default), values are returned on the probability scale.

x, q vector of quantilesp vector of probabilities

n scalar number of observations

#### References

Coles, S. G. (2001) An Introduction to Statistical Modeling of Extreme Values, Springer-Verlag, London. doi:10.1007/9781447136750 3

gpdN	Generalized Pareto distribution (mean of maximum of N exceedances parametrization)

## Description

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized Pareto distribution parametrized in terms of average maximum of N exceedances.

The parameter N corresponds to the number of threshold exceedances of interest over which the maxima is taken. z is the corresponding expected value of this block maxima. Note that the actual parametrization is in terms of excess expected mean, meaning expected mean minus threshold.

## **Arguments**

par	vector of length 2 containing $z$ and $\xi$ , respectively the mean excess of the maxima of N exceedances above the threshold and the shape parameter.
dat	sample vector
N	block size for threshold exceedances.
tol	numerical tolerance for the exponential model
٧	vector calculated by gpdN. Vfun

#### **Details**

The observed information matrix was calculated from the Hessian using symbolic calculus in Sage.

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## Usage

```
gpdN.ll(par, dat, N, tol=1e-5)
gpdN.score(par, dat, N)
gpdN.infomat(par, dat, N, method = c('obs', 'exp'), nobs = length(dat))
gpdN.Vfun(par, dat, N)
gpdN.phi(par, dat, N, V)
gpdN.dphi(par, dat, N, V)
```

### **Functions**

- gpdN.11: log likelihood
- gpdN. score: score vector
- gpdN.infomat: observed information matrix for GP parametrized in terms of mean of the maximum of N exceedances and shape
- gpdN. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gpdN.phi: canonical parameter in the local exponential family approximation
- gpdN.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

### Author(s)

Leo Belzile

gpdr Generalized Pareto distribution (return level parametrization)	
---	--

### **Description**

Likelihood, score function and information matrix, approximate ancillary statistics and sample space derivative for the generalized Pareto distribution parametrized in terms of return levels.

par	vector of length 2 containing $y_m$ and $\xi$ , respectively the $m$ -year return level and the shape parameter.
dat	sample vector
m	number of observations of interest for return levels. See Details
tol	numerical tolerance for the exponential model
method	string indicating whether to use the expected ('exp') or the observed ('obs'-the default) information matrix.
nobs	number of observations
V	vector calculated by gpdr. Vfun

ibvpot 63

### **Details**

The observed information matrix was calculated from the Hessian using symbolic calculus in Sage.

The interpretation for m is as follows: if there are on average  $m_y$  observations per year above the threshold, then  $m = Tm_y$  corresponds to T-year return level.

#### Usage

```
gpdr.ll(par, dat, m, tol=1e-5)
gpdr.ll.optim(par, dat, m, tol=1e-5)
gpdr.score(par, dat, m)
gpdr.infomat(par, dat, m, method = c('obs', 'exp'), nobs = length(dat))
gpdr.Vfun(par, dat, m)
gpdr.phi(par, V, dat, m)
gpdr.dphi(par, V, dat, m)
```

## **Functions**

- gpdr.11: log likelihood
- gpdr.ll.optim: negative log likelihood parametrized in terms of log(scale) and shape in order to perform unconstrained optimization
- gpdr.score: score vector
- $\bullet$  gpdr. infomat: observed information matrix for GPD parametrized in terms of rate of m-year return level and shape
- gpdr. Vfun: vector implementing conditioning on approximate ancillary statistics for the TEM
- gpdr.phi: canonical parameter in the local exponential family approximation
- gpdr.dphi: derivative matrix of the canonical parameter in the local exponential family approximation

#### Author(s)

Leo Belzile

ibvpot

Interpret bivariate threshold exceedance models

## **Description**

This is an adaptation of the evir package interpret.gpdbiv function. interpret.fbvpot deals with the output of a call to fbvpot from the **evd** and to handle families other than the logistic distribution. The likelihood derivation comes from expression 2.10 in Smith et al. (1997).

## Usage

```
ibvpot(fitted, q, silent = FALSE)
```

ibvpot

## **Arguments**

fitted	the output of fbvpot or a list. See Details.
q	a vector of quantiles to consider, on the data scale. Must be greater than the thresholds.
silent	boolean; whether to print the interpretation of the result. Default to FALSE.

#### **Details**

The list fitted must contain

- model a string; see byevd from package evd for options
- param a named vector containing the parameters of the model, as well as parameters scale1, shape1,scale2 and shape2, corresponding to marginal GPD parameters.
- threshold a vector of length 2 containing the two thresholds.
- pat the proportion of observations above the corresponding threshold

#### Value

an invisible numeric vector containing marginal, joint and conditional exceedance probabilities.

## Author(s)

Leo Belzile, adapting original S code by Alexander McNeil

#### References

Smith, Tawn and Coles (1997), Markov chain models for threshold exceedances. *Biometrika*, **84**(2), 249–268.

#### See Also

```
interpret.gpdbiv in package evir
```

### **Examples**

```
if (requireNamespace("evd", quietly = TRUE)) { y \leftarrow rgp(1000,1,1,1) x \leftarrow y*rmevspec(n=1000,d=2,sigma=cbind(c(0,0.5),c(0.5,0)), model='hr') mod <- evd::fbvpot(x, threshold = c(1,1), model = 'hr', likelihood = 'censored') ibvpot(mod, c(20,20)) }
```

kjtail 65

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Estimators of the tail coefficient

# Description

Estimators proposed by Krupskii and Joe under second order expansion for the coefficient of tail dependence  $\eta$  and the joint tail orthant probability

# Usage

```
kjtail(
  xdat,
  qlev,
  ptail = NULL,
  mqu = NULL,
  type = 1,
  ties.method = eval(formals(rank)$ties.method),
  ...
)
```

## **Arguments**

xdat	a matrix of observations
qlev	vector of quantile levels
ptail	tail probability smaller than qlev. Default to NULL
mqu	marginal quantile levels for semiparametric estimation; data above this are modelled using a generalized Pareto distribution. If NULL, empirical estimation is used throughout
type	integer indicating the estimator type
ties.method	method for handling of ties in rank transformation
	additional arguments, for backward compatibility

#### Value

## a list with elements

- p quantile level for estimation
- eta matrix of estimated coefficient of tail dependence  $\eta$ , and standard errors
- k1 parameter of the tail expansion
- pat proportion of observations above the threshold
- lambda tail dependence coefficient (sic)
- tailprob tail probability, if ptail is provided

66 leedspollution

#### Note

EXPERIMENTAL. The numerical optimization of the likelihood surface is difficult, as the function is ill-behaved. Visual inspection of estimates is often necessary to check for possible lack of convergence.

## **Examples**

```
d <- 2
rho <- 0.9
Sigma <- matrix(rho, d, d) + diag(1 - rho, d)
eta_true <- 1/sum(Sigma)
data <- rmnorm(
    n = 1e4,
    mu = rep(0, d),
    Sigma = Sigma)
q <- seq(0.95, 0.995, by = 0.005)
kj <- kjtail(xdat = data, qlev = q)
plot(kj)
abline(h = (1+rho)/2, col = 2)</pre>
```

leedspollution

Leeds air pollution

### **Description**

Daily maximum data (hourly for PM10) on air pollution for the Leeds Centre station in Yorkshire and Humberside station. The data goes from January 1st, 1993, until December 31st, 2024. Data show seasonality and there are some outliers. From December 2nd, 2008 onwards, particulate matters (PM10 and PM2.5) are measured using a tapered element oscillating microbalance (TEOM) and Filter Dynamics Measurement System (FDMS). The data for PM2.5 is missing before the change of instrumentation. A total of 231 daily measurements with only missing values were removed during preprocessing.

#### Usage

leedspollution

#### **Format**

A data frame with 11455 rows and 8 variables:

date [character] a date with format yyy-mm-dd

03 [integer] ozone (in nanograms per cubic meter)

NO [integer] nitrogen oxyde (in nanograms per cubic meter)

CO [double] carbon monoxyde (in micrograms per cubic meter)

NO2 nitrogen dioxyde (in nanograms per cubic meter)

S02 sulphur dioxide (in nanograms per cubic meter)

PM10 [integer] particulate matter 10, (in nanograms per cubic meter)

PM2.5 [integer] particulate matter 2.5, (in nanograms per cubic meter)

likmgp 67

## Source

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likmgp

Likelihood for multivariate peaks over threshold models

## **Description**

Likelihood for the various parametric limiting models over region determined by

$$\{y \in F: \max_{j=1}^{D} \sigma_{j} \frac{y_{j}^{\xi} - 1}{\xi_{j}} + \mu_{j} > u\};$$

where  $\mu$  is loc,  $\sigma$  is scale and  $\xi$  is shape.

# Usage

```
likmgp(
    dat,
    thresh,
    loc,
    scale,
    shape,
    par,
    model = c("log", "br", "xstud"),
    likt = c("mgp", "pois", "binom"),
    lambdau = 1,
    ...
)
```

dat	matrix of observations
thresh	functional threshold for the maximum
loc	vector of location parameter for the marginal generalized Pareto distribution
scale	vector of scale parameter for the marginal generalized Pareto distribution
shape	vector of shape parameter for the marginal generalized Pareto distribution
par	list of parameters: alpha for the logistic model, Lambda for the Brown–Resnick model or else Sigma and df for the extremal Student.
model	string indicating the model family, one of "log", "neglog", "br" or "xstud"
likt	string indicating the type of likelihood, with an additional contribution for the non-exceeding components: one of "mgp", "binom" and "pois".
lambdau	vector of marginal rate of marginal threshold exceedance.
	additional arguments (see Details)

68 maiquetia

### **Details**

Optional arguments can be passed to the function via . . .

- cl cluster instance created by makeCluster (default to NULL)
- ncors number of cores for parallel computing of the likelihood
- mmax maximum per column
- B1 number of replicates for quasi Monte Carlo integral for the exponent measure
- genvec1 generating vector for the quasi Monte Carlo routine (exponent measure), associated with B1

#### Value

the value of the log-likelihood with attributes expme, giving the exponent measure

### Note

The location and scale parameters are not identifiable unless one of them is fixed.

maiquetia

Maiquetia Daily Rainfall

# Description

Daily cumulated rainfall (in mm) at Maiquetia airport, Venezuela. The observations cover the period from January 1961 to December 1999. The original series had missing days in February 1996 (during which there were 2 days with 1hr each of light rain) and January 1998 (no rain). These were replaced by zeros.

## **Format**

a vector of size 14244 containing daily rainfall (in mm),

## Source

J.R. Cordova and M. González, accessed 25.11.2018 from <a href="https://rss.onlinelibrary.wiley.com/hub/journal/14679876/series-c-datasets">https://rss.onlinelibrary.wiley.com/hub/journal/14679876/series-c-datasets</a>

### References

Coles, S. and L.R. Pericchi (2003). Anticipating Catastrophes through Extreme Value Modelling, *Applied Statistics*, **52**(4), 405-416.

Coles, S., Pericchi L.R. and S. Sisson (2003). A fully probabilistic approach to extreme rainfall modeling, *Journal of Hydrology*, **273**, 35-50.

maxstable 69

### **Examples**

```
## Not run:
data(maiquetia, package = "mev")
day <- seq.Date(from = as.Date("1961-01-01"), to = as.Date("1999-12-31"), by = "day")
nzrain <- maiquetia[substr(day, 1, 4) < 1999 & maiquetia > 0]
fit.gpd(nzrain, threshold = 30, show = TRUE)
## End(Not run)
```

maxstable

Transform arguments using max stability

## Description

Given a vector of location, scale and shape parameters, compute the corresponding parameters for block of size m assuming a generalized extreme value distribution.

## Usage

```
maxstable(pars, m = 1L, inverse = FALSE)
```

## **Arguments**

pars vector of location, scale and shape parameters

m [integer] block size

inverse [logical] whether to compute the parameters for the inverse relationship (defaults

to FALSE)

## **Examples**

```
maxstable(pars = maxstable(pars = c(1,2,0), m = 10), m = 10, inv = TRUE) maxstable(pars = maxstable(pars = c(1,2,0.1), m = 5), m = 1/5)
```

nidd

River Nidd Flow

### Description

The data consists of exceedances over the threshold 65 cubic meter per second of the River Nidd at Hunsingore Weir, for 35 years of data between 1934 and 1969.

## Format

a vector of size 154

70 nutrients

### **Source**

Natural Environment Research Council (1975). Flood Studies Report, volume 4. pp. 235-236.

#### References

Davison, A.C. and R.L. Smith (1990). Models for Exceedances over High Thresholds (with discussion), *Journal of the Royal Statistical Society. Series B (Methodological)*, **52**(3), 393–442.

### See Also

nidd. thresh from the evir package

nutrients

Nutrient data

# **Description**

Interview component of survey 'What we eat in America'. These are extracted from the 2015–2016 National Health and Nutrition Examination Survey (NHANES, <a href="https://wwwn.cdc.gov/nchs/nhanes/Default.aspx">https://wwwn.cdc.gov/nchs/nhanes/Default.aspx</a>) report and consist of the total nutrients for all food and beverage intake ingested over a 24 hours period.

## Usage

nutrients

## **Format**

```
A data frame with 9544 rows and 38 variables:
```

```
prot proteins (in grams)
```

carb carbonhydrate (in gram)

sugr total sugars (in gram)

fibe dietary fibers (in grams)

tfat total fat (in grams)

sfat saturated fat (in grams)

mfat monounsaturated fat (in grams)

pfat polyunsaturated fat (in grams)

chol cholesterol (in milligrams)

atoc vitamin E as alpha-tocopherol (in milligrams)

ret retinol (in micrograms)

vara Vitamin A as retinol activity equivalents (in micrograms).

acar alpha-carotene (in micrograms)

bcar beta-carotene (in micrograms)

nutrients 71

```
cryp beta-cryptoxanthin (in micrograms)
lyco lycopene (in micrograms)
1z lutein and zeaxanthin (in micrograms).
vb1 thiamin (vitamin B1, in milligrams)
vb2 riboflavin (vitamin B2, in milligrams)
niac niacin (in milligrams)
vb6 vitamin B5 (in milligrams)
fola total folate (in micrograms)
fa folic acid (in micrograms)
ff food folate (in micrograms)
chl total choline (in milligrams)
vb12 vitamin B12 (in micrograms)
vc vitamin C (in milligrams)
vd vitamin D (comprising D2 and D3, in micrograms)
vk vitamin K (in micrograms)
calc calcium (in milligrams)
phos phosphorus (in milligrams)
magn magnesium (in milligrams)
iron iron (in milligrams)
zinc zinc (in milligrams)
copp copper (in milligrams)
sodi sodium (in milligrams)
pota potassium (in milligrams)
sele selenium (in micrograms)
```

### Details

Note that the sample design oversampled specific population targets and that only respondants are provided. The website contains more information about sampling weights. There are multiple missing records.

#### Note

These data are subject to a data user agreement, available at https://www.cdc.gov/nchs/policy/data-user-agreement.html

### Source

National Center for Health Statistics, now available from the Wayback Machine via https://web.archive.org/web/20201029113801/https://wwwn.cdc.gov/Nchs/Nhanes/2015-2016/DR1TOT\_I.XPT

72 penultimate

pandemics

Deaths from pandemics

## **Description**

The data base contains estimated records of the number of deaths from pandemics.

## Usage

pandemics

#### **Format**

A data frame with 72 rows and 8 variables:

event name of the event
startyear start year of the event
endyear end year of the event
lower lower bound on estimated deaths (in thousands)
average average estimated deaths (in thousands)
upper upper bound on estimated deaths (in thousands)
saverage scaled average of estimated deaths (in thousands)
population estimated population at risk (in thousands)

#### Source

Cirillo, P. and N.N. Taleb (2020). *Tail risk of contagious diseases*. Nat. Phys. **16**, 606–613 (2020). <doi:10.1038/s41567-020-0921-x>

penultimate

Smith's penultimate approximations

## Description

The function takes as arguments the distribution and density functions. There are two options: method='bm' yields block maxima and method='pot' threshold exceedances. For method='bm', the user should provide in such case the block sizes via the argument m, whereas if method='pot', a vector of threshold values should be provided. The other argument (thresh or m depending on the method) is ignored.

## Usage

```
penultimate(family, method = c("bm", "pot"), thresh, qlev, m, ...)
```

penultimate 73

## **Arguments**

family	the name of the parametric family. Will be used to obtain dfamily, pfamily, qfamily
method	either block maxima ('bm') or peaks-over-threshold ('pot') are supported
thresh	vector of thresholds for method 'pot'
qlev	vector of quantile levels for method 'pot', e.g., $0.9, 0.95, \dots$ Ignored if argument thresh is provided.
m	vector of block sizes for method 'bm'
	additional arguments passed to densF and distF

#### **Details**

Alternatively, the user can provide functions densF, quantF and distF for the density, quantile function and distribution functions, respectively. The user can also supply the derivative of the density function, ddensF. If the latter is missing, it will be approximated using finite-differences.

For method = "pot", the function computes the reciprocal hazard and its derivative on the log scale to avoid numerical overflow. Thus, the density function should have argument log and the distribution function arguments log.p and lower.tail, respectively.

### Value

a data frame containing

- loc: location parameters (method='bm')
- scale: scale parameters
- · shape: shape parameters
- thresh: thresholds (if method='pot'), percentile corresponding to threshold (if method='pot')
- m: block sizes (if method='bm')

## Author(s)

Leo Belzile

#### References

Smith, R.L. (1987). Approximations in extreme value theory. *Technical report 205*, Center for Stochastic Process, University of North Carolina, 1–34.

### **Examples**

```
# Threshold exceedance for Normal variables
quants <- seq(1, 5, by = 0.02)
penult <- penultimate(
  family = "norm",
  method = 'pot',
  thresh = quants,
  ddensF = function(x){-x*dnorm(x)}, # optional argument</pre>
```

74 plot.eprof

```
plot(x = quants,
    y = penult$shape,
     type = '1',
    xlab = 'quantile',
   ylab = 'Penultimate shape',
   ylim = c(-0.5, 0)
# Block maxima for Gamma variables
# User must provide arguments for shape (or rate), for which there is no default
m < - seq(30, 3650, by = 30)
penult <- penultimate(family = 'gamma', method = 'bm', m = m, shape = 0.1)</pre>
plot(x = m,
    y = penult$shape,
     type = '1',
     xlab = 'quantile',
    ylab = 'penultimate shape')
# Comparing density of GEV approximation with true density of maxima
m <- 100 # block of size 100
p <- penultimate(</pre>
  family = 'norm',
  ddensF = function(x){-x*dnorm(x)},
 method = 'bm',
 m = m
x < - seq(1, 5, by = 0.01)
plot(
  x = x,
  y = m * dnorm(x) * exp((m-1) * pnorm(x, log.p = TRUE)),
  type = '1',
  ylab = 'density',
  main = 'Distribution of the maxima of\n 100 standard normal variates')
lines(x, mev::dgev(x, loc = ploc, scale = pscale, shape = 0), col = 2)
lines(x, mev::dgev(x, loc = p$loc, scale = p$scale, shape = p$shape), col = 4)
legend(
 x = 'topright',
 lty = c(1, 1, 1),
 col = c(1, 2, 4),
 legend = c('exact', 'ultimate', 'penultimate'),
 bty = 'n')
```

plot.eprof

Plot of (modified) profile likelihood

#### Description

The function plots the (modified) profile likelihood and the tangent exponential profile likelihood

#### Usage

```
## S3 method for class 'eprof'
plot(x, ...)
```

plot.fr 75

## Arguments

x an object of class eprof returned by gpd.pll or gev.pll.

... further arguments to plot.

## Value

a graph of the (modified) profile likelihoods

#### References

Brazzale, A. R., Davison, A. C. and Reid, N. (2007). *Applied Asymptotics: Case Studies in Small-Sample Statistics*. Cambridge University Press, Cambridge.

Severini, T. A. (2000). Likelihood Methods in Statistics. Oxford University Press, Oxford.

plot.fr

Plot of tangent exponential model profile likelihood

# Description

This function is adapted from the plot.fr function from the hoa package bundle. It differs from the latter mostly in the placement of legends.

## Usage

```
## S3 method for class 'fr' plot(x, ...)
```

#### **Arguments**

x an object of class fr returned by gpd. tem or gev. tem.

further arguments to plot currently ignored. Providing a numeric vector which allows for custom selection of the plots. A logical all. See **Details**.

## **Details**

Plots produced depend on the integers provided in which. 1 displays the Wald pivot, the likelihood root r, the modified likelihood root rstar and the likelihood modification q as functions of the parameter psi. 2 gives the renormalized profile log likelihood and adjusted form, with the maximum likelihood having ordinate value of zero. 3 provides the significance function, a transformation of 1. Lastly, 4 plots the correction factor as a function of the likelihood root; it is a diagnostic plot aimed for detecting failure of the asymptotic approximation, often due to poor numerics in a neighborhood of r=0; the function should be smooth. The function spline.corr is designed to handle this by correcting numerically unstable estimates, replacing outliers and missing values with the fitted values from the fit.

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## Value

graphs depending on argument which

## References

Brazzale, A. R., Davison, A. C. and Reid, N. (2007). *Applied Asymptotics: Case Studies in Small-Sample Statistics*. Cambridge University Press, Cambridge.

plot.mev\_shape\_rbm

Plots for random block maximum estimator

# Description

The function returns plot of the shape estimator along with the value (and 95% Wald-based confidence interval) at the selected threshold, or a plot of the empirical Bayes risk.

## Usage

```
## S3 method for class 'mev_shape_rbm'
plot(x, type = c("shape", "risk"), log = TRUE, ...)
```

## **Arguments**

Χ	object of class mev_shape_rbm returned by shape.rbm
type	[string] type of plot, either "shape" for the tail index or "risk" for the empirical Bayes risk
log	[logical] if TRUE (default), the x-axis for the number of exceedances is displayed on the log scale.
	additional arguments, currently ignored

## Value

one or more plots

```
plot.mev_thselect_vmetric
```

Plots for Varty and al. metric-based threshold selection

## Description

This S3 method produces quantile-quantile plots with confidence and tolerance bands on various scale (uniform, exponential, generalized Pareto), or a plot of the metric as a function of the threshold.

## Usage

```
## S3 method for class 'mev_thselect_vmetric'
plot(
    x,
    type = c("qq", "pp", "exp", "metric"),
    B = 1000L,
    probs = c(0.025, 0.975),
    ...
)
```

## **Arguments**

```
x an object of class mev_thselect_vmetric produced by a call to thselect.vmetric
type string; a single string indicating the choice of plot

B number of simulations for variability of estimation
probs quantile levels for intervals.
... additional arguments, currently ignored
```

#### Value

NULL; the function is used to produce a plot

```
plot.mev_thselect_wadsworth
```

Sequential analysis diagnostic plots for threshold selection

# Description

Sequential analysis diagnostic plots for threshold selection

#### Usage

```
## S3 method for class 'mev_thselect_wadsworth'
plot(x, type = c("wn", "ps"), ...)
```

78 plot.mev\_tstab\_mrl

# Arguments

Χ	object returned by a call to thselect.wseq
type	string giving the plots to produce
	additional arguments passed to plotting function

## Value

NULL; the method is used to generate plots

```
plot.mev_tstab_mrl Mean residual life parameter stability plot
```

# Description

Mean residual life parameter stability plot

# Usage

```
## S3 method for class 'mev_tstab_mrl'
plot(
    x,
    xlab = c("thresh", "nexc"),
    level = 0.95,
    type = c("band", "ptwise"),
    ...
)
```

# Arguments

X	object resulting from a call to tstab.mrl
xlab	[string]; whether to plot mean residual life plot as a function of threshold value of number of exceedances
level	[numeric] level of Wald confidence intervals
type	[string] whether to plot pointwise confidence intervals using segments ("ptwise") or using dashed lines ("band")
	additional arguments, currently ignored

## Value

```
NULL; use to produce plots
```

*pp* 79

pp	Poisson process of extremes.	
----	------------------------------	--

## **Description**

Likelihood, score function and information matrix for the Poisson process likelihood.

## **Arguments**

par	vector of loc, scale and shape
dat	sample vector
u	threshold
method	string indicating whether to use the expected ('exp') or the observed ('obs' - the default) information matrix.
np	number of periods of observations. This is a <i>post hoc</i> adjustment for the intensity so that the parameters of the model coincide with those of a generalized extreme value distribution with block size length(dat)/np.
nobs	number of observations for the expected information matrix. Default to length(dat) if dat is provided.

### Usage

```
pp.ll(par, dat)
pp.ll(par, dat, u, np)
pp.score(par, dat)
pp.infomat(par, dat, method = c('obs', 'exp'))
```

## **Functions**

pp.11: log likelihoodpp.score: score vector

 $\bullet$  pp.infomat: observed or expected information matrix

# Author(s)

Leo Belzile

#### References

Coles, S. (2001). An Introduction to Statistical Modeling of Extreme Values, Springer, 209 p.

Wadsworth, J.L. (2016). Exploiting Structure of Maximum Likelihood Estimators for Extreme Value Threshold Selection, *Technometrics*, **58**(1), 116-126, http://dx.doi.org/10.1080/00401706.2014.998345.

Sharkey, P. and J.A. Tawn (2017). A Poisson process reparameterisation for Bayesian inference for extremes, Extremes, 20(2), 239-263, http://dx.doi.org/10.1007/s10687-016-0280-2.

80 qweissman

qweissman

Weissman's quantile estimator

### **Description**

Given a small probability of exceedance p, the number of exceedances k out of n observation above the threshold u (thresh) (corresponding typically to the (k+1)th order statistic, compute the tail quantile at level Q(1-p) using the estimator of Weissman (1978) under the assumption of Pareto tail (positive shape  $\xi$ ), viz.

$$Q(1-p) = u \left(\frac{k}{pn}\right)^{\xi}.$$

## Usage

```
qweissman(p, k, n, thresh, shape)
```

#### **Arguments**

p tail probability, must be larger than the proportion of exceedances k/n.

k vector of the number of exceedances above thresh

n integer, total sample size thresh vector of thresholds

shape vector of positive shape parameters

### Value

a vector of tail quantiles

## References

Weissman, I. (1978). Estimation of Parameters and Larger Quantiles Based on the *k* Largest Observations. *Journal of the American Statistical Association*, 73(**364**), 812–815. <doi:10.2307/2286285>.

#### **Examples**

```
set.seed(2025)
p <- 1/100
xdat <- rgp(n = 1000, loc = 2, scale = 2, shape = 0.4)
hill <- shape.hill(xdat, k = seq(20L, 100L, by = 10L))
thresh <- sort(xdat, decreasing = TRUE)[hill$k+1]
qweissman(
    p = 1/100,
    k = hill$k,
    n = length(xdat),
    thresh = thresh,
    shape = hill$shape)
# Compare with true quantile
qgp(1/100, loc = 2, scale = 2, shape = 0.4, lower.tail = FALSE)</pre>
```

rdir 81

rdir

Random variate generation for Dirichlet distribution on  $S\_d$ 

## **Description**

A function to sample Dirichlet random variables, based on the representation as ratios of Gamma. Note that the RNG will generate on the full simplex and the sum to one constraint is respected here

# Usage

```
rdir(n, alpha, normalize = TRUE)
```

## **Arguments**

n sample size

alpha vector of parameter

normalize boolean. If FALSE, the function returns Gamma variates with parameter alpha.

### Value

sample of dimension d (size of alpha) from the Dirichlet distribution.

### **Examples**

```
rdir(n=100, alpha=c(0.5,0.5,2),TRUE)
rdir(n=100, alpha=c(3,1,2),FALSE)
```

rgparp

Simulation from generalized R-Pareto processes

### **Description**

The generalized R-Pareto process is supported on (loc - scale / shape, Inf) if shape > 0, or (-Inf, loc - scale / shape) for negative shape parameters, conditional on (X-r(loc))/r(scale) > 0. The standard Pareto process corresponds to scale = loc = rep(1, d).

# Usage

```
rgparp(
    n,
    shape = 1,
    thresh = 1,
    risk = c("mean", "sum", "site", "max", "min", "12"),
    siteindex = NULL,
    d,
```

82 rgparp

```
loc,
scale,
param,
sigma,
model = c("log", "neglog", "bilog", "negbilog", "hr", "br", "xstud", "smith",
    "schlather", "ct", "sdir", "dirmix"),
weights,
vario,
coord = NULL,
...
)
```

#### **Arguments**

n number of observations

shape shape parameter of the generalized Pareto variable

thresh univariate threshold for the exceedances of risk functional

risk string indicating the risk functional.

siteindex integer between 1 and d specifying the index of the site or variable

d dimension of sample

loc location vector scale scale vector

param parameter vector for the logistic, bilogistic, negative bilogistic and extremal

Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture.

Degree of freedoms for extremal student model. See Details.

sigma covariance matrix for Brown-Resnick and extremal Student-t distributions. Sym-

metric matrix of squared coefficients  $\lambda^2$  for the Husler-Reiss model, with zero

diagonal elements.

model for multivariate extreme value distributions, users can choose between 1-parameter

logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported

weights vector of length m for the m mixture components that sum to one. For the

"maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this

argument overrides asy).

vario semivariogram function whose first argument must be distance. Used only if

provided in conjunction with coord and if sigma is missing

coord d by k matrix of coordinates, used as input in the variogram vario or as param-

eter for the Smith model. If grid is TRUE, unique entries should be supplied.

... additional arguments for the vario function

rho.dk 83

### Value

an n by d sample from the generalized R-Pareto process, with attributes accept.rate if the procedure uses rejection sampling.

## **Examples**

```
rgparp(n = 10, risk = 'site', siteindex = 2, d = 3, param = 2.5,
   model = 'log', scale = c(1, 2, 3), loc = c(2, 3, 4), shape = 0.5)
rgparp(n = 10, risk = 'max', d = 4, param = c(0.2, 0.1, 0.9, 0.5),
   scale = 1:4, loc = 1:4, model = 'bilog')
rgparp(n = 10, risk = 'sum', d = 3, param = c(0.8, 1.2, 0.6, -0.5),
   scale = 1:3, loc = 1:3, model = 'sdir')
vario <- function(x, scale = 0.5, alpha = 0.8){ scale*x^alpha }
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
rgparp(n = 10, risk = 'max', vario = vario, coord = grid.coord,
   model = 'br', scale = runif(16), loc = rnorm(16))</pre>
```

rho.dk

Second order tail index estimator of Drees and Kaufmann

## **Description**

Estimator of the second order regular variation parameter  $\rho \leq 0$  parameter for heavy-tailed data proposed by Drees and Kaufmann (1998)

#### Usage

```
rho.dk(xdat, k, tau = 0.5)
```

### **Arguments**

xdat vector of positive observations  $k \qquad \qquad \text{number of highest order statistics to use for estimation}$  tau  $\qquad \qquad \text{tuning parameter } \tau \in (0,1)$ 

#### References

Drees, H. and E. Kaufmann (1998). Selecting the optimal sample fraction in univariate extreme value estimation, *Stochastic Processes and their Applications*, 75(2), 149-172, <doi:10.1016/S0304-4149(98)00017-9>.

rho.fagh

rho.fagh	Second order tail index estimator of Fraga Alves et al. Estimator of
	the second order regular variation parameter $\rho \leq 0$ parameter for heavy-tailed data proposed by Fraga Alves et al. (2003)

## **Description**

Second order tail index estimator of Fraga Alves et al. Estimator of the second order regular variation parameter  $\rho \le 0$  parameter for heavy-tailed data proposed by Fraga Alves et al. (2003)

## Usage

```
rho.fagh(xdat, k, tau = 0)
```

## **Arguments**

xdat vector of positive observations
 k number of highest order statistics to use for estimation
 tau scalar real tuning parameter. Default values is 0, which is typically chosen when-

ever  $\rho \ge -1$ . The choice  $\tau = 1$  otherwise.

#### References

Fraga Alves, M.I., Gomes, M. Ivette, and de Haan, Laurens (2003). A new class of semi-parametric estimators of the second order parameter. *Portugaliae Mathematica*. Nova Serie 60(2), 193-213. <a href="http://eudml.org/doc/50867">http://eudml.org/doc/50867</a>>.

## **Examples**

```
# Example with rho = -0.2
n <- 1000
xdat <- mev::rgp(n = n, shape = 0.2)
kmin <- floor(n^0.995)
kmax <- ceiling(n^0.999)
rho_est <- rho.fagh(
    xdat = xdat,
    k = n - kmin:kmax)
rho_med <- mean(rho_est$rho)</pre>
```

rho.gbw 85

rho.gbw

Second order tail index estimator of Goegebeur et al. (2008)

## **Description**

Estimator of the second order regular variation parameter  $\rho \leq 0$  parameter for heavy-tailed data based on ratio of kernel goodness-of-fit statistics.

# Usage

```
rho.gbw(xdat, k)
```

## Arguments

xdat vector of positive observations

k number of highest order statistics to use for estimation

#### References

Goegebeur, Y., J. Beirlant and T. de Wet (2008). Linking Pareto-tail kernel goodness-of-fit statistics with tail index at optimal threshold and second order estimation. *REVSTAT-Statistical Journal*, 6(1), 51-69. <doi:10.57805/revstat.v6i1.57>

rho.ghp

Second order tail index estimator of Gomes et al.

## **Description**

Estimator of the second order regular variation parameter  $\rho \leq 0$  parameter for heavy-tailed data proposed by Gomes et al. (2003)

## Usage

```
rho.ghp(xdat, k, alpha = 2)
```

## Arguments

xdat vector of positive observations

k number of highest order statistics to use for estimation

alpha positive scalar tuning parameter

## References

Gomes, M.I., de Haan, L. & Peng, L. (2002). Semi-parametric Estimation of the Second Order Parameter in Statistics of Extremes. *Extremes* 5, 387–414. <doi:10.1023/A:1025128326588>

86 rlarg

rlarg	Distribution of the r-largest observations
-------	--

## **Description**

Likelihood, score function and information matrix for the r-largest observations likelihood.

## **Arguments**

par	vector of loc, scale and shape
dat	an n by r sample matrix, ordered from largest to smallest in each row
method	string indicating whether to use the expected ('exp') or the observed ('obs'-the default) information matrix.
nobs	number of observations for the expected information matrix. Default to $nrow(dat)$ if $dat$ is provided.
r	number of order statistics kept. Default to ncol(dat)

### Usage

```
rlarg.ll(par, dat, u, np)
rlarg.score(par, dat)
rlarg.infomat(par, dat, method = c('obs', 'exp'), nobs = nrow(dat), r = ncol(dat))
```

## **Functions**

- rlarg.ll: log likelihood
- rlarg.score: score vector
- $\bullet$  rlarg.infomat: observed or expected information matrix

## Author(s)

Leo Belzile

### References

```
Coles, S. (2001). An Introduction to Statistical Modeling of Extreme Values, Springer, 209 p. Smith, R.L. (1986). Extreme value theory based on the r largest annual events, Journal of Hydrology, 86(1-2), 27–43, http://dx.doi.org/10.1016/0022-1694(86)90004-1.
```

rmev 87

rmev

Exact simulations of multivariate extreme value distributions

## Description

Implementation of the random number generators for multivariate extreme-value distributions and max-stable processes based on the two algorithms described in Dombry, Engelke and Oesting (2016).

## Usage

```
rmev(
 n,
 d,
 param,
 asy,
  sigma,
 model = c("log", "alog", "neglog", "aneglog", "bilog", "negbilog", "hr", "br", "xstud",
  "smith", "schlather", "ct", "sdir", "dirmix", "pairbeta", "pairexp", "wdirbs",
    "wexpbs", "maxlin"),
 alg = c("ef", "sm"),
 weights = NULL,
 vario = NULL,
  coord = NULL,
  grid = FALSE,
 dist = NULL,
)
```

### **Arguments**

n	number of observations
d	dimension of sample
param	parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See <b>Details</b> .
asy	list of asymmetry parameters, as in function rmvevd from package evd, of $2^d-1$ vectors of size corresponding to the power set of d, with sum to one constraints.
sigma	covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients $\lambda^2$ for the Husler-Reiss model, with zero diagonal elements.
model	for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported

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alg	algorithm, either simulation via extremal function ('ef') or via the spectral measure ('sm'). Default to ef.
weights	vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides asy).
vario	semivariogram function whose first argument must be distance. Used only if provided in conjunction with coord and if sigma is missing
coord	d by k matrix of coordinates, used as input in the variogram vario or as parameter for the Smith model. If grid is TRUE, unique entries should be supplied.
grid	Logical. TRUE if the coordinates are two-dimensional grid points (spatial models).
dist	symmetric matrix of pairwise distances. Default to NULL.
	additional arguments for the vario function

#### **Details**

The vector param differs depending on the model

- log: one dimensional parameter greater than 1
- alog:  $2^d d 1$  dimensional parameter for dep. Values are recycled if needed.
- neglog: one dimensional positive parameter
- aneglog:  $2^d d 1$  dimensional parameter for dep. Values are recycled if needed.
- bilog: d-dimensional vector of parameters in [0, 1]
- negbilog: d-dimensional vector of negative parameters
- ct, dir, negdir, sdir: d-dimensional vector of positive (a)symmetry parameters. For dir and negdir, a d+1 vector consisting of the d Dirichlet parameters and the last entry is an index of regular variation in  $(-\min(\alpha_1,\ldots,\alpha_d),1]$  treated as shape parameter
- xstud: one dimensional parameter corresponding to degrees of freedom alpha
- dirmix: d by m-dimensional matrix of positive (a)symmetry parameters
- pairbeta, pairexp: d(d-1)/2+1 vector of parameters, containing the concentration parameter and the coefficients of the pairwise beta, in lexicographical order e.g.,  $\beta_{12}, \beta_{13}, \ldots$
- wdirbs, wexpbs: 2d vector of d concentration parameters followed by the d Dirichlet parameters

Stephenson points out that the multivariate asymmetric negative logistic model given in e.g. Coles and Tawn (1991) is not a valid distribution function in dimension d>3 unless additional constraints are imposed on the parameter values. The implementation in mev uses the same construction as the asymmetric logistic distribution (see the vignette). As such it does not match the bivariate implementation of rbvevd.

The dependence parameter of the evd package for the Husler-Reiss distribution can be recovered taking for the Brown-Resnick model  $2/r=\sqrt(2\gamma(h))$  where h is the lag vector between sites and  $r=1/\lambda$  for the Husler-Reiss.

rmev 89

#### Value

an n by d exact sample from the corresponding multivariate extreme value model

#### Warning

As of version 1.8 (August 16, 2016), there is a distinction between models hr and br. The latter is meant to be used in conjunction with variograms. The parametrization differs between the two models.

The family of scaled Dirichlet is now parametrized by a parameter in  $-\min(\alpha)$  appended to the the d vector param containing the parameter alpha of the Dirichlet model. Arguments model='dir' and model='negdir' are still supported internally, but not listed in the options.

## Author(s)

Leo Belzile

#### References

Dombry, Engelke and Oesting (2016). Exact simulation of max-stable processes, *Biometrika*, **103**(2), 303–317.

#### See Also

rmevspec, rmvevd, rbvevd

## **Examples**

```
set.seed(1)
rmev(n=100, d=3, param=2.5, model='log', alg='ef')
rmev(n=100, d=4, param=c(0.2,0.1,0.9,0.5), model='bilog', alg='sm')
## Spatial example using power variogram
#NEW: Semi-variogram must take distance as argument
semivario <- function(x, scale, alpha){ scale*x^alpha }</pre>
#grid specification
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))</pre>
rmev(n=100, vario=semivario, coord=grid.coord, model='br', scale = 0.5, alpha = 1)
#using the Brown-Resnick model with a covariance matrix
vario2cov <- function(coord, semivario,...){</pre>
sapply(1:nrow(coord), function(i) sapply(1:nrow(coord), function(j)
 semivario(sqrt(sum((coord[i,])^2)), ...) +
 semivario(sqrt(sum((coord[j,])^2)), ...) -
 semivario(sqrt(sum((coord[i,]-coord[j,])^2)), ...)))
}
rmev(n=100, sigma=vario2cov(grid.coord, semivario = semivario, scale = 0.5, alpha = 1), model='br')
# asymmetric logistic model - see function 'rmvevd' from package 'evd '
asy <- list(0, 0, 0, 0, c(0,0), c(0,0), c(0,0), c(0,0), c(0,0), c(0,0),
 c(.2,.1,.2), c(.1,.1,.2), c(.3,.4,.1), c(.2,.2,.2), c(.4,.6,.2,.5)
rmev(n=1, d=4, param=0.3, asy=asy, model="alog")
#Example with a grid (generating an array)
rmev(n=10, sigma=cbind(c(2,1), c(1,3)), coord=cbind(runif(4), runif(4)), model='smith', grid=TRUE)
## Example with Dirichlet mixture
```

90 rmevspec

```
alpha.mat <- cbind(c(2,1,1),c(1,2,1),c(1,1,2))
rmev(n=100, param=alpha.mat, weights=rep(1/3,3), model='dirmix')
rmev(n=10, param=c(0.1,1,2,3), d=3, model='pairbeta')
```

rmevspec

Random samples from spectral distributions of multivariate extreme value models.

# **Description**

Generate from  $Q_i$ , the spectral measure of a given multivariate extreme value model based on the L1 norm.

## Usage

```
rmevspec(
    n,
    d,
    param,
    sigma,
    model = c("log", "neglog", "bilog", "negbilog", "hr", "br", "xstud", "smith",
    "schlather", "ct", "sdir", "dirmix", "pairbeta", "pairexp", "wdirbs", "wexpbs",
        "maxlin"),
    weights = NULL,
    vario = NULL,
    coord = NULL,
    grid = FALSE,
    dist = NULL,
    ...
)
```

### **Arguments**

n	number of observations
d	dimension of sample
param	parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See <b>Details</b> .
sigma	covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients $\lambda^2$ for the Husler-Reiss model, with zero diagonal elements.
model	for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported

rmevspec 91

weights	vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides asy).
vario	semivariogram function whose first argument must be distance. Used only if provided in conjunction with coord and if sigma is missing
coord	d by k matrix of coordinates, used as input in the variogram vario or as parameter for the Smith model. If grid is TRUE, unique entries should be supplied.
grid	Logical. TRUE if the coordinates are two-dimensional grid points (spatial models). $ \\$
dist	symmetric matrix of pairwise distances. Default to NULL.
	additional arguments for the vario function

### **Details**

The vector param differs depending on the model

- log: one dimensional parameter greater than 1
- neglog: one dimensional positive parameter
- bilog: d-dimensional vector of parameters in [0, 1]
- negbilog: d-dimensional vector of negative parameters
- ct, dir, negdir: d-dimensional vector of positive (a)symmetry parameters. Alternatively, a d+1 vector consisting of the d Dirichlet parameters and the last entry is an index of regular variation in (0,1] treated as scale
- xstud: one dimensional parameter corresponding to degrees of freedom alpha
- dirmix: d by m-dimensional matrix of positive (a)symmetry parameters
- pairbeta, pairexp: d(d-1)/2+1 vector of parameters, containing the concentration parameter and the coefficients of the pairwise beta, in lexicographical order e.g.,  $\beta_{1,2}, \beta_{1,3}, \dots$
- wdirbs, wexpbs: 2d vector of d concentration parameters followed by the d Dirichlet parameters

### Value

an n by d exact sample from the corresponding multivariate extreme value model

# Note

This functionality can be useful to generate for example Pareto processes with marginal exceedances.

#### Author(s)

Leo Belzile

92 rmnorm

### References

Dombry, Engelke and Oesting (2016). Exact simulation of max-stable processes, *Biometrika*, **103**(2), 303–317.

Boldi (2009). A note on the representation of parametric models for multivariate extremes. *Extremes* **12**, 211–218.

## **Examples**

```
set.seed(1)
rmevspec(n=100, d=3, param=2.5, model='log')
rmevspec(n=100, d=3, param=2.5, model='neglog')
rmevspec(n=100, d=4, param=c(0.2,0.1,0.9,0.5), model='bilog')
rmevspec(n=100, d=2, param=c(0.8,1.2), model='ct') #Dirichlet model
rmevspec(n=100, d=2, param=c(0.8,1.2,0.5), model='sdir') #with additional scale parameter
#Variogram gamma(h) = scale*||h||'alpha
#NEW: Variogram must take distance as argument
vario <- function(x, scale=0.5, alpha=0.8){ scale*x^alpha }
#grid specification
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
rmevspec(n=100, vario=vario,coord=grid.coord, model='br')
## Example with Dirichlet mixture
alpha.mat <- cbind(c(2,1,1),c(1,2,1),c(1,1,2))
rmevspec(n=100, param=alpha.mat, weights=rep(1/3,3), model='dirmix')</pre>
```

rmnorm

Multivariate Normal distribution sampler

## Description

Sampler derived using the eigendecomposition of the covariance matrix Sigma. The function uses the Armadillo random normal generator

## Usage

```
rmnorm(n, mu, Sigma)
```

## **Arguments**

n sample size

mu mean vector. Will set the dimension

Sigma a square covariance matrix, of same dimension as mu. No sanity check is per-

formed to validate that the matrix is positive definite, so use at own risk

## Value

an n sample from a multivariate Normal distribution

rparp 93

## **Examples**

```
rmnorm(n = 10, mu = c(0,2), Sigma = diag(2))
```

rparp

Simulation from R-Pareto processes

# Description

Simulation from R-Pareto processes

# Usage

```
rparp(
    n,
    shape = 1,
    risk = c("sum", "site", "max", "min", "12"),
    siteindex = NULL,
    d,
    param,
    sigma,
    model = c("log", "neglog", "bilog", "negbilog", "hr", "br", "xstud", "smith",
        "schlather", "ct", "sdir", "dirmix"),
    weights,
    vario,
    coord = NULL,
    ...
)
```

## **Arguments**

_		
n	ı	number of observations
sh	hape	shape tail index of Pareto variable
ri	risk	string indicating risk functional.
si	siteindex	integer between 1 and d specifying the index of the site or variable
d	I	dimension of sample
pa	param	parameter vector for the logistic, bilogistic, negative bilogistic and extremal Dirichlet (Coles and Tawn) model. Parameter matrix for the Dirichlet mixture. Degree of freedoms for extremal student model. See <b>Details</b> .
si	igma	covariance matrix for Brown-Resnick and extremal Student-t distributions. Symmetric matrix of squared coefficients $\lambda^2$ for the Husler-Reiss model, with zero diagonal elements.
mc	nodel	for multivariate extreme value distributions, users can choose between 1-parameter logistic and negative logistic, asymmetric logistic and negative logistic, bilogistic, Husler-Reiss, extremal Dirichlet model (Coles and Tawn) or the Dirichlet mixture. Spatial models include the Brown-Resnick, Smith, Schlather and extremal Student max-stable processes. Max linear models are also supported
		1

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weights	vector of length m for the m mixture components that sum to one. For the "maxlin" model, weights should be a matrix with d columns that represent the weight of the components and whose column sum to one (if provided, this argument overrides asy).
vario	semivariogram function whose first argument must be distance. Used only if provided in conjunction with coord and if sigma is missing
coord	d by k matrix of coordinates, used as input in the variogram vario or as parameter for the Smith model. If grid is TRUE, unique entries should be supplied.
	additional arguments for the vario function

#### **Details**

For riskf=max and riskf=min, the procedure uses rejection sampling based on Pareto variates sampled from sum and may be slow if d is large.

### Value

an n by d sample from the R-Pareto process, with attributes accept.rate if the procedure uses rejection sampling.

## **Examples**

```
rparp(n=10, risk = 'site', siteindex=2, d=3, param=2.5, model='log')
rparp(n=10, risk = 'min', d=3, param=2.5, model='neglog')
rparp(n=10, risk = 'max', d=4, param=c(0.2,0.1,0.9,0.5), model='bilog')
rparp(n=10, risk = 'sum', d=3, param=c(0.8,1.2,0.6, -0.5), model='sdir')
vario <- function(x, scale=0.5, alpha=0.8){ scale*x^alpha }
grid.coord <- as.matrix(expand.grid(runif(4), runif(4)))
rparp(n=10, risk = 'max', vario=vario, coord=grid.coord, model='br')</pre>
```

rparpcs

Simulation from Pareto processes using composition sampling

## Description

The algorithm performs forward sampling by simulating first from a mixture, then sample angles conditional on them being less than (max) or greater than (min) one. The resulting sample from the angular distribution is then multiplied by Pareto variates with tail index shape.

# Usage

```
rparpcs(
   n,
   model = c("log", "neglog", "br", "xstud"),
   risk = c("max", "min"),
   param = NULL,
   d,
```

rparpes 95

```
Lambda = NULL,
Sigma = NULL,
df = NULL,
shape = 1,
....
```

## **Arguments**

n	sample size.
model	string indicating the model family.
risk	string indicating the risk functional. Only max and min are currently supported.
param	parameter value for the logistic or negative logistic model
d	dimension of the multivariate model, only needed for logistic or negative logistic models $$
Lambda	parameter matrix for the Brown-Resnick model. See Details.
Sigma	correlation matrix if model = 'xstud', otherwise the covariance matrix formed from the stationary Brown-Resnick process.
df	degrees of freedom for extremal Student process.
shape	tail index of the Pareto variates (reciprocal shape parameter). Must be strictly positive.
	additional parameters, currently ignored

#### **Details**

For the moment, only exchangeable models and models based n elliptical processes are handled.

The parametrization of the Brown–Resnick is in terms of the matrix Lambda, which is formed by evaluating the semivariogram  $\gamma$  at sites  $s_i, s_j$ , meaning that  $\Lambda_{i,j} = \gamma(s_i, s_j)/2$ .

The argument Sigma is ignored for the Brown-Resnick model if Lambda is provided by the user.

# Value

an n by d matrix of samples, where d = ncol(Sigma), with attributes mixt.weights.

## Author(s)

Leo Belzile

## See Also

rparp for general simulation of Pareto processes based on an accept-reject algorithm.

96 rparpcshr

### **Examples**

```
## Not run:
#Brown-Resnick, Wadsworth and Tawn (2014) parametrization
D <- 20L
coord <- cbind(runif(D), runif(D))
semivario <- function(d, alpha = 1.5, lambda = 1){0.5 * (d/lambda)^alpha}
Lambda <- semivario(as.matrix(dist(coord))) / 2
rparpcs(n = 10, Lambda = Lambda, model = 'br', shape = 0.1)
#Extremal Student
Sigma <- stats::rWishart(n = 1, df = 20, Sigma = diag(10))[,,1]
rparpcs(n = 10, Sigma = cov2cor(Sigma), df = 3, model = 'xstud')
## End(Not run)</pre>
```

rparpcshr

Simulation of generalized Huesler-Reiss Pareto vectors via composition sampling

## Description

Sample from the generalized Pareto process associated to Huesler-Reiss spectral profiles. For the Huesler-Reiss Pareto vectors, the matrix Sigma is utilized to build Q viz.

$$Q = \Sigma^{-1} - \frac{\Sigma^{-1} \mathbf{1}_d \mathbf{1}_d^\top \Sigma^{-1}}{\mathbf{1}_d^\top \Sigma^{-1} \mathbf{1}_d}.$$

The location vector m and Sigma are the parameters of the underlying log-Gaussian process.

## Usage

```
rparpcshr(n, u, alpha, Sigma, m)
```

#### **Arguments**

n	sample size
u	vector of marginal location parameters (must be strictly positive)
alpha	vector of shape parameters (must be strictly positive).
Sigma	covariance matrix of process, used to define $Q$ . See <b>Details</b> .
m	location vector of Gaussian distribution.

### Value

n by d matrix of observations

#### References

Ho, Z. W. O and C. Dombry (2019), Simple models for multivariate regular variations and the Huesler-Reiss Pareto distribution, Journal of Multivariate Analysis (173), p. 525-550, doi:10.1016/j.jmva.2019.04.008

rrlarg 97

### **Examples**

rrlarg

Simulate r-largest observations from point process of extremes

### **Description**

Simulate the r-largest observations from a Poisson point process with intensity

$$\Lambda(x) = (1 + \xi(x - \mu)/\sigma)^{-1/\xi}$$

.

### Usage

```
rrlarg(n, r, loc = 0, scale = 1, shape = 0)
```

## **Arguments**

n sample size
r number of observations per block
loc location parameter
scale scale parameter
shape shape parameter

#### Value

an n by r matrix of samples from the point process, ordered from largest to smallest in each row.

98 shape.erm

shape.erm	Exponential regression estimator	

## **Description**

This function implements the exponential regression estimator of the shape parameter for the case of Pareto tails with positive shape index.

## Usage

```
shape.erm(xdat, k, method = c("bdgm", "fh"), bounds = NULL)
```

## **Arguments**

xdat vector of observations vector of integer, the number of largest observations to consider string; one of bdgm for the approach of Beirlant, Dierckx, Goegebeur and Matthys method (1999) or fh for Feuerverger and Hall (1999) vector of length 2 giving the bounds for rho, the second order parameter. Default bounds

to  $\rho \in [-5, -0.5]$ 

#### **Details**

The second-order parameter is difficult to pin down, and while values within [-1,0) are most logical under Hall model, the model parameters become unidentifiable when  $\rho \to 0$ . The default constraint restrict  $-5 < \rho < -0.5$ , with the upper bound changed to -0.25 for sample of sizes larger than 5000 observations. Users can set the value of the bounds for  $\rho$  via argument bounds. The optimization is initialized at the Hill estimator.

#### Value

a data frame with columns

- k number of exceedances
- shape estimate of the shape parameter
- rho estimate of the second-order regular variation index
- scale estimate of the scale parameter

#### References

Feuerverger, A. and P. Hall (1999), Estimating a tail exponent by modelling departure from a Pareto distribution, The Annals of Statistics 27(2), 760-781. <doi:10.1214/aos/1018031215>

Beirlant, J., Dierckx, G., Goegebeur, Y. G. Matthys (1999). Tail Index Estimation and an Exponential Regression Model. Extremes, 2, 177–200 (1999). <doi:10.1023/A:1009975020370>

shape.genjack 99

geniack

Generalized jackknife shape estimator

## **Description**

Generalized jackknife shape estimator

### Usage

```
shape.genjack(xdat, k)
```

## Arguments

xdat vector of positive observations

k vector of order statistics; if missing, a vector going from 10 to sample size minus

one.

#### Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar integer.

### References

Gomes, I.M., João Martins, M. and Neves, M. (2000) Alternatives to a Semi-Parametric Estimator of Parameters of Rare Events-The Jackknife Methodology. *Extremes*, 3, 207–229. doi:10.1023/A:1011470010228

shape.genquant

Beirlant et al. generalized quantile shape estimator

# Description

This estimator estimates the real shape parameter based on generalized quantile plots based on mean excess functions, generalized median excesses or trimmed mean excesses.

# Usage

```
shape.genquant(
  xdat,
  k,
  type = c("genmean", "genmed", "trimmean"),
  weight,
  p = 0.5
)
```

100 shape.hill

### **Arguments**

xdat n vector of observations

k number of upper order statistics

type string indicating the estimator choice, one of genmean, genmed and trimmean.

weight weight a kernel function on [0, 1]

p number between zero and one giving the proportion of order statistics for the

second threshold

#### References

Beirlant, J., Vynckier P. and J.L. Teugels (1996). Excess functions and estimation of the extremevalue index. Bernoulli, 2(4), 293-318.

shape.hill

Hill's estimator of the shape parameter

## **Description**

Given a sample of positive observations, calculate the tail index or shape parameter. The shape estimate returned is positive.

## Usage

```
shape.hill(xdat, k)
```

### **Arguments**

xdat vector of positive observations

k vector of order statistics; if missing, a vector going from 10 to sample size minus

one.

#### Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar integer.

#### References

Hill, B.M. (1975). A simple general approach to inference about the tail of a distribution. Annals of Statistics, 3, 1163-1173.

## **Examples**

```
xdat <- mev::rgp(n = 200, loc = 1, scale = 0.5, shape = 0.5)
shape.hill(xdat)</pre>
```

shape.lthill 101

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Lower-trimmed Hill estimator for the shape parameter

## **Description**

Given a sample of Pareto-tailed samples (positive tail index), compute the lower-trimmed Hill estimator. If k0 = k, the estimator reduces to Hill's estimator for the shape index

# Usage

```
shape.lthill(xdat, k, k0 = k, sorted = FALSE, ...)
```

### **Arguments**

xdat	[numeric] vector of positive observations
k	[integer] number of order statistics for the threshold
k0	[integer] vector of number of largest order statistics, no greater than k
sorted	[logical] if TRUE, data are assumed to be sorted in decreasing order.
• • •	additional arguments for other routines (notably vectorize)

#### Value

a scalar with the shape parameter estimate if k0 is a scalar, otherwise a data frame with columns k0 for the number of exceedances and shape for the tail index.

### References

Bladt, M., Albrecher, H. & Beirlant, J. (2020) *Threshold selection and trimming in extremes*. Extremes, 23, 629-665. doi:10.1007/s10687020003850

# **Examples**

```
# Pareto sample
n <- 200
xdat <- 10/(1 - runif(n)) - 10
shape.lthill(xdat = xdat, k = 100, k0 = 5:100)</pre>
```

102 shape.osz

shape.moment

Dekkers and de Haan moment estimator for the shape

## **Description**

Given a sample of exceedances, compute the moment estimator of the real shape parameter.

# Usage

```
shape.moment(xdat, k)
```

### **Arguments**

xdat vector of positive observations of length n

k number of largest order statistics

#### Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar.

### References

Dekkers, A.L.M. and de Haan, L. (1989). On the estimation of the extreme-value index and large quantile estimation., Annals of Statistics, 17, 1795-1833.

shape.osz

Extreme U-statistic Pickands estimator

## **Description**

Given a random sample of n exceedances, the estimator returns an estimator of the shape parameter or extreme value index using a kernel of order 3, based on k largest exceedances of xdat. Note that the method does not allow for ties.

### Usage

```
shape.osz(xdat, k, ...)
```

### **Arguments**

 ${\sf xdat}$  vector of observations of length n

k number of largest order statistics  $3 \le k < n$ .

... additional arguments for backward compatibility

shape.pickands 103

## **Details**

The calculations are based on the recursions provided in Lemma 4.3 of Oorschot et al.

#### References

Oorschot, J, J. Segers and C. Zhou (2023), Tail inference using extreme U-statistics, Electronic Journal of Statistics 17(1): 1113-1159. doi:10.1214/23EJS2129

# **Examples**

```
xdat \leftarrow rgp(n = 1000, shape = 0.2)
shape.osz(xdat, k = 10)
```

shape.pickands

Pickand's shape estimator

## **Description**

Given a sample of size n of positive exceedances, compute the real shape parameter  $\xi$  based on the k largest order statistics.

#### Usage

```
shape.pickands(xdat, k)
```

### **Arguments**

 $\mbox{ vector of positive observations of length } n$ 

k number of largest order statistics

#### Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar.

### References

Pickands, III, J. (1975). Statistical inference using extreme order statistics. Annals of Statistics, 3, 119-131.

shape.trimhill

shape.rbm

Random block maxima shape estimator of Wager

## Description

Computes the shape estimator for varying k up to sample size of maximum kmax largest observations

## Usage

```
shape.rbm(xdat, k = 10:floor(length(xdat)/2), ...)
```

# **Arguments**

xdat [vector] sample exceedances

k [int] vector of integers for which to compute the estimator

... additional parameters, currently ignored

### Value

a list with elements

k number of exceedances

shape tail index estimate, strictly positive

risk empirical Bayes estimate of risk

thresh threshold given by the smallest order statistic considered in the sample

#### References

Wager, S. (2014). Subsampling extremes: From block maxima to smooth tail estimation, *Journal of Multivariate Analysis*, 130, 335-353, doi:10.1016/j.jmva.2014.06.010

shape.trimhill

Trimmed Hill estimator for the shape parameter

# Description

Given a sample of Pareto-tailed samples (positive tail index), compute the trimmed Hill estimator. If k0 = k, the estimator reduces to Hill's estimator for the shape index

# Usage

```
shape.trimhill(xdat, k, k0, sorted = FALSE)
```

shape.vries 105

## **Arguments**

xdat	[numeric]	vector of	positive	observations

k [integer] number of order statistics for the threshold

k0 [integer] number of largest order statistics, strictly less than k sorted [logical] if TRUE, data are assumed to be sorted in decreasing order

## Value

a scalar with the shape parameter estimate

#### References

Bhattacharya, S., Kallitsis, M. and S. Stoev, (2019) Data-adaptive trimming of the Hill estimator and detection of outliers in the extremes of heavy-tailed data. Electronic Journal of Statistics 13, 1872–1925

shape.vries

de Vries shape estimator

## **Description**

Given a sample of exceedances, compute the moment estimator of the positive shape parameter using the ratio of log ratio of exceedance and it's square.

### Usage

```
shape.vries(xdat, k)
```

### **Arguments**

xdat vector of positive observations of length n

k number of largest order statistics

## Value

a data frame with the number of order statistics k and the shape parameter estimate shape, or a single numeric value if k is a scalar.

## References

de Haan, L. and Peng, L. (1998). *Comparison of tail index estimators*, Statistica Neerlandica 52, 60-70.

de Haan, L. and Peng, L. (1998). Comparison of tail index estimators. *Statistica Neerlandica*, 52: 60-70. doi:10.1111/14679574.00068

106 spunif

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Semi-parametric marginal transformation to uniform

## **Description**

The function spunif transforms a matrix or vector of data x to the pseudo-uniform scale using a semiparametric transform. Data below the threshold are transformed to pseudo-uniforms using a rank transform, while data above the threshold are assumed to follow a generalized Pareto distribution. The parameters of the latter are estimated using maximum likelihood if either scale = NULL or shape = NULL.

## Usage

```
spunif(x, thresh, scale = NULL, shape = NULL)
```

## **Arguments**

x	matrix or vector of data
thresh	vector of marginal thresholds
scale	vector of marginal scale parameters for the generalized Pareto
shape	vector of marginal shape parameters for the generalized Pareto

## Value

a matrix or vector of the same dimension as x, with pseudo-uniform observations

## Author(s)

Leo Belzile

# **Examples**

```
x \leftarrow rmev(1000, d = 3, param = 2, model = 'log')
thresh <- apply(x, 2, quantile, 0.95)
spunif(x, thresh)
```

taildep 107

taildep

Coefficient of tail correlation and tail dependence

## **Description**

For data with unit Pareto margins, the coefficient of tail dependence  $\eta$  is defined via

$$Pr(\min(X) > x) = L(x)x^{-1/\eta},$$

where L(x) is a slowly varying function. Ignoring the latter, several estimators of  $\eta$  can be defined. In unit Pareto margins,  $\eta$  is a nonnegative shape parameter that can be estimated by fitting a generalized Pareto distribution above a high threshold. In exponential margins,  $\eta$  is a scale parameter and the maximum likelihood estimator of the latter is the Hill estimator. Both methods are based on peaks-over-threshold and the user can choose between pointwise confidence obtained through a likelihood ratio test statistic ("lrt") or the Wald statistic ("wald").

# Usage

```
taildep(
  xdat,
  qlev = NULL,
 nq = 40,
 qlim = c(0.8, 0.99),
  depmeas = c("eta", "chi"),
 estimator = list(eta = c("emp", "betacop", "gpd", "hill"), chi = c("emp", "betacop")),
  confint = c("wald", "lrt"),
 level = 0.95,
  trunc = TRUE,
 margtrans = c("emp", "none"),
  ties.method = "random",
 plot = TRUE,
)
```

## **Arguments**

xdat	an $n$ by $d$ matrix of multivariate observations
qlev	vector of percentiles between 0 and 1
nq	number of quantiles of the structural variable at which to form a grid; only used if $u = NULL$ .
qlim	limits for the sequence u of the structural variable
depmeas	dependence measure, either of "eta" or "chi"
estimator	named list giving the estimation method for eta and chi. Default to "emp" for both.
confint	string indicating the type of confidence interval for $\eta$ , one of "wald" or "lrt"

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level the confidence level required (default to 0.95).

trunc logical indicating whether the estimates and confidence intervals should be trun-

cated in [0,1]

margirans marginal transformation; if "none", data are assumed to be in uniform margins

ties.method string indicating the type of method for rank; see rank for a list of options.

Default to "random"

plot logical; should graphs be plotted?

.. additional arguments passed to plot; current support for main, xlab, ylab, add

and further pch, 1ty, type, col for points; additional arguments for confidence

intervals are handled via cipch, cilty, citype, cicol.

#### **Details**

The most common approach for estimation is the empirical survival copula, by evaluating the proportion of sample minima with uniform margins that exceed a given x. An alternative estimator uses a smoothed estimator of the survival copula using Bernstein polynomial, resulting in the so-called betacop estimator. Approximate pointwise confidence intervals for the latter are obtained by assuming the proportion of points is binomial.

The coefficient of tail correlation  $\chi$  is

$$\chi = \lim_{u \to 1} \frac{\Pr(F_1(X_1) > u, \dots, F_D(X_D) > u)}{1 - u}.$$

Asymptotically independent vectors have  $\chi=0$ . The estimator uses an estimator of the survival copula

#### Value

a named list with elements

- glev: a K vector of percentile levels
- eta: a K by 3 matrix with point estimates, lower and upper confidence intervals
- chi: a K by 3 matrix with point estimates, lower and upper confidence intervals

## Note

As of version 1.15, the percentiles used are from the minimum variable. This ensures that, regardless of the number of variables, there is no error message returned because the quantile levels are too low for there to be observations

#### See Also

chiplot for bivariate empirical estimates of  $\chi$  and  $\bar{\chi}$ .

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## **Examples**

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 4, param = 0.7, model = "log")
taildep(dat, confint = 'wald')
## End(Not run)</pre>
```

test.maxstab

P-P plot for testing max stability

## **Description**

The diagnostic, proposed by Gabda, Towe, Wadsworth and Tawn, relies on the fact that, for max-stable vectors on the unit Gumbel scale, the distribution of the maxima is Gumbel distribution with a location parameter equal to the exponent measure. One can thus consider tuples of size m and estimate the location parameter via maximum likelihood and transforming observations to the standard Gumbel scale. Replicates are then pooled and empirical quantiles are defined. The number of combinations of m vectors can be prohibitively large, hence only nmax randomly selected tuples are selected from all possible combinations. The confidence intervals are obtained by a nonparametric bootstrap, by resampling observations with replacement observations for the selected tuples and re-estimating the location parameter. The procedure can be computationally intensive as a result.

## Usage

```
test.maxstab(
  xdat,
  m = prod(dim(dat)[-1]),
  nmax = 500L,
  B = 1000L,
  ties.method = "random",
  plot = TRUE,
  ...
)
```

## Arguments

matrix or array of max-stable observations, typically block maxima. The first dimension should consist of replicates

m integer indicating how many tuples should be aggregated.

max maximum number of pairs. Default to 500L.

B number of nonparametric bootstrap replications. Default to 1000L.

ties.method string indicating the method for rank. Default to "random".

plot logical indicating whether a graph should be produced (default to TRUE).

additional arguments for backward compatibility

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## Value

a Tukey probability-probability plot with 95% confidence intervals obtained using a nonparametric bootstrap

#### References

Gabda, D.; Towe, R. Wadsworth, J. and J. Tawn, Discussion of "Statistical Modeling of Spatial Extremes" by A. C. Davison, S. A. Padoan and M. Ribatet. *Statist. Sci.* **27** (2012), no. 2, 189–192.

## **Examples**

```
## Not run:
xdat <- mev::rmev(n = 250, d = 100, param = 0.5, model = "log")
test.maxstab(xdat, m = 100)
xdat <- rmnorm(n = 250, Sigma = diag(0.5, 10) + matrix(0.5, 10, 10), mu = rep(0, 10))
test.maxstab(xdat, m = 2, nmax = 100)
test.maxstab(xdat, m = ncol(xdat))
## End(Not run)</pre>
```

test.scoreindep

Ramos and Ledford test of independence

## **Description**

The Ramos and Ledford (2005) score test of independence is a modification of tests by Tawn (1988) and Ledford and Tawn (1996) for a logistic model parameter  $\alpha=1$ ; the latter two have scores with zero expectation, but the variance of the score are infinite, which produces non-regularity and yield test, once suitably normalized, that converge slowly to their asymptotic null distribution. The test, designed for bivariate samples, transforms observations to have unit Frechet margins and considers a bivariate censored likelihood approach for the logistic distribution.

## Usage

```
test.scoreindep(xdat, p, test = c("ledford", "tawn"))
```

# **Arguments**

xdat a n by 2 matrix of observations

p probability level for the marginal threshold

test string; if tawn, only censor observations in the upper quadrant when both vari-

ables are large as in Tawn (1988), otherwise censor marginally for ledford as

in Ledford and Tawn (1996).

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#### Value

```
a list with elements
stat value of the score test statistic
pval asymptotic p-value
test test argument
```

## **Examples**

```
samp <- rmev(n = 1000, d = 2,
    param = 0.99, model = "log")
(test.scoreindep(samp, p = 0.9))</pre>
```

thselect.alrs

Automatic L-moment ratio selection method

# Description

Given a sample of observations, calculate the L-skewness and L-kurtosis over a set of candidate thresholds. For each threshold candidate, we find the L-skewness that minimizes the sum of squared distance between the theoretical L-skewness and L-kurtosis of the generalized Pareto distribution,

$$\min_{\tau_3}(t_3-\tau_3)^2+[t_4-\tau_3(1+5\tau_3)/(5+\tau_3)]^2.$$

The function returns the threshold with the minimum distance.

## Usage

```
thselect.alrs(xdat, thresh, plot = FALSE)
```

## **Arguments**

xdat [numeric] vector of observations

thresh [numeric] vector of candidate thresholds. If missing, 20 sample quantiles start-

ing at the 0.25 quantile in increments of 3.75 percent.

plot [logical] if TRUE, return a plot of the sample L-kurtosis against the L-skewness,

along with the theoretical generalized Pareto curve.

#### Value

scalar for the chosen numeric threshold

#### References

Silva Lomba, J., Fraga Alves, M.I. (2020). *L-moments for automatic threshold selection in extreme value analysis*. Stoch Environ Res Risk Assess, 34, 465–491. doi:10.1007/s0047702001789x

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thselect.bab

Lower truncated Hill threshold selection

## **Description**

Given a sample of positive data with Pareto tail, the algorithm computes the optimal number of order statistics that minimizes the variance of the average left truncated tail index estimator, and uses the relationship to the Hill estimator for the Hall class of distributions to derive the optimal number (minimizing the asymptotic mean squared error) of the Hill estimator. The default value for the second order regular variation index is taken to be  $\rho = -1$ .

# Usage

```
thselect.bab(
  xdat,
  kmin = floor(0.2 * length(xdat)),
  kmax = length(xdat) - 1L,
  rho = -1,
  test = FALSE,
  nsim = 999L,
  level = 0.95
)
```

## **Arguments**

xdat	[vector] positive vector of exceedances
kmin	[int] minimum number of exceedances
kmax	[int] maximum number of exceedances for the estimation of the shape parameter.
rho	[double] scalar for the second order regular variation index, a negative number.
test	[logical] if TRUE, computes the goodness-of-fit statistic for the model using Monte Carlo
nsim	[int] number of replications for Monte Carlo test, used only if test=TRUE.
level	[double] confidence level for test

# Value

a list with the number of order statistics for the Hill estimator, k0 and the corresponding shape estimate shape, the average lower-trimmed Hill estimator shape.\_lth and the number of order statistics upon which the latter is based, k0\_lth.

#### References

Bladt, M., Albrecher, H. & Beirlant, J. (2020) *Threshold selection and trimming in extremes*. Extremes, 23, 629-665. doi:10.1007/s10687020003850

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Threshold selection by shape mean square error minimization

# Description

Use a semiparametric bootstrap to calculate the mean squared error of the shape parameter using maximum likelihood for different thresholds, and return the one that minimize the mean squared error.

## Usage

```
thselect.cbm(xdat, thresh, B = 100)
```

## **Arguments**

xdat vector of observations
thresh vector of thresholds

B number of bootstrap replications

#### Value

an object of class mev\_thselect\_cbm containing

- thresh: ordered vector of candidate thresholds
- thresh0: selected threshold
- shape: shape parameter coefficient estimates at each threshold
- nexc: number of exceedances at each threshold
- bias: vector of bootstrap bias estimates
- var: vector of bootstrap variance estimates
- mse: vector of mean squared error bootstrap estimates

## References

Caers, J., Beirlant, J. and Maes, M.A. (1999). Statistics for Modeling Heavy Tailed Distributions in Geology: Part I. Methodology. *Mathematical Geology*, 31, 391–410. <doi:10.1023/A:1007538624271>

```
set.seed(2025)
xdat <- rnorm(1000)
thresh <- qnorm(c(0.8, 0.9, 0.95))
thselect.cbm(xdat, thresh, B = 50)</pre>
```

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thselect.cv

Threshold selection via coefficient of variation

# Description

This function computes the empirical coefficient of variation and computes a weighted statistic comparing the squared distance with the theoretical coefficient variation corresponding to a specific shape parameter (estimated from the data using a moment estimator as the value minimizing the test statistic, or using maximum likelihood). The procedure stops if there are no more than 10 exceedances above the highest threshold.

## Usage

```
thselect.cv(
  xdat,
  thresh,
  method = c("mle", "wcv", "cv"),
  nsim = 999L,
  nthresh = 10L,
  level = 0.05,
  lazy = FALSE,
  plot = FALSE
)
```

## Arguments

xdat	[vector] vector of observations
thresh	[vector] vector of threshold. If missing, set to $p^k$ for $k=0$ to $k=$ nthresh
method	[string], either moment estimator for the (weighted) coefficient of variation (wcv and cv) or maximum likelihood ( $mle$ )
nsim	[integer] number of bootstrap replications
nthresh	[integer] number of thresholds, if thresh is not supplied by the user
level	[numeric] probability level for sequential testing procedure
lazy	[logical] compute the bootstrap p-value until the test stops rejecting at level level? Default to FALSE
plot	[logical] if TRUE, returns a plot of the p-value path

## Value

## a list with elements

- thresh: value of threshold returned by the procedure, NA if the hypothesis is rejected at all thresholds
- thresh0: sorted vector of candidate thresholds
- cindex: index of selected threshold among thresh0 or NA if none returned

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pval: bootstrap p-values, with NA if lazy and the p-value exceeds level at lower thresholds

- shape: shape parameter estimates
- nexc: number of exceedances of each threshold threshol
- method: estimation method for the shape parameter

#### Note

The authors suggest transformation of

$$Y = -1/(X + c) + 1/c,$$

where X are exceedances and  $c=\sigma/\xi$  is the ratio of estimated scale and shape parameters. For heavy-tailed distributions with  $\xi>0.25$ , this may be preferable, but must be conducted outside of the function.

#### References

del Castillo, J. and M. Padilla (2016). *Modelling extreme values by the residual coefficient of variation*, SORT, 40(2), pp. 303–320.

# **Examples**

```
thselect.cv(
  xdat = rgp(1000),
  thresh = qgp(seq(0,0.9, by = 0.1)),
  nsim = 99,
  lazy = TRUE,
  plot = TRUE)
```

thselect.egp

Threshold selection based on extended generalized Pareto models

# Description

Fit an EGP model to data over a range of candidate thresholds thresh and perform likelihood-based tests of equality for  $\kappa=c$ , where c=1 for all regular models and c=0 for the 'gj-tnorm' and 'logist' models, for which the generalized Pareto special case corresponds to a value of  $\kappa$  occurring on the boundary of the parameter space.

# Usage

```
thselect.egp(
  xdat,
  thresh,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
      "logist"),
  type = c("wald", "lrt"),
```

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```
level = 0.95,
transform = FALSE,
plot = FALSE,
...
)
```

#### **Arguments**

xdat vector of observations, greater than the threshold

thresh threshold value

model a string indicating which extended family to fit

type choice of test statistic, either wald for Wald-based intervals, or 1rt for profile

likelihood ratio test.

level [double] confidence interval level, default to 0.95.

transform logical; if TRUE and type="wald", intervals for kappa are computed on the log-

scale and back-transformed.

plot [logical] if TRUE, return a plot of p-values against threshold

... additional arguments, passed to plotting routine

#### **Details**

The threshold selection procedure returns chi-square statistics (stat) for Wald or profile likelihood ratio tests, along with p-values (pval) obtained from large sample distribution. The threshold returned is the lowest for which all further higher thresholds fail to reject the null hypothesis of  $\kappa=c$ , or equivalently of generalized Pareto tail.

#### Value

an invisible list of class mev\_thselect\_egp with elements

- thresh: vector of threshold candidates
- thresh0: selected threshold among candidates
- coef: vector of parameter estimates for  $\kappa$
- stat: squared version of the test statistic
- pval: p-value obtained from the  $\chi_1^2$  approximation
- · level: level of the confidence intervals
- model: string giving the EGP model family
- type: type of confidence interval

```
ths <- thselect.egp(
  xdat = rexp(1000),
  thresh = qexp(c(0.8,0.9,0.95)),
  model = "pt-power")
print(ths)
plot(ths)</pre>
```

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thselect.expgqt

Generalized quantile threshold selection

## **Description**

The methodology proposed by Beirlant, Vynckier and Teugels (1996) uses an asymptotic expansion of the mean squared error for Hill's estimator given a random sample with Pareto tails and positive shape, using an exponential regression. The value of k is selected to minimize the mean squared error given optimal weighting scheme. This depends on the order of regular variation  $\rho$ , which is obtained based on the slope of the difference in Hill estimators, suitably reweighted. The iterative procedure of Beirlant et al. alternates between parameter estimation until convergence. It returns the Hill shape estimate, the number of higher order statistic, the parameter rho and estimates of the standard error of the shape and the mean squared error, based on the ultimate parameter values. Since the weights can become negative, there is no guarantee that the mean squared error estimate is positive, nor that the estimated value of  $\rho$  is nonpositive.

## Usage

```
thselect.expgqt(
  xdat,
  maxiter = 10L,
  tol = 2,
  kmin = max(10, floor(length(xdat)/100)),
  kmax = floor(0.8 * length(xdat)),
  ...
)
```

#### **Arguments**

#### Value

a list with components

- shape the Hill estimator of the shape, based on the k largest order statistics
- k0 number of high order statistics for estimation of the shape using Hill's estimator
- rho estimate of the second order regular variation parameter
- mse mean squared error estimate of the shape parameter
- se standard error of the shape parameter
- convergence logical; if TRUE, indicates that the method converged to a fixed point within tol before reaching the maximum number of iterations maxiter

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## References

Beirlant, J., Vynckier, P., & Teugels, J. L. (1996). Excess Functions and Estimation of the Extreme-Value Index. *Bernoulli*, 2(4), 293–318. doi:10.2307/3318416

## **Examples**

```
# Simulate Pareto data - log(xdat) is exponential with rate 2
xdat <- rgp(n = 200, loc = 1, scale = 0.5, shape = 0.5)
(thselect.expgqt(xdat))</pre>
```

thselect.gbw

Kernel-based threshold selection of Goegebeur, Beirlant and de Wet (2008)

# Description

Kernel-based threshold selection of Goegebeur, Beirlant and de Wet (2008)

## Usage

```
thselect.gbw(
  xdat,
  kmax,
  kernel = c("Jackson", "Lewis"),
  rho = c("gbw", "ghp", "fagh", "dk"),
  ...
)
```

## **Arguments**

xdat [vector] sample exceedances

kmax [int] maximum number of exceedances considered kernel [string] kernel choice, one of Jackson or Lewis

rho string for the estimator of the second order regular variation. Can also be a

negative scalar

. . . additional arguments, for backward compatibility purposes

## Value

a list with elements

- k0: number of exceedances
- shape: Hill's shape estimate
- rho: second-order regular variation parameter estimate
- gof: goodness-of-fit statistic for the chosen threshold.

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#### References

Goegebeur, Y., Beirlant, J., and de Wet, T. (2008). Linking Pareto-Tail Kernel Goodness-of-fit Statistics with Tail Index at Optimal Threshold and Second Order Estimation. REVSTAT-Statistical Journal, 6(1), 51–69. <doi:10.57805/revstat.v6i1.57>

# **Examples**

```
xdat <- rgp(n = 1000, scale = 2, shape = 0.5)
(thselect.gbw(xdat, kmax = 500))</pre>
```

thselect.ksmd

Mahalanobis distance-based methodology

## **Description**

Compute the Mahalanobis distance-based threshold method over a grid of thresholds by transforming data from generalized Pareto to unit exponential based on probability weighted moment estimates, then computing the first L-moment and the L-skewness. The latter are compared to the theoretical counterparts from a unit exponential sample of the same size, which is used to compute the Mahalanobis distance. The threshold returned is the one which minimizes the distance.

## Usage

```
thselect.ksmd(xdat, thresh, approx = c("asymptotic", "mc"), nsim = 1000L)
```

# **Arguments**

xdat	[numeric] vector of observations
thresh	[numeric] vector of candidate thresholds. If missing, 20 sample quantiles starting at the 0.25 quantile in increments of 3.75 percent.
approx	[string] method to use to obtain moments of first L-moment
nsim	[integer] number of replications for Monte Carlo approximation

#### Value

a list with components

- thresh0: selected threshold returned by the procedure
- · thresh: vector of candidate thresholds
- pval: scalar *p*-value for the chi-square approximation to the test statistic for the selected threshold
- dist: vector of Mahalanobis distance
- approx: type of approximation

## References

Kiran, K. G. and Srivinas, V.V. (2021). A Mahalanobis distance-based automatic threshold selection method for peaks over threshold model. Water Resources Research 57. <doi:10.1029/2020WR027534>

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thselect.mdps

Minimum distance threshold selection procedure

## **Description**

Minimum distance threshold selection procedure

## Usage

```
thselect.mdps(xdat)
```

## **Arguments**

xdat

vector of positive exceedances

## Value

a list with components

k0 order statistic corresponding to threshold (number of exceedances)

shape Hill's estimator of the tail index based on k0 exceedances

thresh0 numerical value of the threshold, the n-k0+1 order statistic of the original sample

# References

Clauset, A., Shalizi, C.R. and Newman, M.E.J. (2009). *Power-Law Distributions in Empirical Data*. SIAM Review. Society for Industrial and Applied Mathematics, **51**, 661-703, doi:10.1137/070710111

thselect.mrl

Automated mean residual life plots

# **Description**

This function implements the automated proposal from Section 2.2 of Langousis et al. (2016) for mean residual life plots. It returns the threshold that minimize the weighted mean square error and moment estimators for the scale and shape parameter based on weighted least squares.

# Usage

```
thselect.mrl(xdat, thresh, kmax, plot = TRUE, ...)
```

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# **Arguments**

xdat	[numeric] vector of observations
thresh	[numeric] vector of thresholds; if missing, uses all order statistics from the 20th largest until kmax as candidates
kmax	[integer] maximum number of order statistics
plot	[logical] if TRUE (default), return a plot of the mean residual life plot with the fitted slope and the chosen threshold $\frac{1}{2}$
	additional arguments, currently ignored

#### **Details**

The procedure consists in estimating the usual mean residual life as a function of the threshold, and looking for an order statistic or threshold value above which the fit is more or less linear.

#### Value

a list containing

• thresh: candidate threshold vector

• thresh0: selected threshold

• scale: scale parameter estimate

• shape: shape parameter estimate

• mrl: empirical mean excess values

xdat: ordered observations

• intercept: intercept for mean excess value at chosen threshold

• slope: slope for mean excess value at chosen threshold

• tmanual: logical; TRUE if the user passed a vector of thresholds

#### References

Langousis, A., A. Mamalakis, M. Puliga and R. Deidda (2016). Threshold detection for the generalized Pareto distribution: Review of representative methods and application to the NOAA NCDC daily rainfall database, Water Resources Research, **52**, 2659–2681.

```
thselect.mrl(rgp(n = 100))
```

thselect.pec

tion diagnostic	thselect.ncpgp	Northop and Coleman piecewise generalized Pareto threshold selection diagnostic
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## **Description**

The model tests the null hypothesis of a generalized Pareto above each threshold in thresh against the alternative of a piecewise generalized Pareto model with continuity constraints.

## Usage

```
thselect.ncpgp(xdat, thresh, test = "score", plot = FALSE, level = 0.95, ...)
```

## **Arguments**

xdat	[vector] observations
thresh	[vector] candidate thresholds
test	[string] indicating whether to perform score test or likelihood ratio (1r) test. The latter requires fitting the alternative model, and so is more computationally expensive.
plot	[logical]; if TRUE, return a plot with the p-value path.
level	[double] confidence level for confidence interval, defaults to 0.95
	additional arguments, for backward compatibility purposes

## Value

an object of class mev\_thselect\_ncpgp containing the test statistic (stat), the p-values (pval), the threshold candidates (thresh) and the selected threshold (thresh0).

thselect.pec Prediction error C-criterion threshold selection method	
--	--

# Description

This function computes the non-robust Pareto prediction error of Dupuis and Victoria-Feser (2003), termed C-criterion, for the Hill estimator of the shape parameter. The threshold returned is the value of the threshold, taken from order statistics, that minimizes the average prediction error.

# Usage

```
thselect.pec(xdat, kmax)
```

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#### **Arguments**

xdat vector of observations

kmax maximum number of order statistics to consider. Default to sample size if left

unspecified.

## Value

a list with the number of exceedances k, the chosen threshold thresh0 and the corresponding Hill estimator shape estimate shape.

#### References

Dupuis, D.J. and M.-P. Victoria-Feser (2003). A Prediction Error Criterion for Choosing the Lower Quantile in Pareto Index Estimation, University of Geneva, technical report, https://archive-ouverte.unige.ch/unige:5789.

thselect.pickands

Pickands' order statistics threshold selection method

## **Description**

Restricting to the largest fourth of the data, returns the number of exceedances that minimizes the Kolmogorov-Smirnov statistic, i.e., the maximum absolute difference between the estimated generalized Pareto and the empirical distribution of exceedances. Relative to the paper, different estimation methods are proposed.

## Usage

```
thselect.pickands(xdat, thresh, method = c("mle", "lmom", "quartiles"))
```

# **Arguments**

xdat [numeric] vector of observations

thresh [numeric] vector of candidate thresholds. If missing, defaults to order statistics

from the 10th to a quarter of the sample size.

method [string] estimation method, either the quartiles of Pickands (1975), maximum

likelihood, probability weighted moments or L-moments

#### Value

a list with components

- k0: number of exceedances
- thresh0: selected threshold returned by the procedure
- thresh: vector of candidate thresholds
- · dist; vector of Kolmogorov-Smirnoff distance

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- method; string for the estimation method
- scale: estimated scale parameter at the chosen threshold
- shape: estimated shape parameter at the chosen threshold

#### Note

The quartiles estimator of Pickands is robust, but very inefficient. It is provided for historical reasons.

#### References

James Pickands III (1975). *Statistical inference using extreme order statistics*, Annals of Statistics, 3(1) 119-131, doi:10.1214/aos/1176343003

thselect.rbm

Threshold selection for the random block maxima method

## **Description**

Threshold selection for the random block maxima method

## Usage

```
thselect.rbm(xdat, kmax = length(xdat))
```

# **Arguments**

xdat [vector] sample exceedances

kmax maximum number of exceedances to consider.

## Value

a list with elements

- k0: the number of exceedances at the selected threshold
- thresh $\theta$ : the selected threshold, or accordingly the  $(k_0+1)$ th order statistic
- shape: the RBM shape estimate

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thselect.samsee

Threshold selection via SAMSEE

## **Description**

Smooth asymptotic mean squared error estimator of Schneider et al. (2021) for threshold selection. The implementation uses a second-order regular variation index of -1

## Usage

```
thselect.samsee(xdat)
```

## **Arguments**

xdat

vector of positive exceedances

#### Value

a list with elements

k0 optimal number of exceedances shape Hill estimator of the tail index thresh0 selected threshold

# References

Schneider, L.F., Krajina, A. & Krivobokova, T. (2021). *Threshold selection in univariate extreme value analysis*, Extremes, **24**, 881–913 doi:10.1007/s10687021004057

thselect.sdinfo

Threshold selection diagnostic of Suveges and Davison

# **Description**

The information matrix test (IMT), proposed by Suveges and Davison (2010), is based on the difference between the expected quadratic score and the second derivative of the log-likelihood. The asymptotic distribution for each threshold u and gap K is asymptotically  $\chi^2$  with one degree of freedom. The approximation is good for N>80 and conservative for smaller sample sizes. The test assumes independence between gaps.

# Usage

```
thselect.sdinfo(xdat, thresh, qlev, plot = FALSE, kmax = 1, k = 1)
```

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## **Arguments**

xdat	[vector] vector of observations
thresh	[vector] candidate thresholds
qlev	[vector] probability levels to define threshold if thresh is missing.
plot	[logical]; should the graphical diagnostic be plotted?
kmax	[int] the largest K-gap under consideration for clusters
k	[int] the K-gap for automatic threshold selection

#### **Details**

The procedure proposed in Suveges & Davison (2010) was corrected for erratas. The maximum likelihood is based on the limiting mixture distribution of the intervals between exceedances (an exponential with a point mass at zero). The condition  $D^{(K)}(u_n)$  should be checked by the user.

Fukutome et al. (2015) propose an ad hoc automated procedure

- 1. Calculate the interexceedance times for each K-gap and each threshold, along with the number of clusters
- 2. Select the (u, K) pairs for which IMT < 0.05 (corresponding to a P-value of 0.82)
- 3. Among those, select the pair (u, K) for which the number of clusters is the largest

#### Value

an invisible list of class with elements

- thresh a vector of thresholds based on empirical quantiles at supplied levels.
- stat a matrix of test statistics
- pval a matrix of approximate p-values (corresponding to probabilities under a  $\chi_1^2$  distribution)
- mle a matrix of maximum likelihood estimates for each given pair of thresholds and gaps
- loglik a matrix of log-likelihood values at MLE for each given pair of elements in thresh and gap in  $0, \ldots, k$ max
- quantile quantile levels for thresholds, if supplied by the user
- kmax the largest gap number

## Author(s)

Leo Belzile

#### References

Fukutome, Liniger and Suveges (2015), Automatic threshold and run parameter selection: a climatology for extreme hourly precipitation in Switzerland. *Theoretical and Applied Climatology*, **120**(3), 403-416.

Suveges and Davison (2010), Model misspecification in peaks over threshold analysis. *Annals of Applied Statistics*, **4**(1), 203-221.

White (1982), Maximum Likelihood Estimation of Misspecified Models. *Econometrica*, **50**(1), 1-25.

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## **Examples**

```
thselect.sdinfo(
  xdat = rgp(n = 10000),
  qlev = seq(0.1, 0.9, length = 10),
  kmax = 3)
```

thselect.vmetric

Metric-based threshold selection

# Description

Adaptation of Varty et al.'s metric-based threshold automated diagnostic for the independent and identically distributed case with no rounding.

# Usage

```
thselect.vmetric(
   xdat,
   thresh,
   B = 199L,
   type = c("eqd", "exp", "qq", "pp"),
   dist = c("l1", "l2"),
   uq = FALSE,
   neval = 1000L,
   level = 0.95,
   plot = FALSE,
   ...
)
```

## **Arguments**

xdat	vector of observations
thresh	vector of thresholds
В	number of bootstrap replications
type	string indicating scale, either exp for exponential quantile-quantile plot as in Varty et al. (2021) or pp for probability-probability plot (uniform). The method qq or expected quantile discrepancy (eqd) corresponds to Murphy et al. (2024) with the quantiles on the generalized Pareto scale.
dist	string indicating norm, either 11 for absolute error or 12 for quadratic error
uq	logical; if TRUE, generate bootstrap samples accounting for the sampling distribution of parameters $$
neval	number of points at which to estimate the metric. Default to 1000L
level	level of symmetric confidence interval. Default to 0.95
plot	logical; if TRUE, returns a plot
	additional arguments for backward compatibility

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#### **Details**

The algorithm proceeds by first computing the maximum likelihood algorithm and then simulating datasets from replication with parameters drawn from a bivariate normal approximation to the maximum likelihood estimator distribution.

For each bootstrap sample, we refit the model and convert the quantiles to exponential or uniform variates. The mean absolute or mean squared distance is calculated on these. The threshold returned is the one with the lowest value of the metric.

## Value

an invisible list with components

- thresh: scalar threshold minimizing criterion
- thresh0: vector of candidate thresholds
- metric: value of the metric criterion evaluated at each threshold
- type: argument type
- dist: argument dist,
- level: level of confidence interval, from level

#### References

Varty, Z. and J.A. Tawn and P.M. Atkinson and S. Bierman (2021+), Inference for extreme earthquake magnitudes accounting for a time-varying measurement process.

Murphy, C., Tawn, J. A., & Varty, Z. (2024). *Automated Threshold Selection and Associated Inference Uncertainty for Univariate Extremes*. Technometrics, 67(2), 215–224. <a href="https://doi.org/10.1080/00401706.2024.2421744">doi:10.1080/00401706.2024.2421744</a>

## **Examples**

```
xdat <- rexp(1000, rate = 1/2)
thresh <- quantile(xdat, prob = c(0.25,0.5, 0.75))
thv <- thselect.vmetric(xdat, thresh, B = 99)
plot(thv)
print(thv)</pre>
```

thselect.wseq

Wadsworth's sequential analysis threshold selection

# Description

Function to produce diagnostic plots and test statistics for the threshold diagnostics exploiting structure of maximum likelihood estimators based on the non-homogeneous Poisson process likelihood or the coefficient of tail dependence

thselect.wseq 129

## Usage

```
thselect.wseq(
  xdat,
  thresh,
  qlev,
  model = c("nhpp", "taildep", "rtaildep"),
  npp = 1,
  nsim = 1000L,
  level = 0.95,
  plot = FALSE,
  ...
)
```

## **Arguments**

xdat a numeric vector or matrix of data to be fitted. vector of candidate thresholds. thresh vector of probabilities for empirical quantiles used in place of the threshold, alev used if argument thresh is missing. mode1 string specifying whether the univariate or multivariate diagnostic should be used. Either nhpp for the univariate model, or exp (invexp) for the bivariate exponential model with rate (inverse rate) parametrization. See details. number of observations per period for the non-homogeneous point process model. npp Default to 1. number of Monte Carlo simulations used to assess the null distribution of the nsim test statistic level confidence level of intervals, defaults to 0.95 logical; if TRUE, calls the plot routine plot additional parameters passed to internal routine . . .

#### **Details**

The function is a wrapper for the univariate (non-homogeneous Poisson process model) and exponential dependence model applied to the minimum component (tail dependence coefficient). For the latter, the user can select either the rate ("taildep" or inverse rate parameter ("rtaildep"). The inverse rate parametrization works better for uniformity of the p-value distribution under the likelihood ratio test for the changepoint.

For the coefficient of tail dependence, users must provide pairwise minimum of marginally exponentially distributed margins (see example)

#### Value

an object of class invisible list with components

- thresh0: threshold selected by the likelihood ratio procedure
- thresh: vector of candidate thresholds

thselect.wseq

- coef: maximum likelihood estimates from all thresholds
- vcov: joint asymptotic covariance matrix for shape  $\xi$  or coefficient of tail dependence  $\eta$ , or it's reciprocal.
- wn: values of the white noise process
- stat: value of the likelihood ratio test statistic for the changepoint test
- pval: P-value of the likelihood ratio test
- mle: maximum likelihood estimates for the selected threshold
- model: model fitted, either nhpp, exp or invexp
- nsim: number of Monte Carlo simulations for changepoint test
- xdat: vector of observations

## Author(s)

Jennifer L. Wadsworth, Léo Belzile

#### References

Wadsworth, J.L. (2016). Exploiting Structure of Maximum Likelihood Estimators for Extreme Value Threshold Selection, *Technometrics*, **58**(1), 116-126, http://dx.doi.org/10.1080/00401706.2014.998345.

```
## Not run:
set.seed(123)
xdat <- abs(rnorm(5000))</pre>
thresh <- quantile(xdat, seq(0, 0.9, by = 0.1))
(diag <- thselect.wseq(</pre>
xdat = xdat,
thresh = thresh,
plot = TRUE,
type = "ps")
# Multivariate example, with coefficient of tail dependence
xbvn <- rmnorm(n = 6000L,
                mu = rep(0, 2),
                Sigma = cbind(c(1, 0.7), c(0.7, 1)))
thselect.wseq(
 xdat = xbvn,
 qlev = seq(0, 0.9, length.out = 30),
 model = 'taildep',
 plot = TRUE)
## End(Not run)
```

tstab.cv 131

tstab.cv

Coefficient of variation threshold stability plot

# Description

This function calculates parametric estimates of the coefficient of variation with pointwise Wald confidence intervals along with empirical estimates and returns a threshold stability plot.

# Usage

```
tstab.cv(
  xdat,
  thresh,
  method = c("empirical", "mle", "wcv", "cv"),
  nthresh = 10L,
  nsim = 99L,
  plot = TRUE,
  level = 0.95,
  ...
)
```

# Arguments

xdat	[vector] vector of observations
thresh	[vector] vector of threshold. If missing, set to $p^k$ for $k=0$ to $k=$ nthresh
method	[string], either moment estimator for the (weighted) coefficient of variation (wcv and cv) or maximum likelihood ( $mle$ )
nthresh	[integer] number of thresholds, if thresh is not supplied by the user
nsim	[integer] number of bootstrap replications
plot	[logical] if TRUE, returns a plot of the p-value path
level	[numeric] probability level for sequential testing procedure
• • •	additional parameters, notably for package boot, for the type of confidence intervals.

```
tstab.cv(
    xdat = rgp(1000),
    thresh = qgp(seq(0,0.9, by = 0.1)),
    method = "cv")
tstab.cv(
    xdat = rgp(1000),
    thresh = qgp(seq(0,0.9, by = 0.1)),
    method = "empirical")
```

132 tstab.egp

tstab.egp

Threshold stability plots for extended generalized Pareto models

## **Description**

Threshold stability plots for extended generalized Pareto models

# Usage

```
tstab.egp(
  xdat,
  thresh,
  model = c("pt-beta", "pt-gamma", "pt-power", "gj-tnorm", "gj-beta", "exptilt",
      "logist"),
  param = c("shape", "kappa"),
  type = c("wald", "lrt"),
  transform = FALSE,
  level = 0.95,
  plot = TRUE,
  ...
)
```

## **Arguments**

xdat	vector of observations, greater than the threshold
thresh	threshold value
model	a string indicating which extended family to fit
param	[string] parameter, either shape or additional parameter kappa
type	[string] confidence interval type, either wald or profile.
transform	logical; if TRUE and type="wald", intervals for kappa are computed on the log-scale and back-transformed.
level	[double] confidence interval level, default to 0.95.
plot	[logical] if TRUE (default), return a threshold stability plot
	additional arguments for the plot function, currently ignored.

# Value

an invisible list object of class mev\_egp\_tstab with elements

- kappa: matrix of parameter estimates and confidence intervals for  $\kappa$ , if specified in param
- shape: matrix of parameter estimates and confidence intervals for the shape parameter  $\xi$ , if specified in param
- thresh: vector of threshold candidates
- level: level of the confidence intervals
- model: string giving the EGP model family
- type: type of confidence interval

tstab.gpd 133

## **Examples**

```
xdat <- rgp(n = 1000)
tstab.egp(
  xdat = xdat,
  thresh = c(0, quantile(xdat, 0.5)),
  model = "gj-tnorm",
  param = "kappa",
  transform = TRUE)</pre>
```

tstab.gpd

Parameter stability plots for peaks-over-threshold

## **Description**

This function computes the maximum likelihood estimate at each provided threshold and plots the estimates (pointwise), along with 95% confidence/credible intervals obtained using Wald or profile confidence intervals, or else from 1000 independent draws from the posterior distribution under vague independent normal prior on the log-scale and shape. The latter two methods better reflect the asymmetry of the estimates than the Wald confidence intervals.

# Usage

```
tstab.gpd(
  xdat,
  thresh,
  method = c("wald", "lrt", "post"),
  level = 0.95,
  plot = TRUE,
  which = c("scale", "shape"),
  changepar = TRUE,
  ...
)
```

# **Arguments**

xdat	a vector of observations
thresh	a vector of candidate thresholds at which to compute the estimates.
method	string indicating the method for computing confidence or credible intervals. Must be one of "wald", "profile" or "post".
level	confidence level of the intervals. Default to 0.95.
plot	logical; should parameter stability plots be displayed? Default to TRUE.
which	character vector with elements scale or shape
changepar	logical; if TRUE, changes the graphical parameters.
	additional arguments passed to plot.

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## Value

a list with components

- threshold: vector of numerical threshold values.
- mle: matrix of modified scale and shape maximum likelihood estimates.
- lower: matrix of lower bounds for the confidence or credible intervals.
- upper: matrix of lower bounds for the confidence or credible intervals.
- method: method for the confidence or coverage intervals.

plots of the modified scale and shape parameters, with pointwise confidence/credible intervals and an invisible data frame containing the threshold thresh and the modified scale and shape parameters.

## Note

The function is hard coded to prevent fitting a generalized Pareto distribution to samples of size less than 10. If the estimated shape parameters are all on the boundary of the parameter space (meaning  $\hat{\xi}=-1$ ), then the plots return one-sided confidence intervals for both the modified scale and shape parameters: these typically suggest that the chosen thresholds are too high for estimation to be reliable.

## Author(s)

Leo Belzile

## See Also

```
gpd.fitrange
```

```
dat <- abs(rnorm(10000))
u <- qnorm(seq(0.9,0.99, by= 0.01))
par(mfrow = c(1,2))
tstab.gpd(xdat = dat, thresh = u, changepar = FALSE)
## Not run:
tstab.gpd(xdat = dat, thresh = u, method = "lrt")
tstab.gpd(xdat = dat, thresh = u, method = "post")
## End(Not run)</pre>
```

tstab.hill 135

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Threshold stability plot for Hill estimator

# Description

Threshold stability plot for Hill estimator

# Usage

```
tstab.hill(xdat, kmax, method = "hill", ..., log = TRUE)
```

## **Arguments**

xdat	[vector] sample exceedances
kmax	[int] maximum number of order statistics
method	[string] name of estimator for shape parameter. Default to hill.
	additional arguments passed to fit. shape for certain methods.
log	[logical] should the x-axis for the number of order statistics used for estimation be displayed on the log scale? Default to TRUE

## Value

a plot of shape estimates as a function of the number of exceedances

## **Examples**

```
xdat <- rgp(n = 250, loc = 1, scale = 2, shape = 0.5)
tstab.hill(xdat)</pre>
```

tstab.lthill

Threshold stability plots for left-truncated Hill estimators

## **Description**

Given a vector of exceedances and some potential choices of k for the threshold, compute the left-truncated Hill estimators for each value of k and use these to compute the variance and slope of the estimator

# Usage

```
tstab.lthill(xdat, k, which = c("lthill", "var", "slope"), log = TRUE, ...)
```

136 tstab.mrl

## **Arguments**

xdat	[numeric] vector of positive observations
k	[integer] number of order statistics for the threshold
which	[string] the type of plot, showing the left-truncated Hill plot on the log, the log of the variance of the estimator, or the log slope
log	[logical] if TRUE (default), shows the Hill plot on the log-scale
	additional parameters for color, etc. to be passed to plot

## Value

an invisible list with lthill, order statistics, the log variance and the log scale.

#### References

Bladt, M., Albrecher, H. & Beirlant, J. (2020) *Threshold selection and trimming in extremes*. Extremes, 23, 629-665. doi:10.1007/s10687020003850

## **Examples**

```
xdat <- 10/(1 - runif(n = 1000)) - 10

tstab.lthill(xdat = xdat, k = c(50,100,200))
```

tstab.mrl

Mean residual life plot

## **Description**

Computes mean of sample exceedances over a range of thresholds or for a pre-specified number of largest order statistics, and returns a plot with 95% Wald-based confidence intervals as a function of either the threshold or the number of exceedances. The main purpose is the plotting method, which generates the so-called mean residual life plot. The latter should be approximately linear over the threshold for a generalized Pareto distribution

# Usage

```
tstab.mrl(
  xdat,
  thresh,
  kmin = 10L,
  kmax = length(xdat),
  plot = TRUE,
  level = 0.95,
  xlab = c("thresh", "nexc"),
  type = c("band", "ptwise"),
  ...
)
```

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## **Arguments**

xdat	vector of sample observations
thresh	vector of thresholds
kmin	integer giving the minimum number of exceedances; ignored if thresh is provided. Default to $10$
kmax	integer giving the maximum number of exceedances; ignored if thresh is provided. Default to sample size.
plot	logical; if TRUE, call the plot method
level	double giving the level of confidence intervals for the plot, default to 0.95
xlab	string indicating whether to use thresholds (thresh) or number of largest order statistics (nexc) for the x-axis
type	string whether to plot pointwise confidence intervals using segments ("ptwise") or using dashed lines ("band")
	additional arguments, currently ignored

#### Value

an invisible list with mean sample exceedances and standard deviation, number of exceedances, threshold

#### References

Davison, A.C. and R.L. Smith (1990). Models for Exceedances over High Thresholds (with discussion), *Journal of the Royal Statistical Society. Series B (Methodological)*, **52**(3), 393–442.

# **Examples**

```
tstab.mrl(
  xdat = rgp(n = 100, shape = -0.5),
  xlab = "thresh",
  kmax = 50)
tstab.mrl(
  rexp(100),
  thresh = qexp(seq(0, 0.9, by = 0.01)))
```

venice Venice Sea Levels

# **Description**

The venice data contains the 10 largest yearly sea levels (in cm) from 1887 until 2019. Only the yearly maximum is available for 1922 and the six largest observations for 1936.

w1500m

#### **Format**

a data frame with 133 rows and 11 columns containing the year of the measurement (first column) and ordered 10-largest yearly observations, reported in decreasing order from largest (r1) to smallest (r10).

#### Note

Smith (1986) notes that the annual maxima seems to fluctuate around a constant sea level up to 1930 or so, after which there is potential linear trend. Records of threshold exceedances above 80 cm (reported on the website) indicate that observations are temporally clustered.

The observations from 1931 until 1981 can be found in Table 1 in Smith (1986), who reported data from Pirazzoli (1982). The values from 1983 until 2019 were extracted by Anthony Davison from the City of Venice website (accessed in May 2020) and are licensed under the CC BY-NC-SA 3.0 license. The Venice City website indicates that later measurements were recorded by an instrument located in Punta Salute.

#### Source

City of Venice, Historical archive <a href="https://www.comune.venezia.it/node/6214">https://www.comune.venezia.it/node/6214</a>>. Last accessed November 5th, 2020.

#### References

Smith, R. L. (1986) Extreme value theory based on the *r* largest annual events. *Journal of Hydrology* **86**, 27–43.

Pirazzoli, P., 1982. Maree estreme a Venezia (periodo 1872-1981). Acqua Aria 10, 1023-1039.

Coles, S. G. (2001) An Introduction to Statistical Modelling of Extreme Values. London: Springer.

#### See Also

venice

w1500m

Best 200 times of Women 1500m Track

#### **Description**

200 all-time best performance (in seconds) of women 1500-meter run.

#### Format

a vector of size 200

#### Source

<a href="http://www.alltime-athletics.com/w\_1500ok.htm">http://www.alltime-athletics.com/w\_1500ok.htm</a>, accessed 14.08.2018

xdep.asym 139

xdep.asym

Coefficient of extremal asymmetry

# Description

This function implements estimators of the bivariate coefficient of extremal asymmetry proposed in Semadeni's (2021) PhD thesis. Two estimators are implemented: one based on empirical distributions, the second using empirical likelihood.

## Usage

```
xdep.asym(
  xdat,
  qlev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  estimator = c("emp", "elik"),
  confint = c("none", "wald", "bootstrap"),
  level = 0.95,
  B = 999L,
  ties.method = "random",
  plot = TRUE,
  ...
)
```

## **Arguments**

xdat	an n by 2 matrix of observations
qlev	vector of quantile levels at which to evaluate extremal asymmetry
nq	integer; number of quantiles at which to evaluate the coefficient if u is NULL
qlim	a vector of length 2 with the probability limits for the quantiles
estimator	string indicating the estimation method, one of emp or empirical likelihood (elik)
confint	string for the method used to derive confidence intervals, either none (default) or a nonparametric bootstrap
level	probability level for confidence intervals, default to $0.95$ or bounds for the interval
В	integer; number of bootstrap replicates (if applicable)
ties.method	string; method for handling ties. See the documentation of rank for available options.
plot	logical; if TRUE, return a plot.
	additional arguments for backward compatibility

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## **Details**

Let U, V be uniform random variables and define the partial extremal dependence coefficients

$$\varphi_{+}(u) = \Pr(V > U \mid U > u, V > u),$$

,

$$\varphi_{-}(u) = \Pr(V < U \mid U > u, V > u),$$
  
$$\varphi_{0}(u) = \Pr(V = U \mid U > u, V > u).$$

Define

$$\varphi(u) = \frac{\varphi_+ - \varphi_-}{\varphi_+ + \varphi_-}$$

and the coefficient of extremal asymmetry as  $\varphi = \lim_{u \to 1} \varphi(u)$ .

The empirical likelihood estimator, derived for max-stable vectors with unit Frechet margins, is

$$\widehat{\varphi}_{\text{el}} = \frac{\sum_{i} p_{i} I(w_{i} \le 0.5) - 0.5}{0.5 - 2 \sum_{i} p_{i} (0.5 - w_{i}) I(w_{i} \le 0.5)}$$

where  $p_i$  is the empirical likelihood weight for observation i, I is an indicator function and  $w_i$  is the pseudo-angle associated to the first coordinate, derived based on exceedances above u.

#### Value

an invisible data frame with columns

qlev quantile level of thresholds

coef extremal asymmetry coefficient estimates

lower either NULL or a vector containing the lower bound of the confidence interval upper either NULL or a vector containing the lower bound of the confidence interval

## References

Semadeni, C. (2020). Inference on the Angular Distribution of Extremes, PhD thesis, EPFL, no. 8168.

xdep.chi 141

xdep.chi

Coefficient of tail correlation

# Description

The coefficient of tail correlation  $\chi$  is

$$\chi = \lim_{u \to 1} \frac{\Pr(F_1(X_1) > u, \dots, F_D(X_D) > u)}{1 - u}.$$

Asymptotically independent vectors have  $\chi=0.$  The estimator uses an estimator of the survival copula

# Usage

```
xdep.chi(
   xdat,
   qlev = NULL,
   nq = 40,
   qlim = c(0.8, 0.99),
   estimator = c("emp", "betacop", "gpd", "hill"),
   confint = c("wald", "lrt"),
   level = 0.95,
   margtrans = c("emp", "none"),
   ties.method = "random",
   plot = TRUE,
   ...
)
```

## **Arguments**

xdat	an $n$ by $d$ matrix of multivariate observations
qlev	vector of percentiles between 0 and 1
nq	number of quantiles of the structural variable at which to form a grid; only used if $u = NULL$ .
qlim	limits for the sequence u of the structural variable
estimator	string giving estimator to employ
confint	string indicating the type of confidence interval, one of "wald" or "lrt"
level	the confidence level required (default to 0.95).
margtrans	string giving the marginal transformation, one of emp for rank-based transformation or none if data are already on the uniform scale
ties.method	string indicating the type of method for rank; see rank for a list of options. Default to "random"
plot	logical; if TRUE, return a plot
	additional arguments to taildep, currently ignored

142 xdep.chibar

## Value

a data frame

• qlev: quantile level of estimates

· coef: point estimates

• lower: lower bound of confidence interval

• upper: lower bound of confidence interval

## **Examples**

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 2, param = 0.7, model = "log")
xdep.chi(dat, confint = 'wald')
## End(Not run)</pre>
```

xdep.chibar

Coefficient chi-bar

## **Description**

For data with unit Pareto margins, the coefficient  $\bar{\chi} = 2\eta - 1$  is defined via

$$\Pr(\min(X) > x) = L(x)x^{-1/\eta},$$

where L(x) is a slowly varying function. Ignoring the latter, several estimators of  $\eta$  can be defined. In unit Pareto margins,  $\eta$  is a nonnegative shape parameter that can be estimated by fitting a generalized Pareto distribution above a high threshold. In exponential margins,  $\eta$  is a scale parameter and the maximum likelihood estimator of the latter is the Hill estimator. Both methods are based on peaks-over-threshold and the user can choose between pointwise confidence obtained through a likelihood ratio test statistic ("1rt") or the Wald statistic ("wald").

# Usage

```
xdep.chibar(
  xdat,
  qlev = NULL,
  nq = 40,
  qlim = c(0.8, 0.99),
  estimator = c("emp", "betacop"),
  confint = c("wald", "lrt"),
  level = 0.95,
  margtrans = c("emp", "none"),
  ties.method = "random",
  plot = TRUE,
  ...
)
```

xdep.chibar 143

## **Arguments**

xdat an n by d matrix of multivariate observations

qlev vector of percentiles between 0 and 1

nq number of quantiles of the structural variable at which to form a grid; only used

if u = NULL.

qlim limits for the sequence u of the structural variable

estimator string giving estimator to employ

confint string indicating the type of confidence interval, one of "wald" or "lrt"

level the confidence level required (default to 0.95).

margtrans string giving the marginal transformation, one of emp for rank-based transfor-

mation or none if data are already on the uniform scale

ties.method string indicating the type of method for rank; see rank for a list of options.

Default to "random"

plot logical; if TRUE, return a plot

... additional arguments to taildep, currently ignored

#### **Details**

The most common approach for estimation is the empirical survival copula, by evaluating the proportion of sample minima with uniform margins that exceed a given x. An alternative estimator uses a smoothed estimator of the survival copula using Bernstein polynomial, resulting in the so-called betacop estimator. Approximate pointwise confidence intervals for the latter are obtained by assuming the proportion of points is binomial.

#### Value

a data frame

• qlev: quantile level of estimates

• coef: point estimates

lower: lower bound of confidence intervalupper: lower bound of confidence interval

#### References

Ledford, A.W. and J. A. Tawn (1996), Statistics for near independence in multivariate extreme values. *Biometrika*, **83**(1), 169–187.

Ledford, A.W. and J. A. Tawn (1996), Statistics for near independence in multivariate extreme values. *Biometrika*, **83**(1), 169–187.

144 xdep.eta

## **Examples**

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 2, param = 0.7, model = "log")
xdep.chibar(dat, confint = 'wald')
## End(Not run)</pre>
```

xdep.eta

Coefficient of tail dependence

# Description

For data with unit Pareto margins, the coefficient of tail dependence  $\eta$  is defined via

$$\Pr(\min(X) > x) = L(x)x^{-1/\eta},$$

where L(x) is a slowly varying function. Ignoring the latter, several estimators of  $\eta$  can be defined. In unit Pareto margins,  $\eta$  is a nonnegative shape parameter that can be estimated by fitting a generalized Pareto distribution above a high threshold. In exponential margins,  $\eta$  is a scale parameter and the maximum likelihood estimator of the latter is the Hill estimator. Both methods are based on peaks-over-threshold and the user can choose between pointwise confidence obtained through a likelihood ratio test statistic ("lrt") or the Wald statistic ("wald").

# Usage

```
xdep.eta(
   xdat,
   qlev = NULL,
   nq = 40,
   qlim = c(0.8, 0.99),
   estimator = c("emp", "betacop", "gpd", "hill", "kj"),
   confint = c("wald", "lrt"),
   level = 0.95,
   margtrans = c("emp", "sp", "none"),
   ties.method = "random",
   plot = TRUE,
   mqlev = NULL,
   ...
)
```

#### **Arguments**

xdat an n by d matrix of multivariate observations qlev vector of percentiles between 0 and 1

145 xdep.eta

nq	number of quantiles of the structural variable at which to form a grid; only used if $u = NULL$ .
qlim	limits for the sequence u of the structural variable
estimator	string giving estimator to employ
confint	string indicating the type of confidence interval, one of "wald" or "lrt"
level	the confidence level required (default to 0.95).
margtrans	string giving the marginal transformation, one of emp for rank-based transformation or none if data are already on the uniform scale
ties.method	string indicating the type of method for rank; see rank for a list of options. Default to "random"
plot	logical; if TRUE, return a plot
mqlev	marginal quantile levels for semiparametric estimation for estimator kj; data above this are modelled using a generalized Pareto distribution. If missing, empirical estimation is used throughout
	additional arguments to taildep, currently ignored

## **Details**

The most common approach for estimation is the empirical survival copula, by evaluating the proportion of sample minima with uniform margins that exceed a given x. An alternative estimator uses a smoothed estimator of the survival copula using Bernstein polynomial, resulting in the socalled betacop estimator. Approximate pointwise confidence intervals for the latter are obtained by assuming the proportion of points is binomial.

## Value

a data frame

• qlev: quantile level of estimates

• coef: point estimates

• lower: lower bound of confidence interval • upper: lower bound of confidence interval

# References

Ledford, A.W. and J. A. Tawn (1996), Statistics for near independence in multivariate extreme values. Biometrika, 83(1), 169-187.

```
## Not run:
set.seed(765)
# Max-stable model
dat <- rmev(n = 1000, d = 2, param = 0.7, model = "log")
xdep.eta(dat, confint = 'wald')
## End(Not run)
```

146 xdep.xindex

xdep.xindex

Extremal index estimators

## **Description**

The function implements estimators of the extremal index based on interexceedance time and gap of exceedances. The maximum likelihood estimator and iteratively reweighted least square estimators of Suveges (2007) as well as the intervals estimator. The implementation differs from the presentation of the paper in that an iteration limit is enforced to make sure the iterative procedure terminates. Multiple thresholds can be supplied.

## Usage

```
xdep.xindex(
  xdat,
  qlev = 0.95,
  estimator = c("wls", "mle", "intervals"),
  confint = c("none", "wald", "lrt"),
  level = 0.95,
  plot = FALSE,
  warn = FALSE,
  ...
)
```

## **Arguments**

xdat	numeric vector of observations
qlev	a vector of quantile levels in $(0,1)$ for estimation of the extremal index. Defaults to $0.95$
estimator	a string specifying the chosen method (only one allowed). Must be either wls for weighted least squares, mle for maximum likelihood estimation or intervals for the intervals estimator of Ferro and Segers (2003). Partial match is allowed.
confint	string indicating the type of confidence interval, one of "wald" or "lrt" for estimator="mle", else "none"
level	the confidence level required (default to 0.95).
plot	logical; if TRUE, plot the extremal index as a function of q
warn	logical; if TRUE, receive a warning when the sample size is too small
	additional arguments, for backward compatibility

# **Details**

The iteratively reweighted least square is a procedure based on the gaps of exceedances  $S_n = T_n - 1$ . The model is first fitted to non-zero gaps, which are rescaled to have unit exponential scale. The slope between the theoretical quantiles and the normalized gap of exceedances is  $b = 1/\theta$ , with intercept  $a = \log(\theta)/\theta$ . As such, the estimate of the extremal index is based on  $\hat{\theta} = \exp(\hat{a}/\hat{b})$ . The

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weights are chosen in such a way as to reduce the influence of the smallest values. The estimator exploits the dual role of  $\theta$  as the parameter of the mean for the interexceedance time as well as the mixture proportion for the non-zero component.

The maximum likelihood is based on an independence likelihood for the rescaled gap of exceedances, namely  $\bar{F}(u_n)S(u_n)$ . The score equation is equivalent to a quadratic equation in  $\theta$  and the maximum likelihood estimate is available in closed form. Its validity requires however condition  $D^{(2)}(u_n)$  to apply; this should be checked by the user beforehand.

A warning is emitted if the effective sample size is less than 50 observations.

#### Value

a data frame

• qlev: quantile level of estimates

• coef: point estimates

lower: lower bound of confidence intervalupper: lower bound of confidence interval

## Author(s)

Leo Belzile

#### References

Ferro and Segers (2003). Inference for clusters of extreme values, JRSS: Series B, **65**(2), 545-556. Suveges (2007) Likelihood estimation of the extremal index. *Extremes*, **10**(1), 41-55.

Suveges and Davison (2010), Model misspecification in peaks over threshold analysis. *Annals of Applied Statistics*, **4**(1), 203-221.

Fukutome, Liniger and Suveges (2015), Automatic threshold and run parameter selection: a climatology for extreme hourly precipitation in Switzerland. *Theoretical and Applied Climatology*, **120**(3), 403-416.

```
set.seed(234)
# Moving maxima model with theta=0.5
a <- 1; theta <- 1/(1+a)
sim <- rgev(10001, loc=1/(1+a),scale=1/(1+a),shape=1)
x <- pmax(sim[-length(sim)]*a,sim[-1])
q <- seq(0.9, 0.99, by = 0.01)
xdep.xindex(
    xdat = x,
    qlev = q,
    estimator = "mle",
    confint = "wald")</pre>
```

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